# 1 Phytoplankton community structure and dynamics in the North

# 2 Atlantic subtropical gyre

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#### Abstract

Phytoplankton fuel epipelagic ecosystems and affect global biogeochemical cycles. Nevertheless, there is still a lack of quantitative information about the factors that determine both phytoplankton community structure and dynamics, particularly in subtropical gyres. Here, we estimated size fractionated phytoplankton growth  $(\mu)$  and microzooplankton grazing rates (m) along a transect in the subtropical North Atlantic, from the island of Hispaniola to the Iberian Peninsula, by conducting dilution experiments and fitting mixed models. We also examined the relationship between nutrient availability and the differences in both phytoplankton community structure and size fractionated phytoplankton growth rates at two spatial scales (i.e. subtropical gyre and within-province spatial scale). Our results revealed high values for both phytoplankton growth and microzooplankton grazing rates. Phytoplankton growth (0.00 - 1.19 d<sup>-1</sup>) displayed higher variability among stations, biogeochemical provinces and size fractions than the microzooplankton grazing rate  $(0.32 - 0.74 \text{ d}^{-1})$ . Differences in phytoplankton community structure were associated with dissolved inorganic nitrogen  $(0.72-5.85 \mu M; R^2=0.19)$  and squared Brunt-Väisälä frequency ( $R^2=0.21$ ) at the whole gyre scale. Conversely, the differences in phytoplankton growth rate showed a weak relationship with those properties ( $R^2 \le 0.05$ ) at that scale, but a stronger relationship at the within province scale ( $R^2 \ge 0.07$ ). These results support the idea that phytoplankton grow at high rates in oligotrophic subtropical gyres, this is likely due to the selection of phytoplankton groups with functional traits suited to exploit low nutrient availability. Thus, shedding new, multi-scale knowledge on the commonly misunderstood "ocean deserts".

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## 1. Introduction

Phytoplankton influence most components of epipelagic ecosystems (Reynolds 2001) and affect global biogeochemical cycles (Falkowski et al. 1998). Phytoplankton community structure and dynamics are mainly the result of the balance between growth and mortality. Phytoplankton growth at a community level is determined by resource availability. Nevertheless, phytoplankton growth rate at the community level may also be impacted by the functional traits, related to resource acquisition and growth, of the populations that compose said community, i.e. by the phytoplankton community composition. Despite being influenced by several factors, phytoplankton mortality is driven by microzooplankton grazing mainly (Calbet and Landry 2004). Microzooplankton grazing may also influence phytoplankton growth through nutrient regeneration, particularly in oligotrophic waters (Goldman 1984). To understand and predict phytoplankton community structure and dynamics and ecosystem functioning, the variability in phytoplankton growth and microzooplankton grazing must be disentangled. However, few studies discussed this question (e.g. Landry et al., 2009). In fact, to our knowledge, only the review of Calbet and Landry (2004) did it at a global scale. According to their results, differences among habitats were more pronounced in phytoplankton growth than in microzooplankton grazing rates. The North Atlantic subtropical gyre mainly encompasses two biogeochemical provinces as defined by Longhurst (2007); the North Atlantic Tropical Gyral Province (NATR) and the North Atlantic Subtropical Gyral Province (NAST), which is divided in two sub-provinces (NAST-W and NAST-E). In those provinces, it is often believed that phytoplankton communities are characterized by low biomass, primary production and growth rates; and dominated by picoplankton. This is commonly attributed to the low

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nutrient concentrations in the area (Marañón et al. 2000; Marañón 2005; Teira et al. 2005). However, the influence of nutrient availability on phytoplankton community structure and growth rate at different spatial scales (i.e., at a subtropical gyre or at a within-province spatial scale) has rarely been compared, despite the known importance of scale in ecological processes (see Levin 1992). Also, the influence of phytoplankton community composition, suited to exploit the low nutrient conditions, on the growth of the phytoplankton community might be misunderstood.

Here we used a novel approach to investigate the variability of phytoplankton growth rate  $(\mu)$  and microzooplankton grazing rate (m) along with the relationship between nutrient availability and both the phytoplankton growth and community structure across the subtropical North Atlantic Ocean. First, we grouped the sampling stations into provinces and subprovinces defined by Longhurst (2007). Second, through dilution experiments (Landry and Hassett 1982) and mixed models we estimated phytoplankton growth and microzooplankton grazing rates for each province, size fraction and sampling station. To our knowledge, this is the first study where mixed models were employed to analyze data from dilution experiments. Third, we examined the relationship between phytoplankton community structure and phytoplankton growth and the effect of nutrient availability on both these variables. These analyses were carried out at the subtropical gyre spatial scale, which encompassed all sampled area, and at the within-province spatial scale. Our results showed that the variability of phytoplankton growth rate was higher than the variability of microzooplankton grazing rate. In addition, we found that nutrient availability only had a weak influence on the sizefractionated phytoplankton growth rates at the subtropical gyre spatial scale.

