

The Bhattacharyya distance: enriching the P-box in stochastic sensitivity analysis

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1 **Abstract:** The tendency of uncertainty analysis has promoted the transformation of sensitivity analysis
2 from the deterministic sense to the stochastic sense. This work proposes a stochastic sensitivity analysis
3 framework using the Bhattacharyya distance as a novel uncertainty quantification metric. The
4 Bhattacharyya distance is utilised to provide a quantitative description of the P-box in a two-level procedure
5 for both aleatory and epistemic uncertainties. In the first level, the aleatory uncertainty is quantified by a
6 Monte Carlo process within the probability space of the cumulative distribution function. For each sample
7 of the Monte Carlo simulation, the second level is performed to propagate the epistemic uncertainty by
8 solving an optimisation problem. Subsequently, three sensitivity indices are defined based on the
9 Bhattacharyya distance, making it possible to rank the significance of the parameters according to the
10 reduction and dispersion of the uncertainty space of the system outputs. A tutorial case study is provided in
11 the first part of the example to give a clear understanding of the principle of the approach with reproducible
12 results. The second case study is the NASA Langley challenge problem, which demonstrates the feasibility
13 of the proposed approach, as well as the Bhattacharyya distance metric, in solving such a large-scale, strong-
14 nonlinear, and complex problem.

15 **Keywords:** sensitivity analysis; uncertainty quantification, uncertainty propagation; Bhattacharyya
16 distance; probability box

17 1 Introduction

18 As uncertainty treatment in model Verification and Validation (V&V) becomes increasingly popular,
19 the Bhattacharyya distance has been investigated as a promising Uncertainty Quantification (UQ) metric in
20 stochastic model updating [1]. However, the deterministic methodologies are still widely used in practical
21 engineering, and the Euclidian distance is probably the most common metric providing a geometric distance
22 between two single data points. Alternatively, the Bhattacharyya distance is a stochastic metric between
23 two sets of samples considering their probability distributions. By capturing uncertainty sources from both
24 experimental and numerical data, the Bhattacharyya distance has been investigated as a more
25 comprehensive UQ metric [2]. Nevertheless, its application in V&V has been quite limited so far in both
26 academic and engineering fields. Upon to the authors' previous work on the role of the Bhattacharyya
27 distance in stochastic model updating [1], the main objective of this work is to further promote the

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28 application of the Bhattacharyya distance in stochastic sensitivity analysis (SA) within a two-level
29 procedure for uncertainty propagation and quantification.

30 Besides the stochastic model updating [3], the stochastic SA is another key component of V&V,
31 generally performed prior to model updating, with the purpose to evaluate and rank the significance of the
32 input parameters according to the system outputs [4]. One of the classical SA techniques is the global
33 sensitivity based on Sobol's indices [5,6]. However, as a variance-based approach, it cannot be directly
34 applied to applications where the parameters are described as imprecise random variables due to both
35 aleatory and epistemic uncertainties [7]. Despite the recent development on SA techniques (e.g. polynomial
36 expansion [8], covariance decomposition [9], Bayesian approach [10], analysis of variance [11], etc.), an
37 extension based on the imprecise probability theory [12–14] is required to achieve a deeper understanding
38 on both aleatory and epistemic uncertainties' contributions to the uncertainty of the system outputs [15].
39 Uncertainties in simulation and experiment processes can be divided into three sources:

- 40 • Uncertainties in parameterisation. The input parameters of the numerical model are imprecisely
41 determined, such as the materials properties of novel composites, geometry sizes of complex structures,
42 and random boundary conditions lead by the winds or earthquakes.
- 43 • Uncertainties in modelling. The numerical model always contains simplifications and approximations,
44 such as the linearization of nonlinear behaviours, the hypothesis of frictionless joints, and the
45 simplification of complex connections.
- 46 • Uncertainties in experiments. The measurements are driven by hard-to-control randomnesses, such as
47 environment noise, measurement system errors, and human subjective judgements. More explanations
48 of experimental uncertainty can be found in Ref. [16].

49 Due to the above unavoidable uncertainty sources, this work focuses on the implementation of a
50 stochastic SA approach in which the Bhattacharyya distance can be embedded as a comprehensive and
51 convenient UQ metric by capturing various uncertainty sources. In this work, the input model parameters
52 are no longer treated as unknown-but-fixed constants, but investigated as four categories according to the
53 involvement of aleatory uncertainty (natural variation) or/and epistemic uncertainty (lack of knowledge):

- 54 I) Parameters without any uncertainty, appearing as explicit constants;
- 55 II) Parameters with only epistemic uncertainty, appearing as unknown-but-fixed constants, bounded by a
56 given interval;
- 57 III) Parameters with only aleatory uncertainty, appearing as random variables with fully determined
58 probability characteristics such as distribution type, mean, variance, etc.;
- 59 IV) Parameters with both aleatory and epistemic uncertainties, appearing as imprecise probability variables
60 with only vaguely determined uncertainty characteristics.

61 The above parameter categorisation is critical in this work, since different parameter categories have
62 different representations of their uncertainty characteristics, and thus requiring different treatment for
63 uncertainty quantification and propagation. To represent the uncertainty space of variables with imprecise
64 probability, the *Probability box* (generally known as P-box) is proposed by Ferson et al. [17] as a graphic
65 representation, which has been widely used in uncertainty analysis [18–20]. In the following context, the

66 term “*uncertainty space*” is utilised as a general expression of the imprecise probability measure of an
67 uncertain variable. In the typical probability theory, the probability space is defined as the mathematic
68 triplet $(\Omega, \mathcal{F}, \mathbb{P})$, where Ω is the complete event space, \mathcal{F} is the σ -algebra space of all possible events, \mathbb{P} is
69 the probability measure of the event. The probability space is based on the assumption that the probability
70 measure \mathbb{P} is precisely known, implying only the aleatory uncertainty is considered. In the context of
71 imprecise probability, however, the aleatory and epistemic uncertainties can occur simultaneously. Hence,
72 the uncertainty space herein can be seen as an extension of the precise probability space. More detailed
73 information of the uncertainty space (i.e. imprecise probability space) and P-box can be referred to Ref.
74 [18].

75 The application of P-boxes in SA has been developed, e.g., by Ferson et al. [21] via the Probability
76 Bounds Analysis (PBA), which is further developed by Alvarez [22] in the non-specificity SA approach.
77 However, the P-box still requires a quantitative measure to define an explicit ranking index in SA.
78 Meanwhile, a further investigation of the P-box is required to differentiate the effects of the aleatory and
79 epistemic uncertainties. In this work, the proposed stochastic SA approach is constructed within a two-level
80 framework to quantify the aleatory and epistemic uncertainties, along the vertical and horizontal directions
81 of the P-box. The Monte Carlo simulation and optimisation techniques are respectively utilised in these two
82 levels to quantify and propagate uncertainties from the system inputs to outputs. The Bhattacharyya distance
83 acts as the core of the SA framework by providing a quantitative measure of the P-box, which is the
84 foundation for the proposed sensitivity indices. The three indices, namely the proportional index, the
85 variance-based index, and the comprehensive index, are utilised to rank the significance of the parameters
86 according to the reduction and dispersion characteristics of the output uncertainty space.

87 Two case studies are proposed to demonstrate the feasibility of the overall two-level approach, as well
88 as the Bhattacharyya distance metric. The first case study, investigating the Ishigami function with an
89 explicit formulation, allows the readers to achieve a better understanding of the principle of the approach
90 with reproducible results. The second case study solves the SA task of the NASA UQ challenge problem
91 [23]. Results of the current work are compared with previously published works [7,24,25] to demonstrate
92 the Bhattacharyya distance as a competent UQ metric in the proposed stochastic SA approach.

93 The novelties of this work involve three aspects. First, the Bhattacharyya distance is proposed as a novel
94 UQ metric, which can be conveniently embedded into the overall SA framework to provide a quantitative
95 measure of the P-box (i.e. the uncertainty space). Second, this work performs the stochastic SA in the
96 background of uncertainty treatment, and more specifically, to investigate the significance of an input
97 according to how much the uncertainty space of the output can be changed, when the uncertainty space of
98 this input is reduced. Third, the proposed two-level approach provides a clear logic to differentiate the
99 contributions of the aleatory and epistemic uncertainties to the uncertainty space of the output, by
100 investigating the P-box along two directions in the two levels, respectively.

101 The following parts of this paper are organised as follows. Section 2 investigates the typical P-box from
102 a novel viewpoint: the aleatory and epistemic uncertainty representations along the vertical and horizontal
103 directions of the P-box, respectively. Section 3 presents the two-level procedure for uncertainty propagation

104 and quantification, where the Monte Carlo simulation is performed in the first level for aleatory uncertainty
 105 propagation, and an optimisation is performed in the second level for epistemic uncertainty quantification.
 106 Section 4 explains how to embed the Bhattacharyya distance into the two-level procedure, and subsequently,
 107 three sensitivity indices are defined based on the Bhattacharyya distance. The variation of the results and
 108 calculation cost are also discussed in this section. The Ishigami function and the NASA UQ challenge
 109 problem are investigated in Sections 5 and 6, respectively. Section 7 gives the conclusions and perspectives.

