

1 **Cues to the sex ratio of the local population influence women's preferences for**
2 **facial symmetry**

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4 Christopher D Watkins¹, Benedict C Jones¹, Anthony C Little², Lisa M DeBruine¹ &
5 David R Feinberg³

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7 ¹ Face Research Laboratory, School of Psychology, University of Aberdeen,
8 Aberdeen, AB24 3FX, Scotland, UK.

9 ² School of Natural Sciences, University of Stirling, Stirling, FK9 4LA, Scotland, UK.

10 ³ Department of Psychology, Neuroscience and Behaviour, McMaster University,
11 Hamilton, Ontario, L8S4L8, Canada.

12

13 **Corresponding authors**

14 Christopher D Watkins, University of Aberdeen, Scotland, UK.

15 Tel: +44 (0)1224 273933

16 Fax: +44 (0)1224 273426

17 Email: r01cdw9@abdn.ac.uk

18

19 Benedict C Jones, University of Aberdeen, Scotland, UK.

20 Tel: +44 (0)1224 273933

21 Fax: +44 (0)1224 273426

22 Email: ben.jones@abdn.ac.uk

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26 **Abstract**

27 In non-human species, increasing the proportion of potential mates in the
28 local population often increases preferences for high quality mates, while increasing
29 the proportion of potential competitors for mates intensifies within-sex competition. In
30 two experiments, we tested for analogous effects in humans by manipulating pictorial
31 cues to the sex ratio of the local population and assessing women's preferences for
32 facial symmetry, a putative cue of mate quality in humans. In both experiments,
33 viewing slideshows with varied sex ratios tended to increase preferences for
34 symmetry in the sex that was depicted as being in the majority and tended to
35 decrease preferences for symmetry in the sex that was depicted as being in the
36 minority. In other words, increasing the apparent proportion of a given sex in the
37 local population increased the salience of facial cues of quality in that sex, which
38 may support adaptive appraisals of both potential mates' and competitors' quality.
39 This effect of sex ratio was independent of (i.e. did not interact with) an effect of cues
40 to the degree of variation in the attractiveness of individuals in the local population,
41 whereby the degree of variation in men's, but not women's, attractiveness modulated
42 symmetry preferences. These findings demonstrate that symmetry preferences in
43 humans are influenced by cues to the sex ratio of the local population in ways that
44 complement both the facultative responses that have been observed in many other
45 species and theories of both inter-sexual and intra-sexual selection.

46

47 **Key words:** sex ratio, mate preference, within-sex competition, dominance,
48 attraction, fluctuating asymmetry

49

50 **Introduction**

51 Sexual selection can result in traits that advertise aspects of mate quality in
52 one sex and preferences for such traits in the other sex (Andersson 1994). Variation
53 in the characteristics and demands of the local environment can cause systematic
54 variation in preferences for cues of mate quality, however (Gangestad & Buss 1993;
55 Jennions & Petrie 1997; Penton-Voak et al. 2004; Little et al. 2007a; DeBruine et al.
56 2010, 2011). One ecological constraint that may be particularly important for
57 variation in mate preferences is the ratio of potential mates to potential competitors
58 for mates within the local population (i.e. the operational sex ratio, Emlen & Oring
59 1977; Guttentag & Secord 1983). The potential costs of competing for high quality
60 mates are reduced considerably when potential mates are plentiful and competitors
61 for mates are relatively scarce (Pedersen 1991; see also Noe & Hammerstein 1994
62 and Kvarnemo & Ahnesjo 1996). This change in costs may, in turn, allow individuals
63 to increase their preferences for cues that are associated with high quality in
64 potential mates and require them to engage in less within-sex competition in order to
65 secure mates (Pedersen 1991; see also Noe & Hammerstein 1994 and Kvarnemo &
66 Ahnesjo 1996).

67

68 Changing preferences as a result of experimentally manipulating the
69 operational sex ratio of the local population have been reported for several different
70 species. For example, female guppies show stronger preferences for attractive male
71 colour characteristics when the sex ratio is biased towards males than when it is
72 biased towards females (Jirotkul 1999). In field crickets, female mate preferences
73 also show greater selectivity when the sex ratio is biased towards males (Souroukis
74 & Murray 1994). Additionally, greater selectivity in male mate preferences when the

75 sex ratio is biased towards females have been observed in snapping shrimp
76 (Mathews 2002), milkweed beetles (Lawrence 1986) and pipefish (Berglund 1994).
77 Other experiments have shown that the intensity of within-sex competition is also
78 affected by altering the sex ratio. For example, biasing the sex ratio towards own-sex
79 individuals causes more intense within-sex competition in Japanese medaka (Clark
80 & Grant 2010), guppies (Jirotkul 1999), red spotted newt (Verrell 1983) and
81 amphipods (Dick & Elwood 1996). Importantly, these effects of experimentally
82 altered sex ratio complement findings from studies that investigated relationships
83 between naturally occurring variation in sex ratios and indices of either mate
84 preferences or within-sex competition. This work includes studies of various fish
85 species (e.g. Balshine-Earn 1996; Forsgren et al. 2004), ungulates (e.g. Clutton-
86 Brock et al. 1997; Coltman et al. 1999), rodents (Michener & Locklear 1990), birds
87 (e.g. Colwell & Oring 1988) and primates (Wang et al. 2009; see also Hohmann &
88 Fruth 2003).

89

90 While the findings described above indicate that sex ratio can influence mate
91 preferences and within-sex competition in many non-human species, correlational
92 studies of naturally occurring variation in human sex ratios suggest that they may
93 also be important determinants of human behaviour. For example, Pollet and Nettle
94 (2008) reported that women in regions of the US with male-biased sex ratios
95 demonstrated stronger preferences for high socioeconomic status (i.e. attractive,
96 Hume & Montgomerie 2001) men than did women in regions of the US with female-
97 biased sex ratios. Additionally, female-female competition appears to be more
98 intense in societies with female-biased sex ratios than it is in those with male-biased
99 sex ratios (Schuster 1983, 1985; Campbell 1995). Similarly, male-male competition

100 appears to be more intense in societies with male-biased sex ratios than it is in those
101 with female-biased sex ratios (e.g. Hudson & Den Boer 2002, 2004). These
102 correlations raise the intriguing possibility that manipulating cues to the sex ratio of
103 the local population may influence mate preferences and within-sex competition in
104 humans. However, we know of no experimental studies examining this issue.

105

106 Symmetry is an important cue of mate quality in many species, including
107 humans (reviewed in Gangestad & Simpson 2000). For example, in humans, facial
108 symmetry is positively correlated with indices of long-term, medical health (e.g.
109 Thornhill & Gangestad 2006; Lie et al. 2008), other putative health cues (e.g.
110 Gangestad & Thornhill 2003; Jones et al. 2004; Little et al. 2008) and attractiveness
111 (Grammer & Thornhill 1994; Penton-Voak et al. 2001; Jones et al. 2001, 2004), and
112 is negatively correlated with exposure to developmental stressors (Özener 2010;
113 Özener & Fink 2010). Moreover, experimentally increasing the symmetry of digital
114 face images tends to increase their attractiveness (e.g. Perrett et al. 1999; Little &
115 Jones 2003, 2006). Collectively, these findings suggest that symmetry is a cue to the
116 quality of both potential mates and potential competitors for mates. If manipulating
117 cues to the sex ratio of the local population alters both behaviours that might function
118 to promote successful within-sex competition and the selectivity of mate preferences
119 (see earlier discussion), one might expect altering cues to the sex ratio of the local
120 population to influence responses to symmetric individuals of both sexes.

121

122 Many researchers have suggested that preferences for symmetric mates may
123 reflect adaptive responses that function, at least in part, to increase reproductive
124 fitness by encouraging mating with high quality individuals (e.g. Grammer & Thornhill

125 1994; Penton-Voak et al. 2001; Gangestad & Thornhill 2003). Other researchers
126 have also emphasised that, because of the significant threat attractive women pose
127 as competitors, cues to women's mate quality may be highly salient to other women
128 and play a critical role in within-sex competition for mates (e.g. Buss & Dedden 1990;
129 Maner et al. 2003; Fisher & Cox 2009; Puts et al. 2011). For example, perceptions of
130 potential competitors' mate quality may be important to both gauge one's own
131 market value and to identify competitors for mates who one could or could not
132 successfully compete with for mates (e.g. Buss & Dedden 1990; Maner et al. 2003;
133 Fisher & Cox 2009; Puts et al. 2011). These observations raise the possibility that
134 preferences for symmetry in the faces of a given sex may be stronger when there is
135 a greater proportion of that sex in the local population than when there is a lower
136 proportion of that sex in the local population. Such facultative responses to cues to
137 the sex ratio of the local population would be adaptive if increased attraction to high
138 quality mates when there is a high proportion of potential mates supported efficient
139 allocation of mating effort. They would also be adaptive if increased salience of
140 quality cues in potential competitors when there is a high proportion of potential
141 competitors supported successful within-sex competition for mates.