## 2. Methods

We sampled 16 stations along a SW-NE transect in the North Atlantic Ocean, between the SE of Hispaniola island of Hispaniola (S1, 67.48°W 19.26°N, March 24th) and the NW of the Iberian Peninsula (S16, 14.73°W 41.57°N, April 8<sup>th</sup>) as part of the *Buque Escuela Oceanográfica 2011* initiative (Fig. 1), within the framework of *Malaspina 2010 Expedition*. We performed 12 dilution experiments to estimate phytoplankton growth and microzooplankton grazing rates (Fig. 1) throughout the crossed biogeochemical provinces (NATR and NAST). The dilution experiments analyses were complemented with data on the physical, chemical and biological properties of the water column and satellite-derived altimetry and geostrophic velocities.

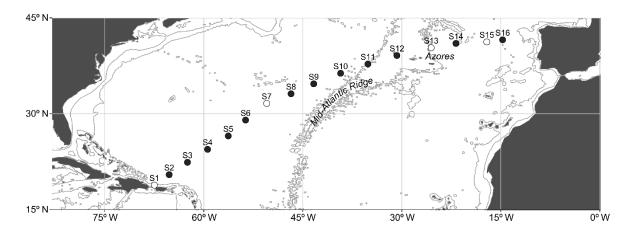


Fig. 1 Map showing the location of the 16 sampling stations (S1-S16) between Hispaniola and the Iberian Peninsula. Black dots indicate stations where dilution experiments were performed. White dots represent stations where experiments were not conducted.

#### 2.1. Water column properties

Vertical distributions of temperature, salinity and fluorescence were obtained using a SBE-19 CTD equipped with a SeaPoint fluorometer mounted in a rosette equipped with 24, 12 L Niskin bottles. We estimated seawater potential density anomaly  $(\sigma_{\theta})$  from temperature, salinity and pressure. Subsequently, squared Brunt-Väisälä frequency  $(N^2)$ 

was calculated using the *oce* R package (Kelley 2014). Nutrient concentrations (NO<sub>3</sub><sup>-</sup>, NO<sub>2</sub><sup>-</sup>, NH<sub>4</sub><sup>+</sup>, PO<sub>4</sub><sup>-</sup> and silicates) were measured for water samples at several depths (5, 25, 50, 75, 100, 125, 150, 175 and 200 m depth) using Niskin bottles. Two aliquots from each depth were collected in polystyrene tubes and preserved at -80°C until their analysis with a Skalar autoanalyzer using the methods described in Tréguer and Le Corre (1975).

#### 2.2. Remote sensing data

Remotely sensed altimeter products and absolute geostrophic satellite data were obtained for the sampling period from Ssalto/Duacs and distributed by Aviso, with support from Cnes (<a href="http://www.aviso.oceanobs.com/duacs/">http://www.aviso.oceanobs.com/duacs/</a>). Gridded geostrophic velocity and sea level anomaly data were estimated by merging data from several altimeters using the methods developed by Le Traon et al. (1998). Using this information, we identified several processes that can alter the sea water properties and directly affect local phytoplankton communities.

# 2.3. Classification of the stations

We sampled across a large area with heterogeneous biogeochemical properties, which encompassed two biogeochemical provinces defined in Longhurst (2007); NATR and NAST (subdivided into NAST-W and NAST-E). Provinces are constrained to a range of latitudes and longitudes, but they do not have a clearly defined extension. We combined the geographic and biogeochemical criteria proposed by Longhurst (2007) with visual inspection of vertical profiles of sea water properties, satellite images of geostrophic velocities and multivariate analysis techniques to classify the stations in the above mentioned provinces.

We obtained a symmetric dissimilarity matrix for the stations using Manhattan distance with the following standardized sea water properties: fluorescence, salinity and potential temperature at 10 m depth, depth of the chlorophyll maximum, sum of the squared Brunt-Väisälä frequency in the upper 200 m and the depth of the maximum squared Brunt-Väisälä frequency. Subsequently, we performed a non-metric multidimensional scaling (NMDS) based on stress minimization by means of majorization (SMACOF) using the *Smacof* R package (de Leeuw 2009) in R computing software (R Core Team 2014). We fitted each covariate to the two dimensions of the ordination space using the *vegan* package (Oksanen et al. 2013). This showed which variables were associated with the differences between stations.