110 2 P-box representation of the imprecise probability uncertainty space

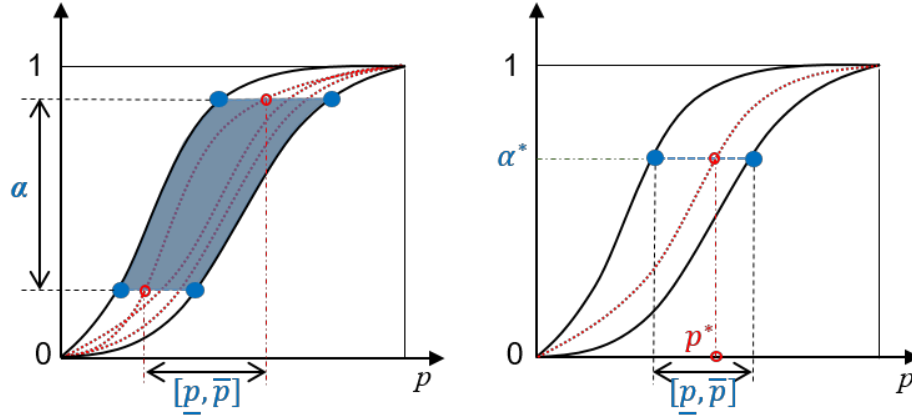
111 The P-box provides a clear graphic representation of the uncertainty space of an imprecise random
 112 variable. Since the stochastic SA approach in this work is proposed in the background of imprecise
 113 probability, the distributional P-box is investigated in this section. A distributional P-box is essentially a
 114 family of Cumulative Distribution Functions (CDF) of a random variable, encompassing an infinite number
 115 of CDF curves. The CDF family of a variable p is expressed as

$$116 \quad \mathcal{F}(p) \supseteq \mathcal{F}(p, \boldsymbol{\theta}), \quad \boldsymbol{\theta} \in [\boldsymbol{\theta}, \bar{\boldsymbol{\theta}}] \quad (1)$$

117 where $\mathcal{F}(p)$ is the CDF curves family, $\boldsymbol{\theta}$ is the distribution coefficients of p . The variable p can fall into
 118 any one of the parameter categories described in Section 1, according to the involvement of different types
 119 of uncertainties. Variables belonging to different categories have different formats of P-boxes, which will
 120 be further investigated in Section 3. For the most complex case when both aleatory and epistemic
 121 uncertainties are involved, i.e. a Category IV variable, the epistemic uncertainty is presented by the interval
 122 $[\boldsymbol{\theta}, \bar{\boldsymbol{\theta}}]$. This interval leads to infinite number of CDF curves within the distributional P-box, and this is the
 123 reason that a P-box is also known as a uncertainty space of an imprecise probability variable p . Fig. 1
 124 illustrates the P-box of a Category IV variable, where the lower and upper bounds of the curve family $\underline{\mathcal{F}}$
 125 and $\bar{\mathcal{F}}$ can be determined by the interval of the distribution coefficients $[\boldsymbol{\theta}, \bar{\boldsymbol{\theta}}]$. The shape of a CDF curve
 126 (horizontal position and slope) is driven by the mean and variance of a distribution. The horizontal position
 127 of the CDF curve is controlled by the mean value; the slope of the CDF curve is controlled by the variance
 128 value, i.e. the dispersion degree of the distribution. Based on this principle, the bounds of the P-box have
 129 four possibilities as follows:

- 130 • The distribution with maximum mean (far right position) and maximum variance (moderate slope);
- 131 • The distribution with maximum mean (far right position) and minimum variance (steep slope);
- 132 • The distribution with minimum mean (far left position) and maximum variance (moderate slope);
- 133 • The distribution with minimum mean (far left position) and minimum variance (steep slope).

134 The bounds of the P-box can be determined by a simple comparison among the four distributions. Note that,
 135 in some cases the P-box bound is not a complete CDF curve of a specific distribution, but a combination of
 136 multiple CDF curves. A demonstration to determine bounds of P-boxes with different distribution forms is
 137 provided in the first case study (Section 5.2).



(a) Vertical direction for aleatory uncertainty (b) Horizontal direction for epistemic uncertainty

Fig. 1: Investigation of the P-box along two directions

According to the typical categorisation of aleatory and epistemic uncertainties, the aleatory uncertainty refers to the irreducible uncertainty caused by the natural variation of the system; the epistemic uncertainty refers to the reducible uncertainty caused by the lack of knowledge. In stochastic SA, it is important to have a deeper understanding of these two kinds of uncertainties, especially their contributions to the output uncertainty space. This can be achieved through a further investigation of the P-box. As shown in Fig. 1, the P-box is investigated along to two directions, i.e. the vertical and horizontal axes, with different emphases on the aleatory uncertainty and the epistemic uncertainty, respectively.

Fig. 1(a) illustrates the investigation of the P-box along the vertical direction within the overall probability range $[0, 1]$. The cumulative probability range α truncated from the vertical axis corresponds to the truncated area within the P-box. Considering the basic hypothesis that the P-box is a set of infinite number of CDF curves, each single CDF curve in the P-box represents a precise probability variable with only aleatory uncertainty. For a single CDF curve in Fig. 1(a), the probability range α corresponds to the variable interval $[\underline{p}, \bar{p}]$ containing only aleatory uncertainty information.

The horizontal investigation of the P-Box focuses on a specific probability value α^* within the range α , as shown in Fig. 1(b). In the following context, the quantity with superscript * denotes an explicit constant. For a single CDF curve, a fixed probability value α^* in the vertical axis corresponds to a fixed variable value p^* in the horizontal axis. However, because of the epistemic uncertainty, the CDF family contains infinite number of variable values corresponding to the same probability value α^* , and thus comprise into the variable interval $[\underline{p}, \bar{p}]$. This interval, different from the one obtained in Fig. 1(a), contains only epistemic uncertainty information.

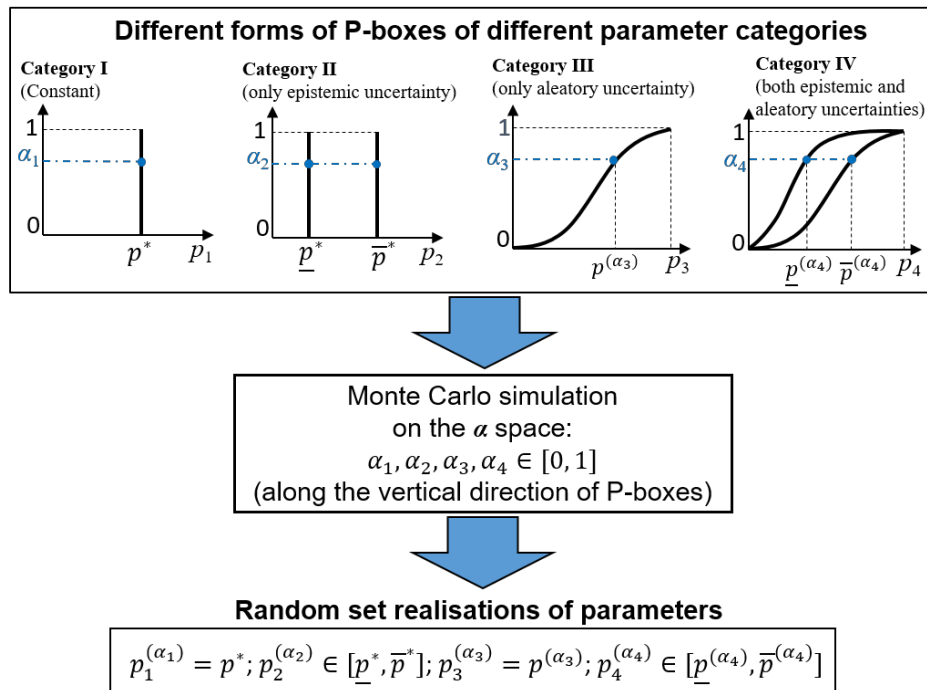
The above two-direction investigation provides a clear logic to differentiate the contributions of the aleatory and epistemic uncertainties to the P-box of an imprecise probability variable. This investigation is the basis of the two-level procedure in Section 3 to quantify and propagate the uncertainties from the system inputs to the outputs.

165 3 A two-level procedure for uncertainty propagation and quantification

166 In this section, a two-level procedure is proposed based on the idea of the two-direction investigation of
 167 the P-box in Section 2. The Monte Carlo simulation and optimisation techniques are utilised in the first and
 168 second levels, focusing on the aleatory and epistemic uncertainties, respectively. Note that, the analysis of
 169 the two kinds of uncertainties cannot be completely separated, but should be performed simultaneously in
 170 a uniform framework. The construction of the two levels into a uniform framework is provided in the end
 171 of this section.

172 3.1 Level I: Monte Carlo simulation for the aleatory uncertainty

173 The first level of this overall procedure focuses on the aleatory uncertainty through the Monte Carlo
 174 simulation to randomly sample the probability value α^* along the vertical direction of the P-box, as shown
 175 in Fig. 2. Different categories of parameters have different forms of CDFs/P-boxes, and thus should be
 176 analysed separately. As described in Section 1, there are four parameter categories according to the
 177 involvement of aleatory and/or epistemic uncertainties. The treatment for each parameter category is
 178 explained as follows.



179

180

Fig. 2: Monte Carlo simulation for different categories of parameters

181 **Category I):** This category of parameter contains no uncertainty. Its CDF appears as an impulse function
 182 at the fixed position p^* , with the amplitude as 1. During Monte Carlo simulation, for arbitrary probability
 183 value α_1 , the corresponding parameter value is always p^* .

184 **Category II):** As an unknown-but-fixed constant fallen within an interval due to the epistemic uncertainty,
 185 this kind of parameter has a family of impulse functions, bounded by the given interval $[\underline{p}^*, \bar{p}^*]$. The P-box
 186 for this kind of parameter appears as a standard rectangle. For randomly sampled value α_2 , the

187 corresponding parameter value is no longer a single value, but a range which always equals to the pre-
188 defined interval $[\underline{p}^*, \overline{p}^*]$.