142

143 In light of the above, in Experiment 1, we examined women's preferences for
144 symmetry in potential mates' and competitors' faces after pictorial cues to the sex
145 ratio of the local population were experimentally manipulated. To do this, we first
146 assessed women's preferences for men's and women's facial symmetry in an initial
147 (i.e. baseline) test. Immediately after this initial test, women watched one of two
148 slideshows, both of which consisted of images of men's and women's faces. In one
149 of these slideshows, the majority of images depicted men (i.e. potential mates). In

150 the other slideshow, the majority of images depicted women (i.e. potential
151 competitors). After watching the slideshow, women repeated the initial symmetry
152 preference test. We predicted that if attractiveness judgements of symmetric versus
153 asymmetric faces are sensitive to cues to the sex ratio of the local population,
154 preferences for symmetry in the faces of a given sex may be stronger when there is
155 a greater proportion of that sex in the local population than when there is a lower
156 proportion of that sex in the local population.

157

158 Some researchers have suggested that the effects of sex ratio on mating-
159 related behaviours may be more closely linked to the degree of variation in the mate
160 quality of individuals in the local population than to the sex ratio, *per se* (e.g. Owens
161 & Thompson 1994; Kvarnemo & Simmons 1999). Thus, in Experiment 2, we adapted
162 the procedure from Experiment 1 in order to investigate whether cues to the sex ratio
163 of the local population and cues to the degree of variation in attractiveness in the
164 local population have independent or interacting effects on women's symmetry
165 preferences. To do this, we added an additional factor (*attractiveness variation*) to
166 our experimental design. Thus, we assessed women's symmetry preferences before
167 and after viewing slideshows of either male faces that varied greatly in
168 attractiveness, male faces that varied little in attractiveness, female faces that varied
169 greatly in attractiveness, or female faces that varied little in attractiveness.
170 Importantly, all slideshows had the same mean facial attractiveness. If cues to the
171 sex ratio of the local population influence behaviour independently of cues to the
172 attractiveness variation in the local population, then we should see independent
173 effects of the sex ratio depicted in the slideshows and of variation in attractiveness.

174

175 **Experiment 1**

176 In Experiment 1, we tested if preferences for symmetry in the faces of a given
177 sex are stronger when there is a greater proportion of that sex in the local population
178 than when there is a lower proportion of that sex in the local population.

179

180 **Methods**

181 ***Participants***

182 One hundred heterosexual women (mean age = 22.94 years, SD = 6.76
183 years) participated in Experiment 1. Participants were recruited for an online study of
184 face preferences by following links from social bookmarking websites (e.g.
185 stumbleupon). Prior research has established that lab-based and online studies of
186 women's face preferences produce very similar patterns of results (e.g. Jones et al.
187 2005; Conway et al. 2008).

188

189 ***Stimuli***

190 To assess symmetry preferences in both the initial (i.e. baseline), pre-
191 slideshow test and the final, post-slideshow test, we manufactured 20 pairs of faces
192 (10 male pairs and 10 female pairs), each pair consisting of a symmetrised and
193 original (i.e. relatively asymmetric) version of an individual (see Figure 1 for
194 examples). First, full-colour digital face photographs were taken of 10 white young
195 adult men and 10 white young adult women with neutral expressions and direct
196 gaze. Photographs were taken against a constant background and photographic
197 conditions were standardised. Next, using procedures first described by Perrett et al.
198 (1999), we used computer graphic methods to warp each image into a more
199 symmetric shape. This method for manipulating symmetry of face shape in digital

200 images does not affect other aspects of facial appearance (e.g. identity, skin colour
201 and texture, aspects of face shape other than symmetry, Perrett et al. 1999) and has
202 been used to assess symmetry preferences in many other previous studies (e.g.
203 Jones et al. 2001; Little & Jones 2003, 2006; Little et al. 2007b, 2011).

204

205 INSERT FIGURE 1 AROUND HERE

206

207 We used full-colour images of 25 white, young adult men and 25 white, young
208 adult women for the slideshows. Each individual was photographed with a neutral
209 expression and direct gaze against a constant background and under standardised
210 lighting conditions. None of the individuals whose photographs were used for the
211 slideshow were used to manufacture the symmetry preference stimuli. To provide
212 descriptive statistics for the attractiveness of these images, all 50 images were rated
213 for attractiveness by 23 women using a 1 (very unattractive) to 7 (very attractive)
214 scale. Inter-rater agreement for these attractiveness ratings was high for both men's
215 faces (Cronbach's alpha = 0.96) and women's faces (Cronbach's alpha = 0.94). The
216 mean rating for men's faces was 2.36 (variance = 0.24) and for women's faces was
217 2.69 (variance = 0.36).

218

219 ***Procedure***

220 In the first part of the experiment (the baseline, pre-slideshow test), the
221 participants were shown the 20 pairs of face images (10 pairs of men's faces and 10
222 pairs of women's faces), each pair consisting of a symmetrised and relatively
223 asymmetric (i.e. original) version of an individual, and were instructed to indicate
224 which face in each pair they thought was more attractive. Trial order and the side of

225 the screen on which any given image was presented were both fully randomised.
226 The inclusion of this baseline test is potentially important; it allows us to control for
227 possible pre-existing individual differences in women's preferences for symmetric
228 faces, such as those related to women's own attractiveness (Little et al. 2001),
229 sociosexuality (Sacco et al. 2009; Quist et al. 2011), menstrual cycle phase (Little et
230 al. 2007b), perceived vulnerability to disease (Young et al. 2011), or recent exposure
231 to pathogen cues (Little et al. 2011).

232

233 Immediately after completing the pre-slideshow test, the participants
234 completed the second part of the experiment (the slideshow). In this second part,
235 participants were randomly allocated to one of two conditions in which they were
236 instructed to watch a slideshow of thirty face images (none of which were seen in the
237 pre-slideshow test), where each face image was shown for 2000ms. In one
238 condition, participants watched a slideshow in which 83% (i.e. 25/30) of the images
239 depicted men and 17% (i.e. 5/30) of the images in this slideshow were of women. In
240 the other condition, participants watched a slideshow in which 83% (i.e. 25/30) of the
241 images depicted women and 17% (i.e. 5/30) of the images in this slideshow were of
242 men. In common with other research employing similar observation phases (e.g.
243 Jones et al. 2007), participants were simply instructed to watch the images closely.
244 Other studies have recently established that priming participants with pictorial cues
245 to the nature of the local environment (e.g. images associated with high or low
246 pathogen loads, Little et al. 2011) can affect face preferences in ways that are
247 consistent with findings from correlational studies of natural variation in ecological
248 conditions (e.g. DeBruine et al. 2010, 2011).

249

250 In the third and final part of the experiment (the post-slideshow test), the
251 participants repeated the initial, pre-slideshow test. This post-slideshow test was
252 completed immediately after the slideshow.

253

254 **Results**

255 For each participant, we first calculated (separately) the proportion of trials on
256 which she chose the symmetrised face as the more attractive when judging men's
257 faces in the pre-slideshow test, men's faces in the post-slideshow test, women's
258 faces in the pre-slideshow test and women's faces in the post-slideshow test. Two-
259 tailed p-values are reported for all analyses. Consistent with previous work (e.g.
260 Perrett et al. 1999; Little et al. 2003, 2006), symmetry preferences at pre-test (i.e.
261 baseline) were significantly greater than chance for both women's faces ($t_{99} = 6.16$,
262 $P < 0.001$, $M = 0.61$, $SEM = 0.02$, $d = 0.62$) and men's faces ($t_{99} = 3.00$, $P = 0.003$,
263 $M = 0.55$, $SEM = 0.02$, $d = 0.31$). Summary statistics for each individual condition are
264 shown in Table 1.