## 2.4. Sampling and Experimental set-up

Water samples were collected from the maximum potential phytoplankton growth rate depth between 3 and 11 h (local time) using 12 L Niskin bottles. The maximum potential phytoplankton growth rate depth in the subtropical North Atlantic has been found slightly above the DCM (Cáceres et al. 2013). When the DCM was not observed (stations from NAST), we sampled at a depth with a similar percentage of surface irradiance to minimize any bias that might occur due to differences in light. These depths were selected by the fluorescence profiles and were further corroborated through chlorophyll profiles, constructed using fluorescence profiles, following the methodology employed in Graziano et al. (1996) based in Morel (1987) (Table 1).

Table 1. Sampling time, depth, approximate percentage of surface irradiance at the sampling and nutrients at the different stations.

Station	Sampling time	Depth (m)	Surface irradiance (%)	DIN (μM)	Silicates (µM)
S2	6:50	80	15	0.84	2.29
S3	7:00	80	14	0.84	0.98
S4	6:40	80	13	0.72	0.95
S5	6:50	80	15	0.81	0.89
S6	6:20	70	13	1.13	0.82
S8	10:40	50	5	0.89	0.83
S9	8:20	40	8	1.98	1.06
S10	8:40	40	8	5.85	2.06
S11	8:30	40	7	2.04	1.02
S12	8:00	25	16	2.95	1.39
S14	11:10	30	8	4.22	1.23
S16	8:10	20	9	3.06	0.46

Water was transferred to 25 L polyethylene carboys, wrapped in black plastic to avoid light exposure, using silicone tubing fitted with 200 µm mesh to eliminate mesozooplankton. Water from one of the carboys was filtered through a 0.2 µm AcroPak 1000 capsule filter with a Supor membrane to obtain fully diluted water. The first liters filtered were discarded in every experiment and filter capsules were changed every six experiments. Next, polycarbonate containers of 2.3 L were gently filled with different proportions of filtered and unfiltered seawater. In this study, we used four dilution treatments with dilution factor (*f*) of 1 (undiluted water), 0.75, 0.5 and 0.25 with two replicates for each treatment. Additionally, we incubated two undiluted containers with added nutrients to check the potential effects of nutrients. Nutrient mixture added to nutrient enriched treatments resulted in a final concentration of 1 mM ammonium (NH<sub>4</sub>Cl), 0.5 mM phosphate (H<sub>3</sub>PO<sub>4</sub>), 5 nM iron (FeSO<sub>4</sub>) and 0.1 nM manganese (MnSO<sub>4</sub>). We did not add nutrients to all the treatments due to potential negative effects on the plankton community (Landry and Hassett 1982; Lessard and Murrell 1998).

Therefore, there is a risk of underestimating the experimental phytoplankton growth rates as a consequence of poor nutrient regeneration in the most diluted treatments.

We used on-deck incubators with calibrated blue light filters to simulate in situ light conditions. They were covered with black plastic at night to protect the experiments from the ship's lights. Incubators were kept at a homogenous temperature that closely resembled the in situ seawater temperature ( $\pm$  0.1°C). Capsule filters, tubes and containers were soaked and rinsed in 10 % HCL-Milli Q water and rinsed with Milli-Q after every experiment. Just before each experiment, they were rinsed with the 0.2  $\mu$ m filtered seawater. Carboys were rinsed with Milli Q water after every use and rinsed with seawater from the sampling depth before every experiment.

#### 2.5. Chlorophyll a, flow cytometry and phytoplankton

Two 1000 mL samples of undiluted seawater were taken from the 25 L containers at the beginning of the experiment ( $t_0$ ) to estimate chlorophyll a (Chl a) concentrations. Samples were sequentially filtered through 10  $\mu$ m, 2  $\mu$ m and 0.2  $\mu$ m polycarbonate filters, which were arranged in line filter funnels. Then, filters were frozen and stored in the dark for 24 h. Chlorophyll a was extracted in 10 mL of 90 % acetone for 12-24 h and measured using Perkin Elmer LS55 fluorometer. Initial Chl a concentrations in the diluted treatments were estimated by multiplying the average undiluted initial Chl a concentrations by the dilution factor. We took 1000 mL samples from every container at the end of the experiment ( $t_f$ ) and followed the same procedure to filter and measure Chl a. In this way, we obtained Chl a measurements in every container at  $t_0$  and  $t_f$ . The picophytoplankton community was analyzed by flow cytometry (FCM) to estimate growth and microzooplankton grazing rates based on abundance measurements. Samples (1.8 mL) were taken at  $t_0$  and  $t_f$  from every container. They were preserved