189 **Category III):** Involving only aleatory uncertainty, this category of parameter is a variable following a
190 fully determined distribution, and thus it has a single CDF curve. During Monte Carlo simulation, for
191 arbitrary value α_3 , the corresponding parameter value is a constant, however, varying according to the CDF
192 curve. The Categories I and III parameters are not affected by the epistemic uncertainty, and thus these
193 parameters do not need to be updated during a model updating procedure. Nevertheless, the Categories I
194 and III parameters still require appropriate representation and quantification, such that their contributions
195 to the output uncertainty space can be clearly differentiate from that of the epistemic uncertainty.

196 **Category IV):** Since both aleatory and epistemic uncertainties are involved, this category of parameters
197 have the normal form P-box as illustrated in Fig. 1. For each sampled probability value α_4 , an interval of
198 the parameter is obtained during the Monte Carlo simulation. Furthermore, bounds of this interval are
199 changing according to different instances of α_4 , because of the P-box with curvilinear bounds as shown in
200 Fig. 2.

201 In the following context, the term “*random set realisation*” is utilised to designate the different
202 realisations (i.e. fixed/varying point or interval) of the parameter. The above parameter categorisation
203 provides a clear understanding of different kinds of uncertainties’ influences on the resulting random set
204 realisations. Category I represents ideally deterministic parameter whose random set realisation is always
205 a constant point. When aleatory uncertainty is involved, the parameter moves from Category I to Category
206 III, where the random set realisation changes from a fixed point to a changing point, according to different
207 probability values α . The epistemic uncertainty renders the random set realisation no long a point, but an
208 interval, by moving the parameter from Category I to Category II. Finally, for Category IV parameter,
209 because of both aleatory and epistemic uncertainties, the resulting random set realisation is an interval with
210 changing bounds according to arbitrary probability values α during Monte Carlo simulation. As the
211 outcome of the first level, the random set realisations will be served as constrains of an optimisation problem
212 in the second level to propagate the uncertainties from the input parameters to the outputs.

213 3.2 Level II: Optimisation for the epistemic uncertainty

214 The second level is committed to the epistemic uncertainty contained in the random set realisations, and
215 to propagate the epistemic uncertainty from the inputs to the outputs. Here the uncertain system refers to
216 the numerical simulation process containing three key components: input parameters \mathbf{p} , outputs \mathbf{x} , and
217 simulator $h(\cdot)$:

$$218 \quad \mathbf{x} = h(\mathbf{p}). \quad (2)$$

219 The simulator can be a sophisticated finite element model of an engineering structure, or a simple
220 mathematical function extracted from a complex system. The uncertainty propagation is achieved by
221 solving an optimisation problem to determine the minimum and maximum of the outputs, using the random
222 set realisations of the parameters obtained from the first level. The optimisation problem is expressed as:
223 finding

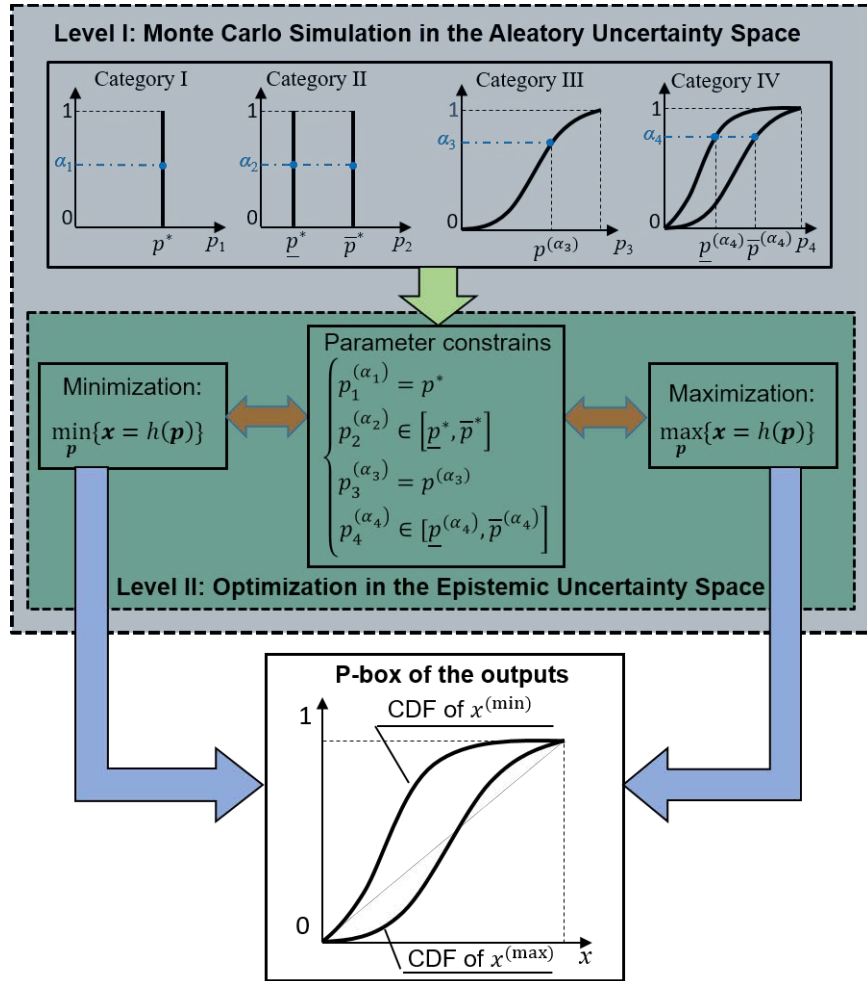
224 $\min_{\mathbf{p}}\{x = h(\mathbf{p})\}$ and $\max_{\mathbf{p}}\{x = h(\mathbf{p})\}$, (3)

225 using the random set realisations as constraints

226
$$\begin{cases} p_1^{(\alpha_1)} = p^* \\ p_2^{(\alpha_2)} \in [\underline{p}^*, \bar{p}^*] \\ p_3^{(\alpha_3)} = p^{(\alpha_3)} \\ p_4^{(\alpha_4)} \in [\underline{p}^{(\alpha_4)}, \bar{p}^{(\alpha_4)}] \end{cases} \quad (4)$$

227 where the superscript * denotes that the value is fixed, the superscript (α) denotes that the value is changing
 228 according to the arbitrary α , the subscripts “1-4” of α and p denote the categories of the parameters.

229 Note that, Eq. (4) provides the optimisation problem with simple interval constraints without any
 230 complex nonlinear constraint. Furthermore, the interval constraints represent only the epistemic uncertainty,
 231 implying the ranges of the intervals are much smaller the whole domain of definition of the parameters in
 232 the system. Consequently, the optimisation problem can be solved easily and rapidly by the typical
 233 techniques, such as simplex algorithm and interior point method.



234 Fig. 3: The two-level procedure for uncertainty quantification and propagation
 235

236 The overall two-level procedure for uncertainty quantification and propagation from the inputs to the
 237 outputs is illustrated in Fig. 3. Note that, this procedure contains two levels, but not two steps, because the
 238 optimisation procedure (Level II) is performed for each sampled point α within the Monte Carlo simulation
 239 (Level I). Suppose the sampling size in the Monte Carlo simulation is N_{MC} , N_{MC} random set realisations of
 240 the input parameters will be obtained and the optimisation is executed N_{MC} times, generating N_{MC} pairs of
 241 minimum and maximum output values. The CDFs of minimum and maximum outputs are estimated, and
 242 the P-box of the output is bounded by these fitted CDFs as shown in the bottom of Fig. 3.

243 The above two-level approach is significant for stochastic SA, since it provides a clear logic for
 244 quantification of both aleatory and epistemic uncertainties, which makes it possible to measure and
 245 differentiate the contributions of these two kinds uncertainties to the output uncertainty space.

246 4 Sensitivity indices based on the Bhattacharyya distance

247 4.1 Bhattacharyya distance: a quantification metric for the P-box

248 The P-box provides a clear representation of the uncertainty space, however, it is still insufficient for
 249 SA because of the following reasons:

- 250 • SA requires a quantitative measure of the uncertainty space to give an explicit parameter ranking, while
 251 the P-box is only a graphic representation;
- 252 • The estimation of the CDFs based on the output samples is computationally expensive, making a precise
 253 representation of the P-box unpractical for complex systems;
- 254 • When multiple outputs are considered in a system, CDFs of the outputs become multi-dimensional, and
 255 the construction of the P-box in a multi-dimensional space is out of imagination.

256 As a result, the Bhattacharyya distance is proposed as a quantification metric for the P-box, which is
 257 quantitative, distribution-free, and feasible for more than one output. The basic principle of the
 258 Bhattacharyya distance and its application in stochastic model updating have been elaborated in Ref. [1].
 259 Hence, the current work is not focusing on the evaluation method of the Bhattacharyya distance, but a
 260 stochastic SA framework where the Bhattacharyya distance can be conveniently implanted.