265

266 INSERT TABLE 1 AROUND HERE

267

268 We then calculated the change in preference for symmetric women between
269 the pre-slideshow test and the post-slideshow test for each participant. To do this,
270 we subtracted the proportion of trials on which symmetric female faces were chosen
271 in the pre-slideshow test from the corresponding score in the post-slideshow test.
272 Positive scores indicate preferences increased after the slideshow, while negative
273 scores indicate preferences decreased after the slideshow. Similarly, we calculated

274 the corresponding change in each participant's preference for symmetric men
275 between the pre-slideshow test and the post-slideshow test.

276

277 Next, we coded these difference scores to reflect (1) the increase in symmetry
278 preference for the sex of face that was in the *majority* during the slideshow that was
279 viewed (i.e. men's faces when the sex ratio of the slideshow was male-biased and
280 women's faces when the sex ratio of the slideshow was female-biased) and (2) the
281 increase in symmetry preference for sex of face that was in the *minority* during the
282 slideshow that was viewed (i.e. men's faces when the sex ratio of the slideshow was
283 female-biased and women's faces when the sex ratio of the slideshow was male-
284 biased), respectively. Because all participants judged both men's and women's faces
285 in the symmetry preference test, this factor was represented in our experimental
286 design by the within-subjects factor *face type*, which had two levels (sex in majority
287 during slideshow and sex in minority during slideshow). Note that whether male face
288 scores were coded as the majority condition and female face scores were coded as
289 the minority condition (or vice versa) depended entirely on whether that participant
290 had been shown the male-biased or female-biased slideshow in the slideshow phase
291 of the experiment (see below).

292

293 The other critical factor in our design reflected whether participants saw a
294 male-biased or female biased slideshow in the slideshow phase of the experiment.
295 Because each individual participant saw either a male-biased or female biased
296 slideshow (but not both), this variable was represented by the between-subjects
297 factor *slideshow type*, which also had two levels (male-biased versus female-
298 biased).

299

300 Coding the data in this way is important, as it is the only way that allows us to
301 directly compare the effects of viewing cues to a male-biased sex ratio on
302 evaluations of men's attractiveness and the effects of viewing cues to a female-
303 biased sex ratio on evaluations of women's attractiveness. This comparison may be
304 critical, since some studies have reported equivalent effects of sex ratio on mate
305 preferences and competition-related behaviours (e.g. Jirotkul 1999), while other
306 researchers have suggested that the effects of sex ratio on mate preferences may
307 be generally weaker than those that have been observed for competition-related
308 behaviours (e.g. Kvarnemo & Ahnesjo 1996).

309

310 We analysed the difference scores using a mixed design ANOVA that had the
311 within-subjects factor *face type* (sex in majority during slideshow versus sex in
312 minority during slideshow) and the between-subjects factor *slideshow type* (male-
313 biased versus female-biased). This analysis revealed the predicted main effect of
314 *face type* ($F_{1,98} = 6.52$, $P = .012$, partial $\eta^2 = 0.06$, Figure 2), whereby women
315 showed a greater increase in symmetry preference when judging the sex that was in
316 the majority during the slideshow than they did when judging the sex that was in the
317 minority during the slideshow. One-sample t-tests comparing the change in
318 symmetry preference in each condition with baseline (i.e. the chance value of zero)
319 showed that, although symmetry preferences tended to be increased from baseline
320 when judging the sex that was in the majority during the slideshow ($t_{99} = 1.56$, $P =$
321 1.12) and tended to be decreased from baseline when judging the sex that was in
322 the minority during the slideshow ($t_{99} = -1.83$, $P = 0.07$), these changes were not
323 significantly different from baseline in either case. Nonetheless, note that the

324 significant main effect of *face type* reported above showed that these scores were
325 significantly different from one another. There was no higher order interaction
326 between *face type* and *slideshow type* ($F_{1,98} = 0.05$, $P = 0.83$, partial $\eta^2 < 0.001$),
327 indicating that the significant main effect of *face type* described above was not
328 qualified by an interaction with *slideshow type*. This null finding is potentially
329 important. It shows that the magnitude of the effect of viewing a male-biased
330 slideshow on judgements of men's faces was not significantly different from the
331 magnitude of the effect of viewing a female-biased slideshow on judgements of
332 women's faces. This, in turn, suggests that cues to the nature of the sex ratio of the
333 local population can have equivalent effects on mating-related and competition-
334 related perceptions. The main effect of *slideshow type* was not significant ($F_{1,98} =$
335 1.31 , $P = 0.26$, partial $\eta^2 = 0.013$).

336

337 INSERT FIGURE 2 AROUND HERE

338

339 **Experiment 2**

340 In Experiment 1, women's preferences for symmetrised faces tended to be
341 increased when judging the attractiveness of the sex that was depicted as being in
342 the majority during the slideshow phase of the experiment, but tended to be
343 decreased when judging the attractiveness of the sex that was depicted as being in
344 the minority during the slideshow phase. These findings suggest that experimentally
345 manipulating pictorial cues to the sex ratio of the local population can alter symmetry
346 preferences. However, it has previously been suggested that effects of sex ratio on
347 mating-related behaviours in some non-human species depend on the degree of
348 variation in the mate quality of individuals in the local population, rather than (or in

349 addition to) cues to the sex ratio of the local population (e.g. Owens & Thompson
350 1994; Kvarnemo & Simmons 1999). To investigate this possibility, in Experiment 2,
351 we adapted the paradigm that we had used in Experiment 1 in order to test whether
352 the observed effect of viewing cues to the sex ratio of the local population is
353 independent of, or interacts with, the possible effects of the degree of variation in the
354 attractiveness of the images shown in the slideshows.

355

356 **Methods**

357 ***Participants***

358 One hundred heterosexual women (mean age = 24.97 years, SD = 8.74
359 years) participated in Experiment 2. As in Experiment 1, participants were recruited
360 for an online study of face preferences by following links from social bookmarking
361 websites (e.g. stumbleupon).

362

363 ***Stimuli***

364 The same stimuli that we used to assess preferences for symmetrised versus
365 original (i.e. relatively asymmetric) versions of men's and women's faces in
366 Experiment 1 were used again here. To select the faces to show in the slideshows, a
367 separate group of 50 male and 50 female face images were first rated for
368 attractiveness by 100 men and 100 women using a 1 (very unattractive) to 7 (very
369 attractive) scale. Trial order was fully randomised and none of these raters took part
370 in either of the main experiments. Inter-rater agreement was extremely high
371 (Cronbach's alphas: ratings of men's faces = 0.97, ratings of women's faces = 0.96)
372 and ratings made by men and women were highly correlated for both face sexes
373 (correlation between men's and women's ratings of men's faces: $r = 0.97$, $P < 0.001$;

374 correlation between men's and women's ratings of women's faces: $r = 0.96$, $P <$
375 0.001). Consequently, we calculated a single attractiveness score for each face by
376 averaging ratings across raters. These ratings were used to select four sets of face
377 images (15 images per set) with the following properties: male images that varied
378 greatly in attractiveness (variance = 0.42), male images that varied little in
379 attractiveness (variance = 0.04), female images that varied greatly in attractiveness
380 (variance = 0.40), and female images that varied little in attractiveness (variance =
381 0.05). These images were used in the male high variance, male low variance, female
382 high variance, and female low variance slideshows, respectively. The mean
383 attractiveness rating for each set was 2.65. In each set, the attractiveness ratings of
384 the seven least attractive images were below this mean value and the attractiveness
385 ratings of the seven most attractive images were above this mean value. The
386 remaining image in each set was the middle ranked image and was within 0.08 of
387 the mean attractiveness of the sample. All face images were full colour and were
388 shown with neutral expressions and direct gaze.

389

390 ***Procedure***

391 The procedure was identical to that in Experiment 1, except that participants
392 were randomly allocated to one of the male high variance, male low variance, female
393 high variance, or female low variance slideshow conditions and that each image was
394 shown only once (for 4000ms) in the slideshow.

395

396 **Results**

397 Responses were coded in precisely the same way as in Experiment 1. Two-
398 tailed p-values are reported for all analyses. As in Experiment 1, and consistent with

399 previous work (e.g. Perrett et al. 1999; Little et al. 2003, 2006), symmetry
400 preferences at pre-test (i.e. baseline) were significantly greater than chance for both
401 women's faces ($t_{99} = 5.39$, $P < 0.001$, $M = 0.59$, $SEM = 0.02$, $d = 0.54$) and men's
402 faces ($t_{99} = 2.08$, $P = 0.040$, $M = 0.53$, $SEM = 0.02$, $d = 0.21$). Summary statistics for
403 each individual condition are shown in Table 2.