with a 1 % paraformaldehyde plus 0.05 % glutaraldehyde solution and stored at -80°C. Just before the analysis, we added a solution of 1 μm fluorescent latex beads to use them as standards. Analyses were conducted using a FACSCalibur flow cytometer (Becton, Dickinson and Company) equipped with a blue (488 nm) laser. Phytoplankton were grouped and enumerated according to the side-scattered light (SSC), an indicator of cell size, the orange fluorescence (FL2, 585 nm) and red fluorescence (FL3, > 650 nm) signals. Four groups were identified: *Prochlorococcus, Synechococcus,* small picoeukaryotes and large picoeukaryotes (Calvo-Díaz and Morán 2006). If the initial cell counts in dilution treatments were very low, we estimated initial cell abundances by multiplying cell concentrations in undiluted containers by the corresponding nominal dilution (see Supplementary material).

Nano- and microphytoplankton abundances were estimated from samples taken from the carboy at the beginning of the experiments (except at S11 and S14, in which samples were taken at t<sub>f</sub>). They were preserved with the 10 % glacial acetic acid Lugol solution. Sample aliquots were maintained in the laboratory during 24 h using 25 mL Utermöhl chambers (Utermöhl 1958). The entire bottom area of the slide was examined and cells were determined up to genus or species level by using an inverted microscope. *Nitzschia spp.* at S16 was counted only in one strip and subsequently converted to cells mL<sup>-1</sup> using the appropriate conversion factor due to their high abundances.

#### 2.6. Phytoplankton growth and microzooplankton grazing rates

Exponential phytoplankton growth was assumed across the dilution treatments, resulting in apparent growth rate (r) equal to:

$$212 r = t^{-1} \ln(P_t P_0^{-1})$$

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where t is the incubation time,  $P_0$  is the initial phytoplankton biomass (Chl a biomass or cell abundance) and  $P_t$  is the biomass at the end of the incubation. Commonly, phytoplankton growth rate  $(\mu)$  and microzooplankton grazing rate (m) are estimated with a linear regression analysis of r against dilution factor (f), where u is the intercept and m is the slope (Landry and Hassett 1982). Here we estimate  $\mu$  and m by fitting mixed models using the *lme4* R package (Bates et al. 2013). We included the dilution factor as a covariate, province and phytoplankton group (phytoplankton size fraction or flow cytometry group) as fixed factors and station as a random factor (see Supplementary material). This allowed us to simultaneously estimate  $\mu$  and m for every phytoplankton group and station and mean  $\mu$  and m for all phytoplankton groups and provinces. Additionally, the parameters are estimated taking into account the hierarchical organization of the data (Gelman and Hill 2007), which is not accounted for when conducting separate linear regressions for every experiment (the method commonly employed). In this way, all the information contained in the data set is considered when estimating the rates in the different experiments, and greater weight is given to experiments with less uncertainty. This provides more robust estimates, which are less influenced by extreme results or potential errors. Additionally, the correlation among stations from the same province, i.e. the non-independence of the data, is taken into account. For all those reasons, and considering our interest in estimating not only the rates  $(\mu \text{ and } m)$  for each experiment but also the mean rates for each province and group, we find mixed models a more appropriate method than averaging  $\mu$  and m for every province and group from the parameters obtained by fitting a linear regression in each experiment. Furthermore, we performed model selection followed by model averaging, recommended when more than one model has substantial support, to obtain a more robust estimate of the parameters and a more stabilized inference (Burnham and

Anderson 2002) (see Supplementary material). This multimodel inference approach also enabled us to estimate the relative importance of each variable by adding the scaled AICc weights (see Supplementary material) of all the models within the 95 % confidence set of models where the variable of interest was included (Burnham and Anderson 2002). In our case, we obtained the relative importance of station, province and phytoplankton group as predictors for phytoplankton growth rate and microzooplankton grazing rate (i.e. interaction between predictors and dilution factor). Finally, to check the validity of our approach we compared the rates obtained by using mixed models and model averaging with the ones obtained by fitting separate linear regressions to each experiment.

# 2.7. Multivariate analyses of relations between nutrients, phytoplankton community structure and growth

Multivariate statistics were used to analyze differences among stations with regard to phytoplankton community taxonomic structure, phytoplankton community size structure and size fractionated phytoplankton growth rates at the depths of maximum phytoplankton activity. In addition, we related differences among stations in those properties with the nutrient availability at both the subtropical gyre and the within-province spatial scale.