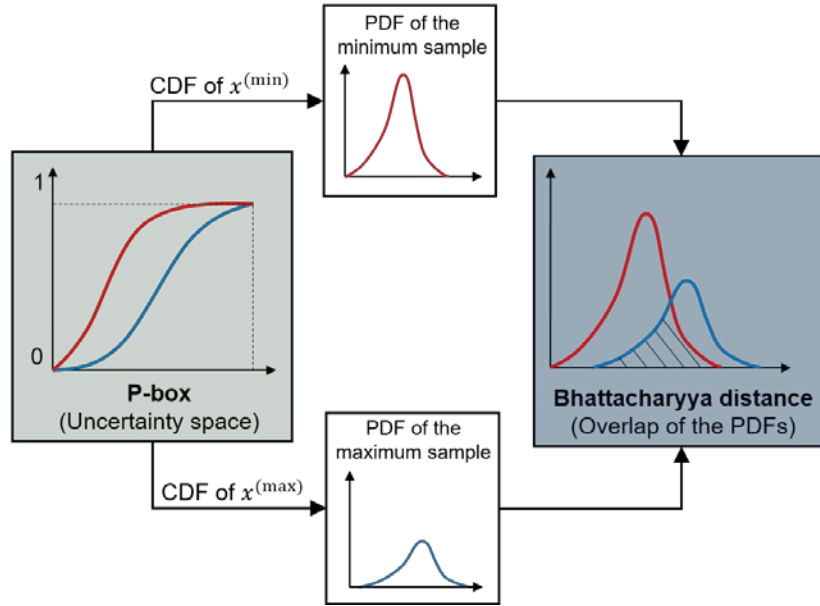
261 Fig. 4 illustrates the relationship between the Bhattacharyya distance and the P-box representation of a
 262 single output x . Instead of using the estimated CDFs of the output samples to construct the P-box, the
 263 Bhattacharyya distance is proposed to directly evaluate the overlap between the minimum and maximum
 264 output samples. This treatment is executed directly after the two-level procedure in Section 3, where the
 265 minimum and maximum output samples are obtained.

266 The Bhattacharyya distance between two discrete distributions is defined as

$$267 \quad D(\mathbf{x}_{min}, \mathbf{x}_{max}) = -\log \left\{ \sum_{i_m=1}^{n_{bin}} \dots \sum_{i_1=1}^{n_{bin}} \sqrt{PM_{min}(b_{i_1, i_2, \dots, i_m}) PM_{max}(b_{i_1, i_2, \dots, i_m})} \right\} \quad (5)$$

268 where \mathbf{x}_{min} and \mathbf{x}_{max} are the minimum and maximum samples of the outputs; m is the number of the
 269 outputs; n_{bin} is the number of bins defined for each of the m outputs; $PM_{\blacksquare}(b_{i_1, i_2, \dots, i_m})$ is the joint
 270 probability mass function (PMF) of the bin b_{i_1, i_2, \dots, i_m} . The bin has m subscripts because it is generated under

271 a m -dimensional joint-PMF space. The detailed evaluation method of the joint-PMF function can be
 272 referred to Ref. [1]. For theoretical completeness, the evaluation procedure is simply recalled as follows.



273 Fig. 4: Relationship between the Bhattacharyya distance and the P-box

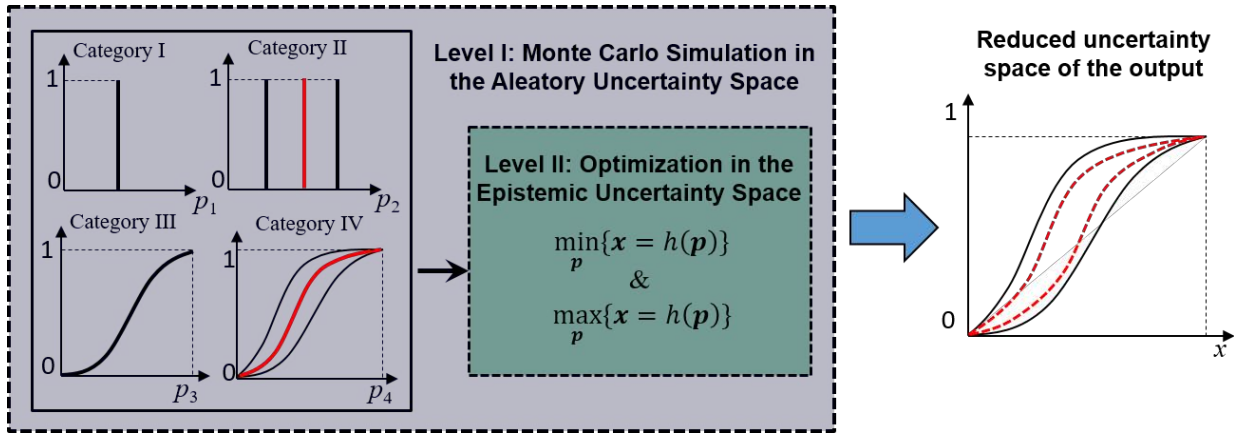
- 274
- 275 1) The output samples \mathbf{x}_{min} and \mathbf{x}_{max} appear as matrices with size as $\mathbf{x}_{\blacksquare} \in \mathbb{R}^{N \times m}$, where N is the number
 276 of data points from Monte Carlo simulation, and m is the number of outputs. Considering the i -th
 277 columns of the two matrices (\mathbf{x}_{min} and \mathbf{x}_{max}), find the lower and upper bounds of all values in both
 278 columns. These lower and upper bounds define an interval $I_i (\forall i = 1, \dots, m)$ containing all the i -th
 279 output values in \mathbf{x}_{min} and \mathbf{x}_{max} .
 - 280 2) Within the defined interval I_i , decide the number of bins $n_{bin} \cong \lceil \frac{N}{10} \rceil$, where $\lceil \blacksquare \rceil$ denotes the upper
 281 integer of the investigating values;
 - 282 3) Count the number of points fallen into each bin, i.e. the frequency. Note that, when m is larger than one,
 283 the frequency should be counted in a m -dimensional PMF space, and the total number of bins in this
 284 joint-PMF space is $N_{bin} = n_{bin}^m$.

285 The above procedure is based on the discrete PMF of data samples and thus is applicable to any sample
 286 set regardless of its exact distribution format. This makes it especially appropriate for the current application
 287 when the sample size after the two-level procedure is too limited to give a precise distribution estimation.
 288 Another advantage of the Bhattacharyya distance is that it provides a scalar measure for all the outputs
 289 simultaneously, which fulfils the expectation as an elaborate, quantitative, and uniform measure of multiple
 290 uncertain outputs.

291 4.2 Stochastic sensitivity indices

292 The final objective of the stochastic SA, in the background of uncertainty quantification, is described as:
 293 Quantify the importance of the input according to how much the uncertainty space of the output can be
 294 changed, when the *epistemic uncertainty* of this input is reduced. This objective is illustrated in Fig. 5 where

295 the epistemic uncertainties of the Categories II and IV parameters are completely reduced. For the Category
 296 II parameter, the completely reduction of the epistemic uncertainty leads to an impulse function (the red
 297 vertical line in the figure). For the Category IV parameter, this will lead to a single CDF curve (the red solid
 298 curve in the figure). In the right part of Fig. 5, the change of the output P-box represents the significance of
 299 the input parameters. In order to give a quantitative measure, an explicit sensitivity index based on the
 300 Bhattacharyya distance is required to rank the significance of each input parameter. Three sensitivity indices,
 301 namely the proportional index, the variance-based index, and the comprehensive index, are proposed as
 302 follows. Since the reduction of the epistemic uncertainty is executed to either the Category II parameter or
 303 to the Category IV parameter, the following three sensitivity indices apply to both Categories II and IV
 304 parameters.



305
 306 Fig. 5: Reduced output uncertainty space when the epistemic uncertainties of
 307 Categories II and IV parameters are reduced

308 • *The proportional index*

309 As shown in Fig. 5, the most intuitive manner to measure how much the output uncertainty space is
 310 reduced is to calculate the proportion ratio between the reduced uncertainty space and the original one:

311
$$S_i^{(pro)} = \frac{D_0(\mathbf{x}) - D_i(\mathbf{x}, p_i)}{D_0(\mathbf{x})} \quad (6)$$

312 where $S_i^{(pro)}$ is the proportional sensitivity index of the i -th input parameter; $D_0(\mathbf{x})$ is the Bhattacharyya
 313 distance of the original uncertainty space of the output; $D_i(\mathbf{x}, p_i)$ is the Bhattacharyya distance of the output
 314 uncertainty space with the epistemic uncertainty of the i -th input parameter completely reduced.

315 Eq. (6) provides a measure of the parameter significance with a direct and simple principle. However,
 316 an obvious defect of this equation is that different specified input distributions have different influences on
 317 the output uncertainty space. As shown in Fig. 5, there are infinite CDF curves in the P-box of the Category
 318 IV parameter, while different curves lead to different output uncertainty spaces. Only randomly
 319 investigating one single CDF curve from the input uncertainty space is obviously insufficient. A simple
 320 way to deal with this problem is to investigate a series of specified distributions of the parameter, and use
 321 the mean of the output uncertainty spaces to replace the single one in Eq. (6). Hence, the improved
 322 proportional index is expressed as

323
$$S_i^{(pro)} = \frac{D_0(\mathbf{x}) - \mu_{D_i}(\mathbf{x}, p_i)}{D_0(\mathbf{x})} \quad (7)$$

324 where $\mu_{D_i}(\mathbf{x}, p_i)$ denotes the mean of the multiple Bhattacharyya distance with a series of determined
 325 distributions of p_i . As explained by Eq. (1), the CDF family of input parameter is driven by the interval of
 326 its distribution coefficient (e.g. mean or variance) $[\underline{\theta}, \bar{\theta}]$. Hence these different input distributions can be
 327 specified by assigning a certain number of equidistant values within the coefficient interval. This treatment
 328 is quite sensitive to the number of the coefficients required to describe the P-box. Most distribution types,
 329 e.g. Gaussian distribution $N(\mu, \sigma^2)$, uniform distribution $U(a, b)$, and Beta distribution $Beta(a, b)$, contain
 330 only two distribution coefficients, and thus the full factorial design is feasible to generate a grid of
 331 equidistant values. In case of complex distributions with three or even more coefficients, the Design of
 332 Experiment (DoE) method is suggested. Considering the typical DoE technique using the orthogonal Latin
 333 squares, when there are four coefficients considered and each coefficient contains eight levels, the full
 334 factorial design results into $8^4 = 4096$ realisations. In contrast, the orthogonal Latin square design only
 335 requires $8^2 = 64$ realisations, which is still the same number as the full factorial design of two-coefficient
 336 case. More information of the orthogonal Latin square DoE method can be referred to Ref. [26]. The
 337 detailed demonstration for the assignment of specified input distributions can be found in Section 5.3.