404

405 INSERT TABLE 2 AROUND HERE

406

407 To keep terminology consistent between Experiment 1 and Experiment 2, the
408 'sex in majority' and 'sex in the minority' conditions reflect the sex seen or not seen
409 in the slideshow, respectively. As in Experiment 1, difference scores were first
410 analysed using a mixed design ANOVA with the within-subjects factor *face type* (sex
411 in majority during slideshow versus sex in minority during slideshow), the between-
412 subjects factor *slideshow type* (male-biased versus female-biased), and the
413 additional between-subjects factor *attractiveness variation* (high versus low). As in
414 Experiment 1, this analysis revealed the predicted main effect of *face type* ($F_{1,96} =$
415 4.42 , $P = 0.038$, partial $\eta^2 = 0.044$, Figure 2), whereby women showed a greater
416 increase in symmetry preference when judging the sex that were in the majority
417 during the slideshow than they did when judging the sex that were in the minority
418 during the slideshow. Symmetry preferences tended to be increased from baseline
419 when judging the sex that was in the majority during the slideshow ($t_{99} = 1.53$, $P =$
420 1.13) and tended to be decreased from baseline when judging the sex that was in
421 the minority during the slideshow ($t_{99} = -1.83$, $P = 0.07$), although these differences
422 from baseline were not significant. As in Experiment 1, note that the significant main
423 effect of *face type* indicated that these difference scores were significantly different

424 from one another. The main effect of *face type* was not qualified by any higher order
425 interactions (all $F < 1.31$, all $P > 0.25$, all partial $\eta^2 < 0.014$). These null findings
426 suggest that (1) the magnitude of the effect of viewing a male-biased slideshow on
427 judgements of men's faces was not significantly different from the magnitude of the
428 effect of viewing a female-biased slideshow on judgements of women's faces and (2)
429 the effect of sex ratio cues on symmetry preference is independent of the effects of
430 *attractiveness variation*. The interaction between *slideshow type* and *attractiveness*
431 *variation* was significant ($F_{1,96} = 4.82$, $P = 0.031$, partial $\eta^2 = 0.048$, Figure 3).
432 There were no other significant effects (all $F < 1.49$, all $P > 0.22$, all partial $\eta^2 <$
433 0.016).

434

435 INSERT FIGURE 3 AROUND HERE

436

437 We used one-way ANOVAs to interpret the significant interaction between
438 *slideshow type* and *attractiveness variation*. Symmetry preferences were increased
439 to a greater extent after viewing a slideshow of male images with high variation in
440 attractiveness than after viewing a slideshow of male images with low variation in
441 attractiveness ($F_{1,48} = 6.50$, $P = 0.014$, partial $\eta^2 = 0.12$). By contrast, symmetry
442 preferences were not increased to a greater extent after viewing a slideshow of
443 female images with high variation in attractiveness than after viewing a slideshow of
444 female images with low variation in attractiveness ($F_{1,48} = 0.43$, $P = 0.51$, partial η^2
445 = 0.009).

446

447 ***Additional analyses***

448 We conducted an additional test for an effect of the sex ratio depicted in the
449 slideshow in which we combined the datasets from both experiments and analysed
450 the resultant dataset using a mixed design ANOVA with the within-subjects factor
451 *face type* (sex in majority during slideshow versus sex in minority during slideshow)
452 and the between-subjects factors *slideshow type* (male-biased versus female-
453 biased) and *experiment* (Experiment 1 versus Experiment 2). *Attractiveness variation*
454 was not included in this analysis as it was not a factor in Experiment 1 and did not
455 interact with the effect of *face type* in Experiment 2. The main effect of *face type* was
456 significant ($F_{1,196} = 10.82$, $P < 0.001$, partial $\eta^2 = 0.052$) and did not interact with
457 any other variables (all $F < 1.58$, all $P > 0.11$, all partial $\eta^2 = 0.008$). One sample t-
458 tests showed that symmetry preferences were significantly increased from baseline
459 when judging the sex that was in the majority during the slideshow ($t_{199} = 2.18$, $P =$
460 0.03) and were significantly decreased from baseline when judging the sex that was
461 in the minority during the slideshow ($t_{199} = -2.59$, $P = 0.01$).

462

463 **Discussion**

464 In both experiments, we assessed women's preferences for symmetry in
465 men's and women's faces before and after viewing slideshows in which cues to the
466 sex ratio of the local population had been manipulated. Consistent with previous
467 studies of women's preferences for symmetric faces (e.g. Perrett et al. 1999; Little et
468 al. 2003, 2006), preferences for symmetry in both men's and women's faces were
469 significantly greater than chance in the pre-slideshow (i.e. baseline) tests. As we had
470 predicted, however, symmetry preferences were sensitive to experimentally
471 manipulated cues to the sex ratio of the local population; women's preferences for
472 symmetrised faces tended to be increased when judging the attractiveness of the

473 sex that was depicted as being in the majority during the slideshow phase of the
474 experiments and tended to be decreased when judging the attractiveness of the sex
475 that was depicted as being in the minority during the slideshow phase (see Figure 2).
476 More importantly, these changes in symmetry preferences differed significantly from
477 one another in both experiments. Additionally, the effects of viewing female-biased
478 slideshows on evaluations of women's attractiveness and of viewing male-biased
479 slideshows on evaluations of men's attractiveness were equivalent (i.e. did not differ
480 significantly from one another). Thus, our findings suggest that increasing the
481 apparent proportion of a given sex in the local population increases the salience of
482 facial cues of mate quality in that sex.

483

484 Preferences for symmetric mates are thought to reflect, at least in part,
485 adaptive responses that evolved to encourage mating with high quality individuals
486 (e.g. Grammer & Thornhill 1994; Penton-Voak et al. 2001; Gangestad & Thornhill
487 2003). Additionally, and because particularly attractive women are likely to present
488 particularly significant competition for mates, cues to women's mate quality, such as
489 symmetry, may be highly salient to other women and perceptions of these cues may
490 be important for effective within-sex competition for mates (e.g. Buss & Dedden
491 1990; Maner et al. 2003; Fisher & Cox 2009; Puts et al. 2011). Increasing the
492 apparent proportion of a given sex in the local population appears to increase the
493 salience of facial cues of mate quality in that sex. This finding is consistent with both
494 increased preferences for high-quality mates when mates are relatively abundant
495 and increased salience of cues of competitors' mate quality when competition for
496 mates is likely to be relatively intense. Consequently, our findings complement those
497 from experiments with non-human species that have demonstrated effects of sex

498 ratio on both mate preferences and behaviours that might function to promote
499 effective within-sex competition (e.g. Jirotkul 1999; Clark & Grant 2010) and studies
500 that have reported correlations between natural variation in sex ratios and human
501 behaviours that are related to mating and within-sex competition (Schuster 1983,
502 1985; Campbell 1995; Pollet & Nettle 2008). Additionally, our findings present new
503 evidence that priming participants with pictorial cues to different ecological conditions
504 can influence behaviour in ways that are consistent with previous findings from
505 correlational studies of natural variation in ecological conditions, experimental
506 studies exposing human participants to relevant pictorial cues, and theories of sexual
507 selection (e.g. DeBruine et al. 2010, 2011; Little et al. 2011).

508

509 In Experiment 2, we also tested for effects of viewing cues to the degree of
510 variation in attractiveness in the local population on women's symmetry preferences.
511 We found that viewing a slideshow of male images that varied greatly in
512 attractiveness increased women's symmetry preferences more than did viewing a
513 slideshow of male images that varied little in attractiveness. Importantly, both
514 slideshows possessed the same mean attractiveness and this effect of
515 attractiveness variability was independent of the effects of viewing cues to the sex
516 ratio of the local population that were observed in both experiments and that we
517 have discussed in the previous paragraphs. By contrast with the observed effect of
518 cues to the degree of variation in male attractiveness in the local population, there
519 was no similar effect of cues to the degree of variation in female attractiveness.
520 Collectively, these findings suggest that the effects of the sex ratio of the local
521 population and of cues to the variability in the attractiveness of available potential
522 mates can indeed be dissociated. Nonetheless, these results do support the

523 proposal that cues to both the sex ratio and degree of variation in the attractiveness
524 of potential mates in the local population can be important for mating-related
525 behaviours (e.g. Owens & Thompson 1994; Kvarnemo & Simmons 1999).

526

527 The sex ratio depicted in our slideshows (5:1) may appear somewhat
528 extreme, at least when compared with the sex ratios that are typically observed for
529 relatively large geographic regions (e.g. countries), which are generally around 1:1
530 (Barber 2011). However, sex ratios that are similar to those in our experiment are
531 often encountered in real life (e.g. at a social gathering at which there is a far greater
532 proportion of women than men, or vice versa). Thus, it is important to note that sex
533 ratios of the type presented in our slideshow are not necessarily unrealistic and,
534 consequently, do not necessarily lack ecological validity. Indeed, in small-scale
535 societies, violent between-group conflict may well have led to extremely biased sex
536 ratios and, even in more recent times, the influence of war on sex ratios appears to
537 have resulted in detectable changes in women's mate choices (Pawlowski et al.
538 2000). Moreover, our findings suggest that one minute's indirect visual experience
539 can be sufficient to subtly recalibrate women's evaluations of facial attractiveness,
540 suggesting that the sex ratios encountered in relatively brief social interactions may
541 well have effects on appraisals of others' attractiveness. This rapid recalibration of
542 women's evaluations of others' attractiveness may have played an important role in
543 the evolution of mate preferences; if female dispersal was common in ancestral
544 societies, as has been suggested (e.g. Seielstad et al. 1998), facultative responses
545 that rapidly recalibrated assessments of potential mates and potential competitors for
546 mates may have helped women compete for mates more successfully in unfamiliar
547 groups. These facultative responses would also be potentially important in societies

548 where primary partner choices were commonly determined by parental, rather than
549 female, choice (e.g. arranged marriages), as they may have allowed women to better
550 evaluate the quality of potential extra-pair mates and potential competitors for extra-
551 pair mates.

552

553 Although analyses of responses in the pre-slideshow symmetry preference
554 tests suggested that the women in our experiments tended to generally prefer
555 symmetrised versions of faces over relatively asymmetric versions (i.e. the original
556 versions), the strength of these preferences was somewhat variable (see Table 1
557 and Table 2). This variability is consistent with previous work demonstrating
558 individual differences in women's preferences for facial symmetry (see Quist et al.
559 2011 for a recent review) and underlines the potential importance of controlling for
560 these individual differences as we have done in our experiments. Additionally, the
561 current work identifies recent experience with potential cues to the sex ratio of the
562 local population as an additional source of variation in symmetry preferences.

563

564 In both experiments, women's preferences for symmetrised faces tended to
565 be increased when judging the attractiveness of the sex that was depicted as being
566 in the majority during the slideshow phase and tended to be decreased when judging
567 the attractiveness of the sex that was depicted as being in the minority during the
568 slideshow phase. One interpretation of this pattern of results is that it implies that
569 participants' default expectation about the sex ratio of the local population may well
570 be approximately 1:1. However, it is also important to note that this pattern of results
571 could equally be a consequence of averaging responses across individual
572 participants, each of whom may have experienced different proportions of men and

573 women immediately prior to participating in our experiment. Additionally, our findings
574 could be due to varying cues to the absolute number of men or women in the local
575 population, rather than the sex ratio, per se. While we acknowledge this distinction
576 could well prove to be an important one, we also note that cues to the absolute
577 number of individuals of one sex in a local population and more direct cues to the
578 sex ratio itself are both likely to have similar effects on women's perceptions of the
579 size of the pool of potential mates and the size of the pool of potential competitors for
580 mates in the local population. We suggest that these issues, together with
581 experiments testing whether similar effects occur for perceptions of other putative
582 cues of potential mates' and competitors' physical condition (e.g. sexually dimorphic
583 shape characteristics and perceived health in faces) are potentially important topics
584 for future research.

585

586 The cognitive and/or perceptual mechanisms that might underpin the sex ratio
587 effects observed in our experiments remain unclear. It is unlikely that our findings
588 reflect a simple priming effect whereby viewing potential mates' faces triggers a
589 mating psychology under which cues of potential mates' qualities might become
590 more relevant. For example, the women in our experiments showed corresponding
591 effects for judgements of women's faces following female-biased slideshows,
592 demonstrating that the observed effects are not specific to experience with opposite-
593 sex faces. Similarly, it is unlikely that our findings simply reflect improved symmetry
594 detection for a given sex of face after recent experience with a larger number of
595 exemplars of that sex, which may have represented a wider range of facial
596 asymmetries than a smaller number of faces would have. Given the correlation
597 between symmetry and facial attractiveness, the variation in symmetry should have

598 been greater in the high variance than low variance conditions in Experiment 2. This
599 issue is noteworthy since the effect of viewing cues to the sex ratio of the local
600 population in Experiment 2 was found to be wholly independent of the effect of
601 *attractiveness variation*. Indeed, facial symmetry detection and facial symmetry
602 preferences do not appear to be correlated in the way that a simple symmetry
603 detection explanation for our findings would require (e.g. Little & Jones 2006;
604 Oinonen & Mazmanian 2007). Finally, although our findings suggest that very little
605 experience with cues to the sex ratio of the local population is needed to alter
606 symmetry preferences, suggesting that our findings reveal a short-term, facultative
607 response to the current environment, it is unclear whether similar cues can also have
608 long-term effects on behaviour. For example, more prolonged exposure to cues to
609 relatively stable (i.e. unchanging) sex ratios may have longer-term effects, as may
610 experience with cues to the sex ratio of the local population during critical periods of
611 development, such as childhood or adolescence. Our data do not speak directly to
612 these issues and they require further investigation.

613

614 In summary, in two experiments, we found that manipulating pictorial cues to
615 the sex ratio of the local population altered women's assessments of others' facial
616 attractiveness. Increasing the apparent proportion of a given sex in the local
617 population increased the salience of facial cues of quality in that sex. This effect of
618 manipulating cues to the sex ratio of the local population on perceptions of
619 symmetric versus relatively asymmetric faces is consistent with those observed in
620 both previous experiments with non-human species (e.g. Jirotkul 1999; Clark &
621 Grant 2010) and prior correlational studies of naturally occurring variation in human
622 behaviour (e.g. Schuster 1983, 1985; Campbell 1995; Pollett & Nettle 2008).

623 Moreover, the facultative responses observed in the current research may be
624 adaptive; they suggest that women's evaluations of the attractiveness of both
625 potential mates and potential competitors for mates are sensitive to cues of the sex
626 ratio of the local population in ways that might function to promote efficient allocation
627 of mating effort (i.e. increase preferences for high-quality mates when competition for
628 mates is likely to be less intense) and successful within-sex competition (i.e.
629 increased salience of cues to the quality of potential competitors for mates when
630 competition for mates is likely to be relatively intense), respectively. More
631 fundamentally, our findings suggest that both inter-sexual and intra-sexual selection
632 have been important for shaping the effects of sex ratio on person perception.

633

634

635 **References**

- 636 **Andersson, M.** 1994. *Sexual Selection*. Princeton, NJ: Princeton University Press.
- 637 **Balshine-Earn, S.** 1996. Reproductive rates, operational sex ratios and mate choice
638 in St Peter's fish. *Behavioral Ecology and Sociobiology*, **39**, 107–116.
- 639 **Barber, N.** 2011. Marriage markets and mating aggression help explain societal
640 differences in violent crime. *Aggression and Violent Behavior*, **16**, 420-427.
- 641 **Berglund, A.** 1994. Operational sex ratio influences choosiness in a pipefish.
642 *Behavioral Ecology*, **5**, 254-258.
- 643 **Buss, D. M. & Dedden, L. A.** 1990. Derogation of competitors. *Journal of Social and*
644 *Personal Relationships*, **7**, 395-422.
- 645 **Campbell, A.** 1995. A few good men: Evolutionary psychology and female
646 adolescent aggression. *Ethology and Sociobiology*, **16**, 99-123.
- 647 **Clark, L. & Grant, J. W. A.** 2010. Intrasexual competition and courtship in female
648 and male Japanese medaka, *Oryzias latipes*: Effects of operational sex ratio
649 and density. *Animal Behaviour*, **80**, 707-712.
- 650 **Clutton-Brock, T. H., Rose, K. E. & Guinness, F. E.** 1997. Density related changes
651 in sexual selection in red deer. *Proceedings of the Royal Society of London B*,
652 **264**, 1509–1516.
- 653 **Coltman, D. W., Smith, J. A., Bancroft, D. R., Pilkington, J., MacColl, A. D. C.,**
654 **Clutton-Brock, T. H. & Pemberton, J. M.** 1999. Density-dependent variation
655 in lifetime breeding success and natural sexual selection in soay rams. *The*
656 *American Naturalist*, **154**, 730-746.
- 657 **Colwell, M. A. & Oring, L. W.** 1988. Sex ratios and intrasexual competition for
658 mates in a sex-role reversed shorebird, Wilson's phalarope (*Phalaropus*
659 *tricolor*). *Behavioral Ecology and Sociobiology*, **22**, 165–173.