338 In the following context, the number of the specified input distributions is assumed as N_{sp} , and thus
 339 $\mu_{D_i}(\mathbf{x}, p_i)$ in Eq. (7) is estimated from a set of Bhattacharyya distance samples with size as N_{sp} . The
 340 Bhattacharyya distance samples with different sizes lead to different $\mu_{D_i}(\mathbf{x}, p_i)$, and subsequently, different
 341 sensitivity index values. The variation of the index is investigated in the following subsection.

342 • *The variance-based index*

343 The significance of the input parameter is not only reflected by the reduction degree of the output
 344 uncertainty space, but also reflected by the dispersion degree of the output uncertainty space when the input
 345 P-box is reduced to different single curves. Similar as the above proportional index, a series of specified
 346 distributions of p_i are utilised, and multiple Bhattacharyya distances are obtained for the corresponding
 347 output uncertainty spaces. The dispersion degree of the output uncertainty spaces is measured by the
 348 coefficient of variation of the Bhattacharyya distance samples:

349
$$S_i^{(var)} = \frac{\sigma_{D_i}(\mathbf{x}, p_i)}{\mu_{D_i}(\mathbf{x}, p_i)} \quad (8)$$

350 where the superscript (*var*) denote the variance-based index; $\sigma_{D_i}(\mathbf{x}, p_i)$ is the standard deviation of the
 351 Bhattacharyya distances with the epistemic uncertainty of p_i completely reduced; $\mu_{D_i}(\mathbf{x}, p_i)$ is the mean of
 352 the Bhattacharyya distances. $\sigma_{D_i}(\mathbf{x}, p_i)$ and $\mu_{D_i}(\mathbf{x}, p_i)$ are estimated from a set of Bhattacharyya distance
 353 samples with size as N_{sp} .

354 • *The comprehensive index*

355 As discussed above, the proportional index reflects the reduction degree of the output uncertainty; and
 356 the variance-based index reflects the dispersion degree of the output uncertainty. Clearly, neither of the

357 above indices can give a full measure of the significance of the input according to the overall change of the
 358 output uncertainty space. It is reasonable to make a product of these two indices so that a more
 359 comprehensive index is obtained as

$$360 \quad S_i^{(com)} = \frac{D_0(\mathbf{x}) - \mu_{D_i}(\mathbf{x}, p_i)}{D_0(\mathbf{x})} \frac{\sigma_{D_i}(\mathbf{x}, p_i)}{\mu_{D_i}(\mathbf{x}, p_i)} \quad (9)$$

361 where the superscript (*com*) denotes the comprehensive index. A larger proportional index $S_i^{(pro)}$ indicates
 362 a stronger reduction effect of the input epistemic uncertainty on the output uncertainty; a larger variance-
 363 based index $S_i^{(var)}$ indicates a larger variation of the output uncertainty space due to different reduction
 364 manners of the input epistemic uncertainty. Consequently, when the comprehensive index $S_i^{(com)}$ is larger,
 365 the input parameter p_i is comprehensively more significant to the output uncertainty space, on both
 366 reduction and dispersion effects.

367 4.3 Variation of the stochastic sensitivity indices

368 A common feature of the stochastic SA methodologies employing Monte Carlo sampling is that the
 369 resulting sensitivity index is also stochastic. Specifically, the stochastic feature of the three sensitivity
 370 indices in Section 4.2 is derived from two aspects:

- 371 1) The evaluation of the Bhattacharyya distance utilises the random samples generated from Monte Carlo
 372 simulation. Different random samples lead to slightly different Bhattacharyya distance values.
- 373 2) The definition of the sensitivity indices requires not a single specified parameter distribution (i.e. a
 374 single CDF curve within a P-box), but a series of specified distributions of the parameter. Different
 375 numbers and configurations of the selected CDF curves in the P-box lead to variation of the sensitivity
 376 indices.

377 Consequently, it is important to investigate the variation feature of the sensitivity indices, especially,
 378 when the significances of two parameters are comparable. The bootstrap technique [27] is employed herein
 379 to estimate the mean and standard deviation of the stochastic sensitivity indices. Considering the existing
 380 sample set of Bhattacharyya distances with size N_{sp} described in Section 4.2, the basic principle of bootstrap
 381 technique in this work is to generate a new sample set (termed as “bootstrap sample set”) with the same
 382 size as the existing sample set (i.e. N_{sp}), using the strategy *sampling with replacement*. A new value of the
 383 sensitivity index can be calculated based on this bootstrap sample set. Assuming the bootstrap process is
 384 repeated N_{bs} times, N_{bs} bootstrap samples are generated, and subsequently a sample set of the sensitivity
 385 index values is obtained with the size as N_{bs} . In the following case studies, the mean of the sensitivity index
 386 samples is utilised to rank the significance of the input parameters; the standard deviation of the sensitivity
 387 index is provided to assess the variation degree of the index. The smaller the standard deviation, the more
 388 precise the index is within the overall stochastic SA approach.

389 4.4 Discussion of the calculation cost

390 The overall two-level procedure requires an optimisation problem to be solved for each Monte Carlo
 391 sample, leading to considerable calculation cost. However, the optimisation problem itself for each sample

392 can be solved easily and rapidly using the typical optimisation techniques such as simplex algorithm and
 393 interior point method. The simplicity of the optimisation problem is caused by the following reasons.

- 394 • The optimisation problem contains only interval constraints of the parameters, without any complex
 395 and nonlinear constraint.
- 396 • The interval constraints for each Monte Carlo sample are determined in the first level in Section 3,
 397 representing only epistemic uncertainty of the P-box. Consequently, the ranges of the interval
 398 constraints are much smaller than the whole domain of definition of the parameters in the system.

399 These interval constraints with small ranges represent an explicit and reduced searching space, which
 400 provides convenience when solving the optimisation problem. Even for the complex and strong-nonlinear
 401 problem presented in the second case study (Section 6), the optimisation problem can still be solved by the
 402 typical interior point method with the CPU calculation time for less than one second in a desk-top computer.

403 Although the calculation cost for the optimisation is acceptable, in case of complex numerical models
 404 (e.g. sophisticated finite element models), the meta-models are suggested as a common rule for Monte Carlo
 405 based methodologies. Since each Monte Carlo sampling procedure is independent to each other, the parallel
 406 computation [28] can be utilised in this approach to further reduce the calculation time.

407 5 Case study I: The Ishigami function

408 5.1 Problem description

409 A tutorial case study is given in this section utilising the Ishigami function [29], which is a general
 410 example for SA techniques. The Ishigami function is defined as

$$411 \quad f(\mathbf{p}) = \sin(p_1) + a\sin^2(p_2) + bp_3^4\sin(p_1), \quad (10)$$

412 where p_1, p_2 , and p_3 are the input parameters; a and b are the pre-defined coefficients. In this work, the
 413 values of a and b are assigned to be the same as Marrel et al. [30]: $a=7$ and $b=0.1$. The uncertainty
 414 characteristic of p_{1-3} is generally set as the uniform distribution on the interval $[-\pi, \pi]$ in the literature.
 415 However, this uncertainty characteristic cannot fulfil the application in this work, because the fully
 416 determined distribution implies p_{1-3} belong to Category III parameters with only aleatory uncertainty.
 417 Considering the objective of the stochastic SA herein, the epistemic uncertainty is the investigating
 418 emphasis to assess the uncertainty space of an imprecise distribution, rather than a single CDF curve of a
 419 fully determined distribution. As a result, more complex uncertainty characteristics of the parameters are
 420 assigned as listed in Table 1. p_1 and p_2 are prescribed to follow the uniform distribution, and p_3 follows the
 421 Gaussian distribution. The distribution coefficients, e.g. the bounds, mean, and variance, are not fully
 422 determined but fallen into intervals. That is to say, p_{1-3} belong to Category IV parameter in this example.

423 Table 1: Uncertainty characteristics of the input parameters

Category	Parameter Distribution	Epistemic Coefficient
IV	$p_1 \sim U(a_1, b_1)$	$a_1 \in [-4.0, -3.0]; b_1 \in [2.0, 3.0]$
IV	$p_2 \sim U(a_2, b_2)$	$a_2 \in [-3.0, -1.0]; b_2 \in [3.0, 5.0]$
IV	$p_3 \sim N(\mu_3, \sigma_3^2)$	$\mu_3 \in [0.0, 1.0]; \sigma_3 \in [\sqrt{5}, \sqrt{2}]$

424 The uncertainty characteristics shown in Table 1 lead to a stochastic SA task different from the normal
 425 demonstration on the Ishigami function in the literature, and thus a comparison between the current results
 426 and the published results is unavailable. The purpose of this case study is to provide a reproducible result
 427 such that a better understanding on the principle of the overall approach is achieved. Detailed assessment
 428 and comparison between the results of the proposed approach and the published approaches are presented
 429 in the next case study on the NASA UQ Challenge.