- 660 **Conway, C. A., Jones, B. C., DeBruine, L. M. & Little, A. C.** 2008. Evidence for
661 adaptive design in human gaze preference. *Proceedings of the Royal Society*
662 *of London B*, **275**, 63-69.
- 663 **DeBruine, L. M., Jones, B. C., Crawford, J. R., Welling, L. L. M. & Little, A. C.**
664 2010. The health of a nation predicts their mate preferences: Cross-cultural
665 variation in women's preferences for masculinized male faces. *Proceedings of*
666 *the Royal Society of London B*, **277**, 2405-2410.
- 667 **DeBruine, L. M., Jones, B. C., Little, A. C., Crawford, J. R. & Welling, L. L. M.**
668 2011. Further evidence for regional variation in women's masculinity
669 preferences. *Proceedings of the Royal Society of London B*, **278**, 813-814.
- 670 **Dick, J. T. A. & Elwood, R. W.** 1996. Effects of natural variation in sex ratio and
671 habitat structure on mate-guarding decisions in amphipods (*Crustacea*).
672 *Behaviour*, **133**, 985–996.
- 673 **Emlen, S. T. & Oring, L. W.** 1977. Ecology, sexual selection and the evolution of
674 mating systems. *Science*, **197**, 215–23.
- 675 **Fisher, M. & Cox, A.** 2009. The influence of female attractiveness on competitor
676 derogation. *Journal of Evolutionary Psychology*, **7**, 141-155.
- 677 **Forsgren, E., Amundsen, T., Borg, A. A. & Bjelvenmark, J.** 2004. Unusually
678 dynamic sex roles in a fish. *Nature*, **429**, 551-554.
- 679 **Gangestad, S. W. & Buss, D. M.** 1993. Pathogen prevalence and human mate
680 preferences. *Ethology and Sociobiology*, **14**, 89–96.
- 681 **Gangestad, S. W. & Simpson, J. A.** 2000. The evolution of human mating: trade-
682 offs and strategic pluralism. *Behavioral Brain Sciences*, **23**, 675-687.
- 683 **Gangestad S. W. & Thornhill, R.** 2003. Facial masculinity and fluctuating
684 asymmetry. *Evolution and Human Behavior*, **24**, 231–241.

- 685 **Grammer, K., & Thornhill, R.** 1994. Human (*Homo sapiens*) facial attractiveness
686 and sexual selection: The role of symmetry and averageness. *Journal of*
687 *Comparative Psychology*, **108**, 233–242.
- 688 **Guttentag, M. & Secord, P. F.** 1983. *Too many women? The sex ratio question.*
689 Beverly Hills: Sage.
- 690 **Hohmann, G. & Fruth, B.** 2003. Intra- and inter-sexual aggression by bonobos in
691 the context of mating. *Behaviour*, **140**, 1389-1413.
- 692 **Hudson, V.M. & Den Boer, A.** 2002. A surplus of men, a deficit of peace: Security
693 and sex ratio in Asia's largest states. *International Security*, **26**, 5–38.
- 694 **Hudson, V.M. & Den Boer, A.** 2004. *Bare branches: Security implications of Asia's*
695 *surplus male population.* Cambridge, MA: MIT Press.
- 696 **Hume, D. K. & Montgomerie, R.** 2001. Facial attractiveness signals different
697 aspects of "quality" in women and men. *Evolution and Human Behavior*, **22**,
698 93-112.
- 699 **Jennions, M. D. & Petrie, M.** 1997. Variation in mate choice and mating
700 preferences: A review of causes and consequences. *Biological Reviews*, **72**,
701 283–327.
- 702 **Jirotkul, M.** 1999. Operational sex ratio influences female preference and male-male
703 competition in guppies. *Animal Behaviour*, **58**, 287-294.
- 704 **Jones, B. C., DeBruine, L. M., Little, A. C., Burriss, R. P. & Feinberg, D. R.** 2007.
705 Social transmission of face preferences among humans. *Proceedings of the*
706 *Royal Society of London B*, **274**, 899-903.
- 707 **Jones, B. C., Little, A. C., Feinberg, D. R., Penton-Voak, I. S., Tiddeman, B. P. &**
708 **Perrett, D. I.** 2004. The relationship between shape symmetry and perceived

709 skin condition in male facial attractiveness. *Evolution and Human Behavior*,
710 **25**, 24-30.

711 **Jones, B. C., Little, A. C., Penton-Voak, I. S., Tiddeman, B. P., Burt, D. M. &**
712 **Perrett, D. I.** 2001. Facial symmetry and judgments of apparent health:
713 support for a “good genes” explanation of the attractiveness–symmetry
714 relationship. *Evolution and Human Behavior*, **22**, 417–429.

715 **Jones, B. C., Perrett, D. I., Little, A. C., Boothroyd, L., Cornwell, R. E., Feinberg,**
716 **D. R., Tiddeman, B. P., Whiten, S., Pitman, R. M., Hillier, S. G., Burt, D. M.,**
717 **Stirrat, M. R., Law Smith, M. J. & Moore, F. R.** 2005. Menstrual cycle,
718 pregnancy and oral contraceptive use alter attraction to apparent health in
719 faces. *Proceedings of the Royal Society of London B*, **272**, 347-354.

720 **Kvarnemo, C. & Ahnesjo, I.** 1996. The dynamics of operational sex ratios and
721 competition for mates. *Trends in Ecology and Evolution*, **11**, 404-408.

722 **Kvarnemo, C. & Simmons, L. W.** 1999. Variance in female quality, OSR and male
723 mate choice in a bushcricket. *Behavioral Ecology & Sociobiology*, **45**, 245-
724 252.

725 **Lawrence, W. S.** 1986. Male choice and competition in *Tetraopes tetraophthalmus*:
726 Effects of local sex ratio variation. *Behavioral Ecology and Sociobiology*, **18**,
727 289–296.

728 **Little, A. C., Burt, D. M., Penton-Voak, I. S. & Perrett, D. I.** 2001. Self-perceived
729 attractiveness influences human female preferences for sexual dimorphism
730 and symmetry in male faces. *Proceedings of the Royal Society of London B*,
731 **268**, 39-44.

732 **Little, A. C., Cohen, D. L., Jones, B. C. & Belsky, J.** 2007a. Human preferences
733 for facial masculinity change with relationship type and environmental
734 harshness. *Behavioral Ecology and Sociobiology*, **61**, 967–973.

735 **Little, A. C., DeBruine, L. M. & Jones, B. C.** 2011. Exposure to visual cues of
736 pathogen contagion changes preferences for masculinity and symmetry in
737 opposite-sex faces. *Proceedings of the Royal Society of London B*, **278**, 2032-
738 2039.

739 **Little, A. C., & Jones, B. C.** 2003. Evidence against perceptual bias views for
740 symmetry preferences in human faces. *Proceedings of the Royal Society of*
741 *London B*, **270**, 1759–1763.

742 **Little, A. C. & Jones, B. C.** 2006. Attraction independent of detection suggests
743 special mechanisms for symmetry preferences in human face perception.
744 *Proceedings of the Royal Society of London B*, **273**, 3093-3099.

745 **Little, A. C., Jones, B. C., Burt, D. M. & Perrett, D. I.** 2007b. Preferences for
746 symmetry in faces change across the menstrual cycle. *Biological Psychology*,
747 **76**, 209.

748 **Little, A. C., Jones, B. C., Waite, C., Tiddeman, B. P., Feinberg, D. R., Perrett, D.**
749 **I., Apicella, C. A. & Marlowe F. W.** 2008. Symmetry is related to sexual
750 dimorphism in faces: Data across culture and species. *PLoS One*, **3**, 2106.

751 **Maner, J. K., Kenrick, D. T., Becker, D. V., Delton, A. W., Hofer, B., Wilbur, C. J.**
752 **& Neuberg, S. L.** 2003. Sexually selective cognition: Beauty captures the
753 mind of the beholder. *Journal of Personality and Social Psychology*, **85**, 1107-
754 1120.

755 **Mathews, L. M.** 2002. Tests of the mate-guarding hypothesis for social monogamy:
756 Does population density, sex ratio, or female synchrony affect behavior of

- 757 male snapping shrimp (*Alpheus angulatus*). *Behavioral Ecology and*
758 *Sociobiology*, **51**, 426-432.
- 759 **Michener, G. R. & Locklear, L.** 1990. Differential costs of reproductive effort for
760 male and female Richardson's ground squirrels. *Ecology*, **71**, 855-868.
- 761 **Noe, R. & Hammerstein, P.** 1994. Biological markets: Supply and demand
762 determine the effect of partner choice in cooperation, mutualism and mating.
763 *Behavioral Ecology and Sociobiology*, **35**, 1-11.
- 764 **Oinonen, K. A. & Mazmanian, D.** 2007. Facial symmetry detection ability changes
765 across the menstrual cycle. *Biological Psychology*, **75**, 136–145.
- 766 **Owens, I. P. F. & Thompson, D. B. A.** 1994. Sex differences, sex ratios and sex
767 roles. *Proceedings of the Royal Society of London B*, **258**, 93-99.
- 768 **Özener, B.** 2010. Fluctuating and directional asymmetry in young human males:
769 Effect of heavy working condition and socioeconomic status. *American*
770 *Journal of Physical Anthropology*, **143**, 112-120.
- 771 **Özener, B. & Fink, B.** 2010. Facial symmetry in young girls and boys from a slum
772 and a control area of Ankara, Turkey. *Evolution and Human Behavior*, **31**,
773 436-441.
- 774 **Pawłowski, B., Dunbar, R. I. M. & Lipowicz, A.** 2000. Tall men have more
775 reproductive success. *Nature*, **403**, 156.
- 776 **Pedersen, F. A.** 1991. Secular trends in human sex ratios: Their influence on
777 individual and family behavior. *Human Nature*, **2**, 271–91.
- 778 **Penton-Voak I. S., Jacobson, A. & Trivers, R.** 2004. Populational differences in
779 attractiveness judgments of male and female faces: comparing British and
780 Jamaican samples. *Evolution and Human Behaviour*, **25**, 355-370.

781 **Penton-Voak, I. S., Jones, B. C., Little, A. C., Baker, S., Tiddeman, B, Burt, D. M.**
782 **& Perrett, D. I.** 2001. Symmetry, sexual dimorphism in facial proportions, and
783 male facial attractiveness. *Proceedings of the Royal Society of London B*,
784 **268**, 1617–1623.

785 **Perrett D. I., Burt, D. M., Penton-Voak, I. S., Lee, K. J., Rowland, D. A. &**
786 **Edwards, R.** 1999. Symmetry and human facial attractiveness. *Evolution and*
787 *Human Behavior*, **20**, 295–307.

788 **Pollet, T. V. & Nettle, D.** 2008. Driving a hard bargain: Sex ratio and male marriage
789 success in a historical US population. *Biology Letters*, **4**, 31-33.

790 **Puts, D. A., Barndt, J. L., Welling, L. L. M., Dawood, K. & Burriss, R. P.** 2011.
791 Intrasexual competition among women: Vocal femininity affects perceptions of
792 attractiveness and flirtatiousness. *Personality and Individual Differences*, **50**,
793 111-115.

794 **Quist, M., Watkins, C. D., Smith, F. G., Little, A. C., DeBruine, L. M. & Jones, B.**
795 **C.** 2011. Sociosexuality predicts women's preferences for symmetry in men's
796 faces. *Archives of Sexual Behavior*, in press.

797 **Sacco, D. F., Hugenberg, K. & Sefcek, J. A.** 2009. Sociosexuality and face
798 perception: Unrestricted sexual orientation facilitates sensitivity to female
799 facial cues. *Personality and Individual Differences*, **47**, 777–782.

800 **Schuster, I.** 1983. Women's aggression: An African case study. *Aggressive*
801 *Behavior*, **9**, 319-331.

802 **Schuster, I.** 1985. Female aggression and resource scarcity: A cross-cultural
803 perspective. In: *The Aggressive Female* (Ed by M. Haug, D. Benton, P. Brain,
804 B. Oliver & J. Mos.), pp. 185-208. The Hague: CIP-Gegevens Koninklijke
805 Bibliotheek.

- 806 **Seielstad, M. T., Minch, E. & Cavalli-Sforza, L. L.** 1998. Genetic evidence for a
807 higher female migration rate in humans. *Nature Genetics*, **20**, 278–280.
- 808 **Souroukis, K., & Murray, A. M.** 1994. Female mating behavior in the field cricket,
809 *Gryllus pennsylvanicus* (Orthoptera: Gryllidae) at different operational sex
810 ratios. *Journal of Insect Behaviour*, **8**, 269-279.
- 811 **Thornhill, R., & Gangestad, S. W.** 2006. Facial sexual dimorphism, developmental
812 stability, and susceptibility to disease in men and women. *Evolution and*
813 *Human Behavior*, **27**, 131–144.
- 814 **Verrell, P. A.** 1983. The influence of ambient sex ratio and intramale competition on
815 the sexual behaviour of the redspotted newt, *Notophthalmus viridescens*
816 (Amphibia: Urodela: Salamandridae). *Behavioral Ecology and Sociobiology*,
817 **13**, 307–313.
- 818 **Wang, X, Li, J., Xia, D., Chen, R., Zhu, Y., Zhang, M. & Wang, S.** 2009.
819 Comparisons on reproductive and aggressive behaviors in male Tibetan
820 Macaques (*Macaca thibetana*) of two different operational sex ratio troops at
821 Huangshan, China. *Zoological Research*, **30**, 83-89.
- 822 **Young, S. G., Sacco, D. F. & Hugenberg, K.** 2011. Vulnerability to disease is
823 associated with a domain-specific preference for symmetrical faces relative to
824 symmetrical non-face stimuli. *European Journal of Social Psychology*, **41**,
825 558–563.
- 826

827 **Table 1.** Summary statistics for each individual condition in Experiment 1. *t*, *P*, and *d*
 828 statistics are for one-sample t-tests comparing scores with the chance value of 0.5.

829

slideshow type	sex of face judged	test phase	mean proportion of symmetrised faces chosen (SEM)	<i>t</i>₄₉	<i>d</i>
female-biased	male	pre-slideshow	0.56 (0.02)	2.46*	0.35
female-biased	male	post-slideshow	0.54 (0.02)	1.69 ⁺	0.24
female-biased	female	pre-slideshow	0.56 (0.03)	2.40*	0.33
female-biased	female	post-slideshow	0.62 (0.03)	4.34***	0.63
male-biased	male	pre-slideshow	0.54 (0.02)	1.73 ⁺	0.25
male-biased	male	post-slideshow	0.56 (0.02)	2.69**	0.39
male-biased	female	pre-slideshow	0.66 (0.02)	6.84***	0.94
male-biased	female	post-slideshow	0.59 (0.03)	3.72***	0.51

830 ⁺ < 0.10, * < 0.05, ** < 0.01, *** < 0.001

831

832
833
834

Table 2. Summary statistics for each individual condition in Experiment 2. *t*, *P*, and *d* statistics are for one-sample t-tests comparing scores with the chance value of 0.5.

slideshow type	sex of face judged	attractiveness variation	test phase	mean proportion of symmetrised faces chosen (SEM)	<i>t</i>₂₄	<i>d</i>
female-biased	male	high	pre-slideshow	0.58 (0.03)	2.22*	0.45
female-biased	male	high	post-slideshow	0.52 (0.03)	0.74	0.15
female-biased	female	high	pre-slideshow	0.60 (0.04)	2.49*	0.50
female-biased	female	high	post-slideshow	0.60 (0.04)	2.71*	0.54
male-biased	male	high	pre-slideshow	0.48 (0.03)	-0.90	0.15
male-biased	male	high	post-slideshow	0.56 (0.03)	1.73 ⁺	0.35
male-biased	female	high	pre-slideshow	0.60 (0.04)	2.45*	0.50
male-biased	female	high	post-slideshow	0.60 (0.04)	2.90**	0.58
female-biased	male	low	pre-slideshow	0.56 (0.03)	1.90 ⁺	0.38
female-biased	male	low	post-slideshow	0.50 (0.04)	-0.11	0.02
female-biased	female	low	pre-slideshow	0.58 (0.03)	3.00**	0.60
female-biased	female	low	post-slideshow	0.65 (0.03)	5.06***	1.01
male-biased	male	low	pre-slideshow	0.52 (0.03)	0.54	0.11
male-biased	male	low	post-slideshow	0.48 (0.03)	-0.64	0.13
male-biased	female	low	pre-slideshow	0.58 (0.03)	3.06**	0.61
male-biased	female	low	post-slideshow	0.54 (0.03)	1.24	0.25

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⁺ < 0.10, * < 0.05, ** < 0.01, *** < 0.001

837 **Figure Captions**

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839 **Figure 1.** Symmetrised (left) and original versions (right) of men's and women's
840 faces used in our experiment.