430 5.2 P-boxes of the inputs and outputs

431 The first step of this demonstration is to determine the P-boxes of the input parameters, based on the
 432 uncertainty characteristics listed in Table 1. The P-box of an imprecise uniform distribution can be easily
 433 determined by the bounds of the epistemic coefficient intervals. As shown in Fig. 6(a), the P-box of p_1 is
 434 enveloped by the CDFs of $\underline{p}_1 \sim U(-4, 2)$ and $\bar{p}_1 \sim U(-3, 3)$. Similarly, the P-box of p_2 is enveloped by the
 435 CDFs of $\underline{p}_2 \sim U(-3, 3)$ and $\bar{p}_2 \sim U(-1, 5)$ as shown in Fig. 6(b). The determination of the P-box of p_3 is
 436 more complex than the first two, since the upper or lower bound of the P-box is no longer a complete CDF
 437 curve, but a combination of multiple CDF curves as illustrated in Fig. 6(c).

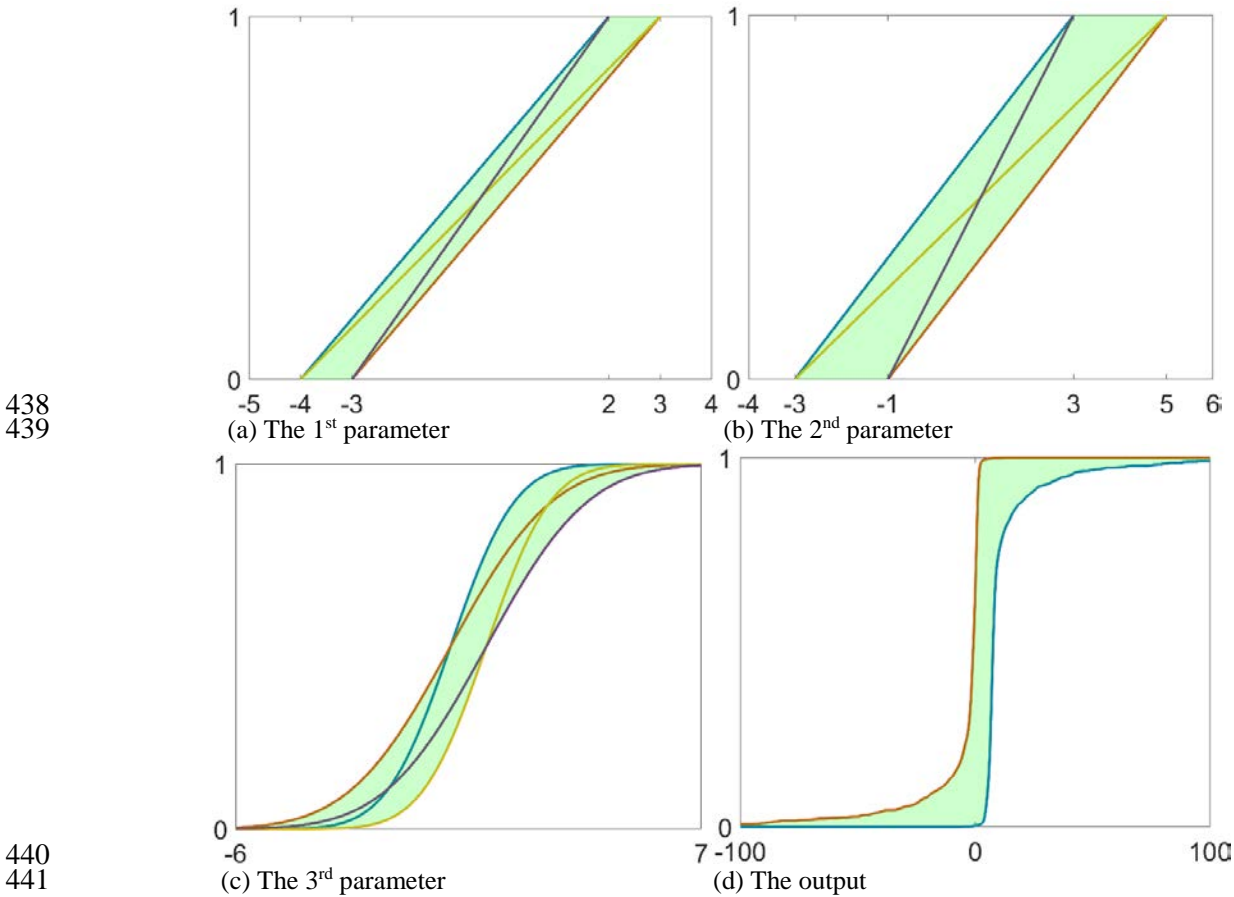


Fig. 6: The P-boxes of the inputs and output variables

443 As long as the P-boxes of the input parameters are determined, the two-level procedure can be performed
 444 to propagate the uncertainty from the inputs to the outputs. In this example, 1000 probability data points

445 $(\alpha = [\alpha_1, \alpha_2, \alpha_3])$ are sampled during the Monte Carlo simulation in the first level. For each sampled α^* ,
446 the intervals of the parameters are determined according to the P-boxes illustrated in Fig. 6(a-c). In fact, for
447 the first two parameters (p_1 and p_2), different Monte Carlo samples correspond to the intervals with
448 identical range. For the last parameter (p_3), the range of the interval is changing according to different
449 Monte Carlo samples, because of the curvilinear bounds of the P-box as shown in Fig. 6(c).

450 After the intervals of the three parameters are determined, an optimisation problem is solved in the
451 second level to find the minimal and maximal values of the output. Since the Monte Carlo sample size is
452 1000, both the minimal and maximal output sample sets contain 1000 data points. The distributions of the
453 minimal and maximal output samples are estimated using the Kernel Density Estimation (KDE) technique.
454 The CDFs of the minimal and maximal output are utilised to envelop the P-box of the output as illustrated
455 in Fig. 6(d). As described in Section 4.1, the Bhattacharyya distance between the minimal samples and the
456 maximal samples is proposed to provide a quantitative measure of the P-box (i.e. uncertainty space) of the
457 output. During the calculation of the Bhattacharyya value, the number of bins N_{bin} is taken as 100, and the
458 calculated distance value is 2.01.

459 5.3 Sensitivity indices based on the Bhattacharyya distance

460 After the output uncertainty space has been quantified by the Bhattacharyya distance, the following step
461 focuses on how the uncertainty space of the output can be changed when the epistemic uncertainty of a
462 single input parameter is reduced. Taking p_1 for example, its epistemic uncertainty is controlled by two
463 epistemic coefficients, i.e. a_1 and b_1 in Table 1. Ten levels of these two coefficients are investigated by
464 assigning ten equidistant values within their intervals, respectively. The full factorial design results in 100
465 configurations of a_1 and b_1 , i.e. $N_{sp}=100$ specific CDF curves within the P-box of p_1 . Because of the
466 simplicity of each optimisation as explained in Section 4.4, the typical interior point method is sufficient to
467 solve the problem. For a complete analysis process of the 1st parameter p_1 , the calculation time is 747.12
468 s using a desk-top computer with four processors. When the parallel computation is performed at a small
469 scale computer cluster with 36 processors, the calculation time is reduced to 183.60 s.

470 The two-level procedure is executed for each single curve of p_1 meanwhile keeping the full P-boxes of
471 p_2 and p_3 . 100 reduced P-boxes of the output are obtained, and accordingly a sample with 100
472 Bhattacharyya distance values is available. The proportional, variance-based, and comprehensive
473 sensitivity indices of p_1 are evaluated according to the equations in Section 4.2, and the bootstrap method
474 is utilised to estimate the means and standard deviations of these three indices, respectively. The number
475 of bootstrap samples utilised in this example is $N_{bs}=1000$. The same strategy is performed for p_2 and p_3 ,
476 and the ranking results according to the means of the bootstrap samples are presented in Table 2, where the
477 superscript denotes the standard deviation of the bootstrap sample. All of the standard deviation values are
478 dramatically smaller than the corresponding mean values, implying the variation of the index is low, and
479 thus the ranking results according to the means of the bootstrap samples are precise.

480 Clearly, different indices lead to different ranking results. According to the proportional index, p_2 is the
481 most significant one, implying p_2 has the largest influence in reducing the output uncertainty space.

482 According to the variance-based index, p_1 is the most significant one, implying different realisation of p_1
 483 within its epistemic uncertainty space have the largest influence on the variation of the output uncertainty
 484 space. The comprehensive index integrating the effects of the first two indicates the significance ranking
 485 as $p_2 > p_1 > p_3$. Note that, the significance ranking can be easily changed by the prescribed intervals of
 486 the epistemic coefficients in Table 1. The case study is presented to provide a tutorial demonstration of the
 487 overall stochastic SA approach. Hence, the ranking results in Table 2 are not necessarily to be the same as
 488 results in other literature about the Ishigami function.

489 Table 2: Ranking results of the parameters of the Ishigami function

Rank	Results according to different indices		
	$S_i^{(pro)}$	$S_i^{(var)}$	$S_i^{(com)}$
1	$p_2 (0.9125)^{(0.0017)}$	$p_1 (0.2957)^{(0.0278)}$	$p_2 (0.1642)^{(0.0107)}$
2	$p_1 (0.5027)^{(0.0145)}$	$p_3 (0.2891)^{(0.0177)}$	$p_1 (0.1490)^{(0.0177)}$
3	$p_3 (0.4878)^{(0.0149)}$	$p_2 (0.1800)^{(0.0117)}$	$p_3 (0.1410)^{(0.0098)}$

490

491 6 Case study II: The NASA UQ challenge problem

492 6.1 Problem description

493 Released in 2014, the NASA UQ challenge [23] has been developed as a benchmark problem for
 494 uncertainty treatment techniques, containing multiple sub-problems such as uncertainty characterization,
 495 SA, uncertainty propagation, and robust design. In this case study, the SA sub-problem is solved to
 496 demonstrate performance of the Bhattacharyya distance metric within the proposed stochastic SA approach.
 497 Fig. 7 presents the formulation of the SA problem, including the uncertain parameters \mathbf{p} , outputs \mathbf{x} , and
 498 multiple simulators $h_i(\mathbf{p})$ given as black-box models. This system contains 21 uncertain parameters, i.e. 4
 499 Category II parameters (with only epistemic uncertainty), 4 Category III parameters (with only aleatory
 500 uncertainty), and 13 Category IV parameters (with both epistemic and aleatory uncertainties), whose
 501 uncertain characteristics are listed in Table 3. The outputs $\mathbf{x} \in \mathbb{R}^5$ are evaluated by the multiple black-box
 502 models h :

$$503 \begin{cases} x_1 = h_1(p_i) & i = 1, \dots, 5; \\ x_2 = h_2(p_i) & i = 6, \dots, 10; \\ x_3 = h_3(p_i) & i = 11, \dots, 15; \\ x_4 = h_4(p_i) & i = 16, \dots, 20; \\ x_5 = p_{21}. \end{cases} \quad (10)$$

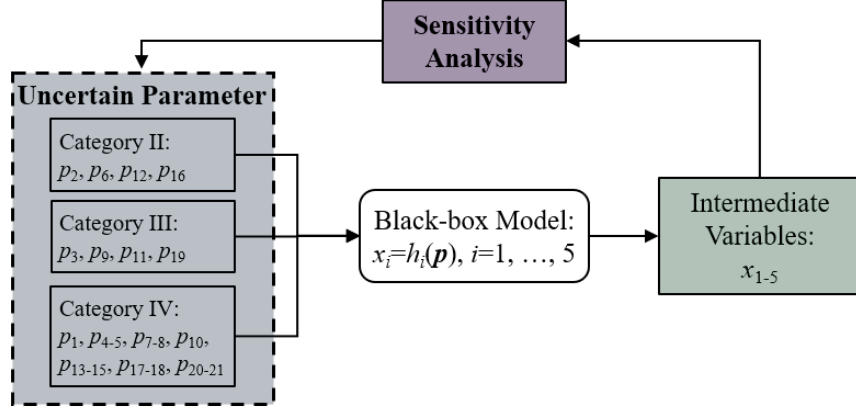


Fig. 7: The formulation and tasks of the SA problem in the NASA UQ challenge

Table 3: The uncertainty characteristics of the 21 input parameters

Sub-models	Parameter	Category	Distribution	Uncertainty characteristics
h_1	p_1	IV	Beta	$\mu_1 \in [0.6, 0.8], \sigma_1^2 \in [0.02, 0.04]$
	p_2	II	Constant	$p_2 \in [0.0, 1.0]$
	p_3	III	Uniform	$a_3 = 0, b_3 = 1$
	p_4, p_5	IV	Multiple Gaussian	$\mu_i \in [-5.0, 5.0], \sigma_i^2 \in [0.0025, 4.0]$ $\rho \in [-1.0, 1.0], i = 4, 5$
h_2	p_6	II	Constant	$p_6 \in [0.0, 1.0]$
	p_7	IV	Beta	$a_7 \in [0.982, 3.537], b_7 \in [0.619, 1.080]$
	p_8	IV	Beta	$a_8 \in [7.450, 14.093], b_8 \in [4.285, 7.864]$
	p_9	III	Uniform	$a_9 = 0, b_9 = 1$
	p_{10}	IV	Beta	$a_{10} \in [1.520, 4.513], b_{10} \in [1.536, 4.750]$
h_3	p_{11}	III	Uniform	$a_{11} = 0, b_{11} = 1$
	p_{12}	II	Constant	$p_{12} \in [0.0, 1.0]$
	p_{13}	IV	Beta	$a_{13} \in [0.412, 0.737], b_{13} \in [1.000, 2.068]$
	p_{14}	IV	Beta	$a_{14} \in [0.931, 2.169], b_{14} \in [1.000, 2.407]$
	p_{15}	IV	Beta	$a_{15} \in [5.435, 7.095], b_{15} \in [5.287, 6.945]$
h_4	p_{16}	II	Constant	$p_{16} \in [0.0, 1.0]$
	p_{17}	IV	Beta	$a_{17} \in [1.060, 1.662], b_{17} \in [1.000, 1.488]$
	p_{18}	IV	Beta	$a_{18} \in [1.000, 4.266], b_{18} \in [0.553, 1.000]$
	p_{19}	III	Uniform	$a_{19} = 0, b_{19} = 1$
	p_{20}	IV	Beta	$a_{20} \in [7.530, 13.492], b_{20} \in [4.711, 8.148]$
--	p_{21}	IV	Beta	$a_{21} \in [0.421, 1.000], b_{21} \in [7.772, 29.621]$

In the background of uncertainty analysis, the stochastic SA in this example has the task as: For the sub-model h_1 , rank the 4 Category II and IV parameters (i.e. p_1, p_2, p_4 , and p_5 .) according to their significance on the uncertainty space of x_1 , when the epistemic uncertainty of the parameter is reduced. And do the same for other sub-models h_2, h_3 , and h_4 .

513 Differing from the classical SA investigating the variation of unknown-but-fixed constants, the problem
 514 herein is established upon the reduction or variation of the uncertainty space of the imprecise random
 515 variables. This task involves multiple uncertain parameter categories and multiple distribution formats,
 516 which provide a veritable challenge for SA in the uncertainty quantification background. In order to assess
 517 the results, three published works on the NASA UQ challenge problem, namely the papers by Patelli et al.
 518 [7], Pedroni et al. [24], and McFarland [25], are used as references to compare with the results of the current
 519 work.

520 6.2 Ranking result of $p_1, p_2, p_4,$ and p_5 according to x_1

521 The SA results for the first output x_1 are shown in Table 4. The optimisation algorithm used in this
 522 example is still the interior point method. The calculation time of a complete analysis process for p_1 at a
 523 four-processor desk-top computer is 4342.08 s. When the parallel computation method is employed at a
 524 computer cluster with 36 processors, the calculation time is dramatically reduced to 312.07 s. The 2nd-4th
 525 columns list the results from different references, whose ranking orders are not exactly the same. All these
 526 three works get same conclusion that p_1 and p_5 are the first two most important parameters. However, the
 527 rankings about p_4 and p_2 are different. This demonstrates that different approaches using different
 528 quantification indices may get different results.

529 Table 4: The parameter rankings of $p_1, p_2, p_4,$ and p_5 according to x_1

Rank	Results in the references			Results in the current work		
	Patelli [7]	Pedroni [24]	McFarland [25]	$S_i^{(pro)}$	$S_i^{(var)}$	$S_i^{(com)}$
1	p_1	p_1	p_1	p_1 (0.715)	p_2 (0.215)	p_1 (0.137)
2	p_5	p_5	p_5	p_5 (0.199)	p_1 (0.192)	p_5 (0.036)
3	p_4	p_4	p_2	p_4 (0.166)	p_5 (0.182)	p_4 (0.025)
4	p_2	p_2	p_4	p_2 (0.099)	p_4 (0.154)	p_2 (0.021)

530
 531 The last three columns of Table 4 present the ranking results of the current work according to three
 532 different indices, namely the proportional index $S_i^{(pro)}$, the variance-based index $S_i^{(var)}$, and the
 533 comprehensive index $S_i^{(com)}$. The index value is provided in parenthesis after each parameter in the table.
 534 The ranking result according to the proportional index is the same as the results from the first two references.
 535 However, the ranking result according to the variance-based index is different from all of the three results
 536 in the references. Nevertheless, the variance-based index values of these four parameters are similar with
 537 each other, implying the four parameters have comparable influence on the dispersion degree of the output
 538 uncertainty space. This is also the reason that the ranking based on the comprehensive index is the same as
 539 the one according to the proportional index.

540
 541

542 6.3 Ranking result of $p_6, p_7, p_8,$ and p_{10} according to x_2

543 Sensitivity results for x_2 are more complex than that for x_1 , since different referenced works obtain quite
 544 different parameter rankings, as shown in Table 5. The results in the references are different in ranking the
 545 last two parameters (p_8 and p_{10}). In the work of McFarland [25], p_8 and p_{10} were regarded to be negligible,
 546 and thus they were not included in McFarland’s ranking result.

547 In Table 5, the differences are also detected among the results according to the three indices in the
 548 current work. Using the proportional index, p_7 is demonstrated as the most significant parameter. However,
 549 in the ranking according to the variance-based index, p_6 is the most significant parameter. By integrally
 550 considering both reduction and dispersion effects, the comprehensive index provides the significant ranking
 551 as: $p_7 > p_6 > p_8 > p_{10}$. Note that this index shows p_7 is the most significant parameter, while in the
 552 referenced works, p_6 was regarded as the most significant one.