841

842 **Figure 2.** Change in symmetry preferences for the *majority sex* and *minority sex*
843 conditions in Experiments 1 and 2. Bars show means and SEMs.

844

845 **Figure 3.** Change in symmetry preferences for the high and low attractiveness
846 variance conditions after exposure to male or female slideshows in Experiment 2.
847 Bars show means and SEMs.

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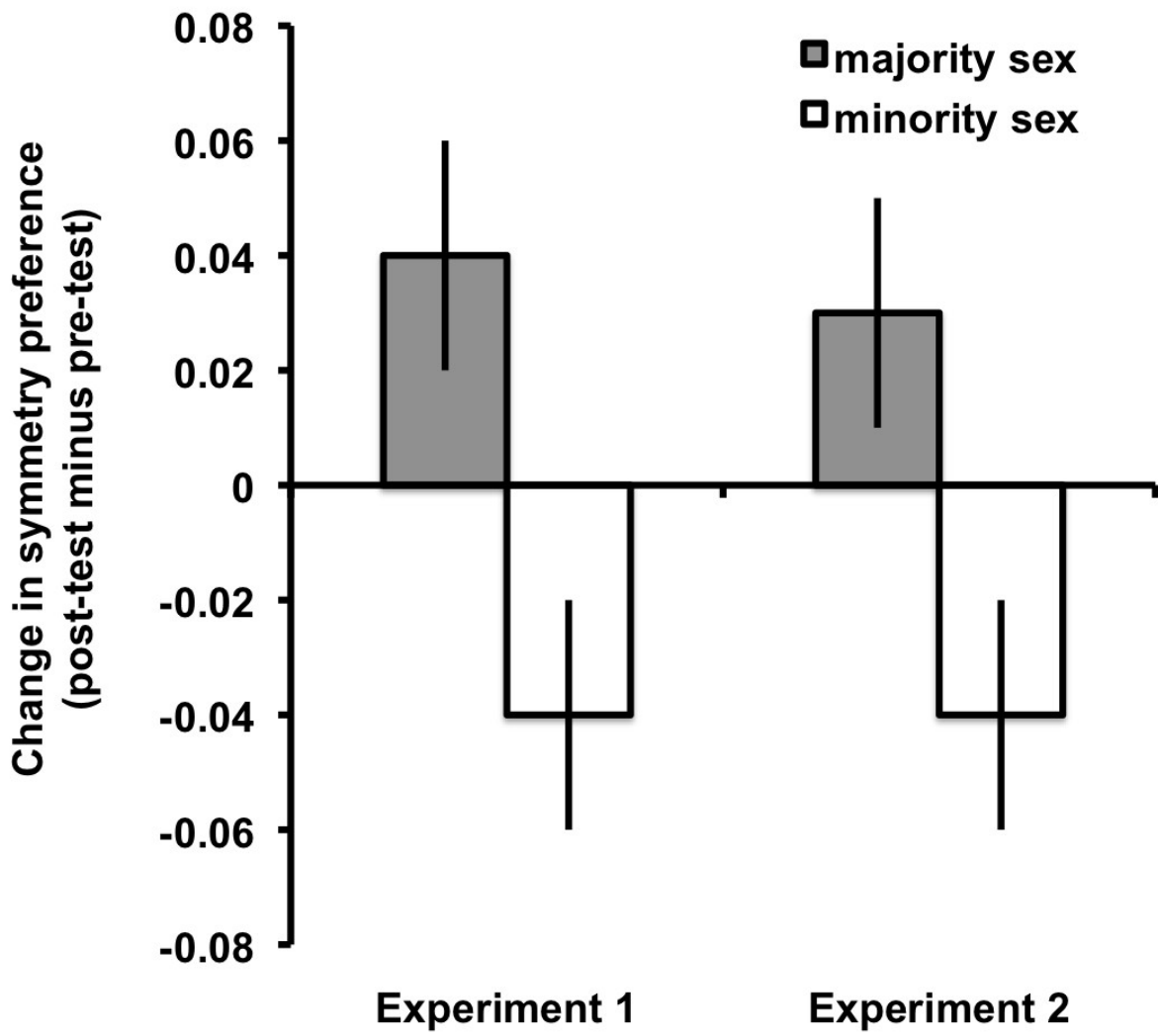
850 **Figure 1.**



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852 **Figure 2.**

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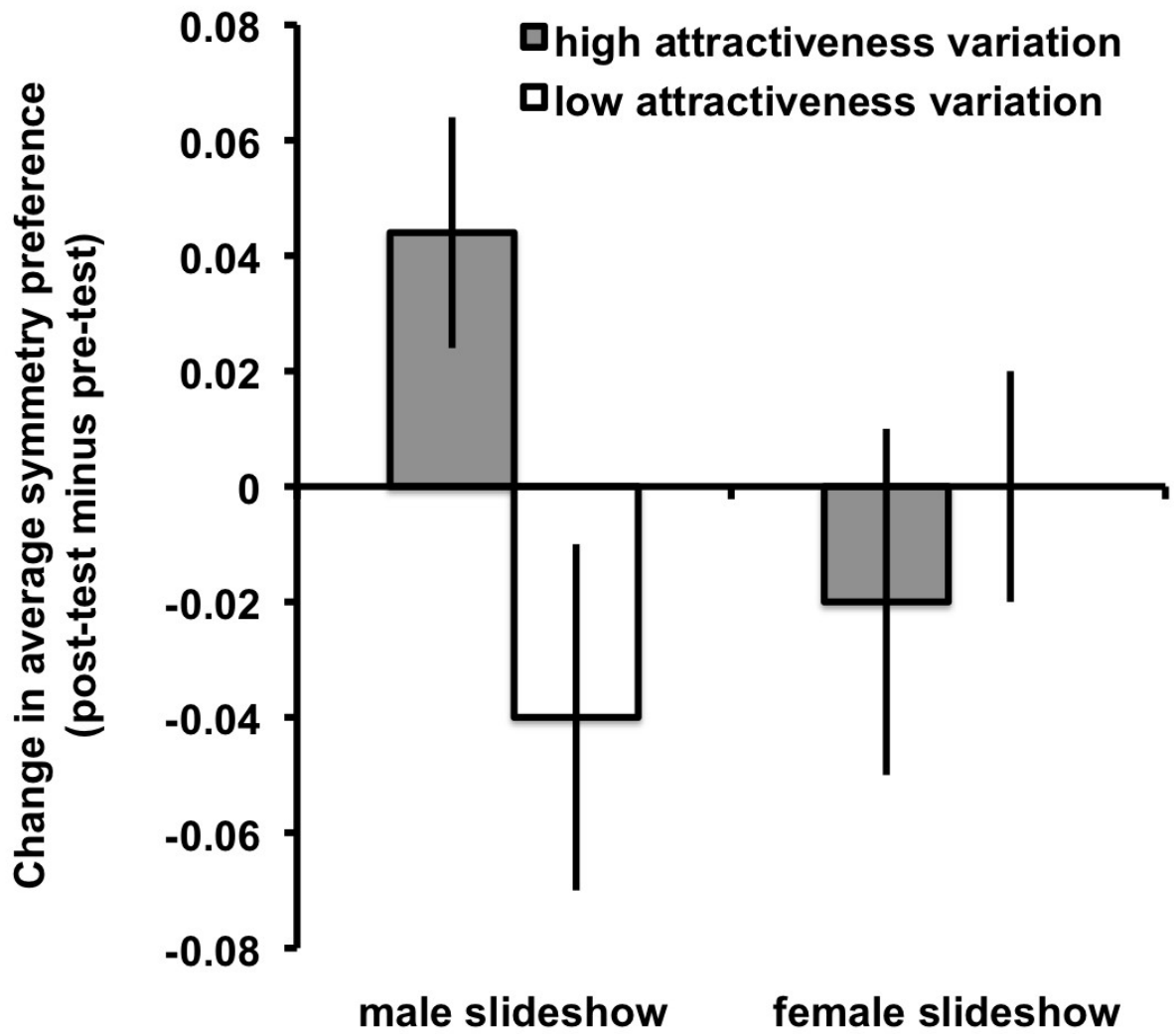


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857 Figure 3.



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