553 Table 5: The parameter rankings of $p_6, p_7, p_8,$ and p_{10} according to x_2

Rank	Results in the references			Results in the current work		
	Patelli [7]	Pedroni [24]	McFarland [25]	$S_i^{(pro)}$	$S_i^{(var)}$	$S_i^{(com)}$
1	p_6	p_6	p_6	p_7 (0.643)	p_6 (0.187)	p_7 (0.101)
2	p_7	p_7	p_7	p_6 (0.488)	p_7 (0.157)	p_6 (0.091)
3	p_8	p_{10}	--	p_8 (0.483)	p_{10} (0.144)	P_8 (0.070)
4	p_{10}	p_8	--	p_{10} (0.458)	p_8 (0.138)	P_{10} (0.063)

554
 555 To explain this difference between the current results and the referenced results, a further investigation
 556 is performed for p_6 and p_7 . Fig. 8 illustrates the samples of the Bhattacharyya distances of the output
 557 uncertainty spaces, when the epistemic uncertainties of p_6 and p_7 are reduced, respectively. Recall the SA
 558 framework in Fig. 5, the P-box of a parameter is reduced into a single CDF curve, meaning its epistemic
 559 uncertainty is completely reduced, and subsequently the reduction degree of the output P-box is investigated.
 560 However, not only a single specified CDF, but a series of CDFs of this parameter is utilised, leading to a
 561 series of changed P-boxes of the output. In this example, the number of the investigating CDFs is 100. As
 562 shown in Table 3, p_6 is a Category II parameter within the interval [0.0, 0.1], and thus 100 values equi-
 563 spaced along this interval are assigned. Differently, p_7 is a Category IV parameter following imprecise Beta
 564 distribution, whose epistemic uncertainty is driven by $a_7 \in [0.982, 3.537]$ and $b_7 \in [0.619, 1.080]$.
 565 Hence 10 equidistant values within each interval of a_7 and b_7 are respectively assigned. And the full
 566 factorial design is performed to specify 100 Beta distributions of p_6 . After these 100 P-boxes of the output
 567 are obtained, the Bhattacharyya distance is utilised to quantify the P-boxes, and thus the histograms for p_6
 568 and p_7 in Fig. 8 are plotted based on 100 Bhattacharyya distance samples, respectively.

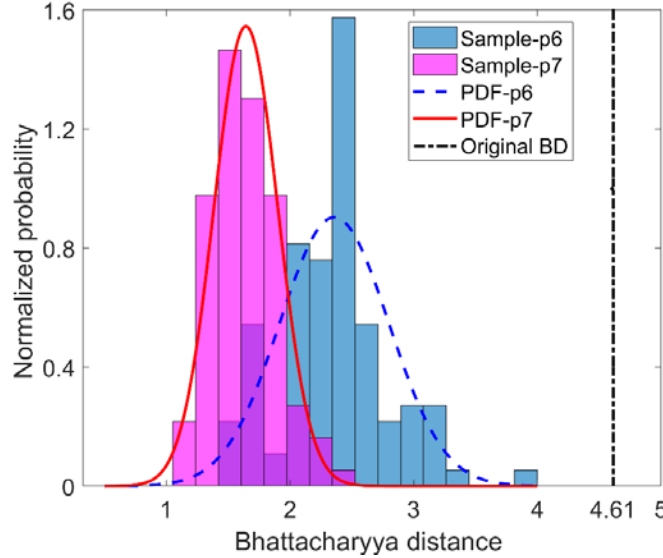


Fig. 8: Comparison of the samples of Bhattacharyya distances for p_6 and p_7

569
570

571 In Fig. 8, the Bhattacharyya distance sample according to p_6 is denominated as “Sample-p6”, and the
572 one according to p_7 is denominated as “Sample-p7”. Based on these two samples, two normal distributions
573 are estimated, whose Probability Density Functions (PDFs) are also illustrated in the figure. Note that, the
574 samples are not necessarily following normal distribution. The normal PDFs are presented herein to assist
575 the comparison between these samples. Another important element in Fig. 8 is the dash-dotted vertical line
576 (denominated as “Original BD), denoting the Bhattacharyya distance value of the original P-box of the
577 output x_2 . In other words, when every parameter keeps its epistemic uncertainty, the original output
578 uncertainty space has the Bhattacharyya distance shown by the dash-dotted vertical line (i.e. 4.61). Now
579 Fig. 8 can be investigated from two aspects: the reduction and dispersion of the samples.

580

581 1) The reduction of uncertainty spaces is reflected by the distance from the samples to the dash-dotted
582 line. The distance from Sample-p7 to Original BD is larger than that from Sample-p6 to Original BD,
583 implying the output uncertainty space is much more reduced by reducing the epistemic uncertainty of
584 p_7 , rather than p_6 . This coincides with the ranking result with the proportional index in the 5th column
585 of Table 5, where p_7 is more significant than p_6 .

585

586 2) The dispersion of the uncertainty space is reflected by the variances of the two samples. Sample-p7 (or
587 PDF-p7) is more centralized than Sample-p6 (or PDF-p6), implying various specified CDF of p_6 have
588 more significant influence on the output uncertainty space. This is why, using the variance-based index,
589 p_6 is demonstrated to be more important than p_7 as shown in the 6th column of Table 5.

589

590 The above investigation on reduction and dispersion characteristics of the output uncertainty space
591 provides a clear understanding on the significance of p_6 and p_7 , from different aspects. Consequently, it is
592 suggested to propose the comprehensive index to combine both aspects of the influence and provide an
593 integrated evaluation of the parameters.

593

594 The parameter ranking according to x_3 and x_4 are listed in Tables 6 and 7, respectively. The results from
595 different references, as well as the results from the current work according to different indices, exhibit more
or less differences. However, the situation is the same as the above discussed results for x_1 and x_2 , and hence

596 the same investigation procedure can be repeated to analyse the result for x_3 and x_4 . For clarity, the detailed
 597 analysis for x_3 and x_4 is omitted.

598 Table 6: The parameter rankings of p_{12} , p_{13} , p_{14} , and p_{15} according to x_3

Rank	Results in the references			Results in the current work		
	Patelli [7]	Pedroni [24]	McFarland [25]	$S_i^{(pro)}$	$S_i^{(var)}$	$S_i^{(com)}$
1	p_{12}	p_{12}	p_{12}	p_{12} (0.944)	p_{12} (1.129)	p_{12} (1.066)
2	p_{15}	p_{13}, p_{15}	--	p_{14} (0.142)	p_{14} (0.215)	p_{14} (0.030)
3	p_{14}	--	--	p_{13} (0.058)	p_{13} (0.085)	p_{13} (0.005)
4	p_{13}	p_{14}	--	p_{15} (0.458)	p_{15} (0.068)	p_{15} (0.002)

599
 600 Table 7: The parameter rankings of p_{16} , p_{17} , p_{18} , and p_{20} according to x_4

Rank	Results in the references			Results in the current work		
	Patelli [7]	Pedroni [24]	McFarland [25]	$S_i^{(pro)}$	$S_i^{(var)}$	$S_i^{(com)}$
1	p_{16}	p_{16}	p_{16}	p_{16} (0.737)	p_{18} (0.355)	p_{16} (0.196)
2	p_{18}	p_{18}	p_{18}	p_{18} (0.445)	p_{16} (0.265)	p_{18} (0.158)
3	p_{17}	p_{17}	--	p_{17} (0.180)	p_{17} (0.185)	p_{17} (0.033)
4	p_{20}	p_{20}	--	p_{20} (0.019)	p_{20} (0.143)	p_{20} (0.003)

601
 602 **7 Conclusions and perspectives**

603 This work promotes the application of the Bhattacharyya distance as a novel UQ metric in a two-level
 604 stochastic SA framework. The two-level procedure provides a clear logic for investigation of the P-box
 605 representation, where the aleatory uncertainty is quantified through the Monte Carlo simulation in the first
 606 level, and the epistemic uncertainty is propagated from inputs to outputs via optimisation in the second
 607 level. The Bhattacharyya distance plays a critical role in constructing the sensitivity indices by providing a
 608 quantitative measure of the output uncertainty space. All of the three sensitivity indices are capable of
 609 measuring the significance of both Categories II and IV parameters with not only epistemic uncertainty but
 610 also aleatory uncertainty, implying this approach is an extension of the deterministic SA which is only
 611 feasible for interval sensitivities of the unknown-but-fixed constants (i.e. Category II parameters). The three
 612 proposed indices explores different aspects of the sensitivity information of the parameters, that is to say,
 613 1) the proportional index focuses on the reduction effect of the output uncertainty space; 2) the variance-
 614 based index focuses on the dispersion degree of the output uncertainty space; and 3) the comprehensive
 615 index measures both reduction and dispersion effects of the output uncertainty space. The different
 616 information conveyed by different indices is significant for decision makers to achieve a better
 617 understanding of the influence of inputs on the outputs from different perspectives.

618 The tutorial case study utilising the Ishigami function provides a detailed demonstration of the overall
 619 approach with reproducible results. In the second case study with the NASA UQ challenge problem, the
 620 comparison among the references and current work using different indices exhibits the uniform ranking

621 tendency, demonstrating the feasibility of the proposed approach. The detailed differences among the
622 ranking results (even in the references themselves) confirms the natural conclusion that different approaches
623 using different sensitivity indices result in different rankings. However, this differences do not necessarily
624 conceal the role of the Bhattacharyya distance as an elaborate, quantitative, and uniform metric for
625 stochastic SA in the background of uncertainty quantification and propagation.

626 The current work is based on the first-order SA with the assumption that the investigating parameters
627 are independent with each other. Consequently, one of the extensions of this work is a second-order SA
628 investigating the interactions between uncertain parameters. Furthermore, the proposed approach has the
629 potential to be utilised in the case of tail distributions in the context of reliability analysis, with the necessary
630 extension using the directional Monte Carlo sampling techniques, e.g. the subset simulation and Markov
631 chain Monte Carlo algorithm, to achieve a high efficiency for the rare occurrence events.

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