To analyze phytoplankton taxonomic structure, we considered the abundances of 31 different genera (identified using optical microscope and FCM) and two non-taxonomic groups (small and large picoeukaryotes). These abundances were standardized by dividing each value by the range of abundances of the corresponding group, to counteract the higher contribution of the most abundant groups to the dissimilarities among stations (Quinn and Keough 2002). Those dissimilarities were estimated using

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Bray-Curtis measure. Then, we performed NMDS using SMACOF. Subsequently, we conducted Permutational Multivariate Analysis of Variance [PERMANOVA, (Anderson 2001)] using the *vegan* package (Oksanen et al. 2013) to estimate the relationship (R<sup>2</sup>) between the differences in taxonomic community structure among stations and the availability of nutrients using the following sea water properties: dissolved inorganic nitrogen (DIN,  $NH_4^+ + NO_3^- + NO_2^-$ ), silicates and accumulated  $N^2$ in the 100 m below the sampling depth, which indicates the strength of stratification and, consequently, was used as a proxy for nutrient inputs from deeper waters. DIN and silicate measurements were from the same depth as the phytoplankton samples or the closest depth for which nutrient samples were available. Phosphates were not included in the analysis because of their high correlation with DIN at those depths (r = 0.99). PERMANOVA was conducted without including and including province as a predictor, which removes the effects of province, in order to estimate the variances explained by the covariates at the subtropical gyre and at the within-province spatial scales, respectively. By including province as a predictor we also estimated the variance explained by province. In addition, we included the interaction between province and different covariates, which highlights the differences in magnitude or direction of the relationship among provinces. We conducted the same analyses with phytoplankton community size structure (using size fractionated Chl a) and size-fractionated growth rates (obtained from dilution experiments), although in these cases we employed Euclidean distances to generate the dissimilarity matrices. Finally, we explored the relationship of community structure (taxonomic and size) with growth rate at the two scales considered in our research. For the subtropical gyre scale, we estimated the correlation between dissimilarity matrices. For the province scale, we fitted a linear mixed model that assessed the relationship between size-fractionated Chl

a and growth rate in each province. The model included  $\mu$  as a dependent variable, centered Chl a as a covariate, province as a fixed factor and size fraction as a random factor (see Supplementary material for further details). Chl a concentrations were centered by subtracting the mean Chl a value for each phytoplankton size fraction in each province. This analysis allows us to consider the different size fractions simultaneously. We fitted a similar model using the size-fractionated m as a dependent variable. This analysis can help us disentangle the role of grazing in nutrient regeneration and in the relaxation of phytoplankton competition for nutrients (Cooper 1973; Bergquist and Carpenter 1986).

# 3. Results

#### 3.1. Sea water properties and classification of the stations

Visual inspection of vertical profiles and satellite images revealed general patterns in the evolution of the sea water properties along the transect (Supplementary material Figs. 1 and 2). This was further corroborated using the NMDS ordination of the seawater properties, which enabled us to classify the stations into their corresponding provinces and sub-provinces. S2 to S6 have similar values on axis 1; we classified them as stations from NATR (Supplementary material Fig. 3). S7 to S16 were classified as NAST stations. The boundary between both NAST sub-provinces, NAST-W and NAST-E, was located between S11 and S12, coinciding with the topography of the Mid Atlantic Ridge (Fig. 1). For a further description see Supplementary material.

#### 3.2. Phytoplankton abundances and community structure

Differences in the taxonomic structure of phytoplankton communities along the transect corresponded with provinces defined by Longhurst (2007). Indeed, province explained a large amount of the variance in community structure among stations ( $R^2 = 0.43$ ,

PERMANOVA), which might reflect the differences in nutrient availability (see below). 310 NATR stations formed a well-defined group (Fig. 2A) characterized by low abundance 311 312 of Synechococcus, small picoeukaryotes, large picoeukaryotes and diatoms (Fig. 2B). The abundance of most groups increased in the NAST-W stations, with the exception of 313 dinoflagellates, which exhibited homogeneous abundances along the transect, and 314 Prochlorococcus (although Prochlorococcus reached its maximum concentration in 315 S11). Most NAST-E stations showed higher abundances of large picoeukaryotes and 316 317 diatoms than the NAST-W stations (Fig. 2B), which led to their distinction in the NMDS analysis (Fig. 2A). Our results showed a diatom bloom in S16 dominated by 318 Nitzschia delicatissima, with low abundances of Prochlorococcus and Synechococcus 319 320 (Fig. 2B, Supplementary material Table 9). This differentiated the S16 community from the rest of the NAST-E stations. Hence, S16 was possibly located at the boundary 321 322 between NAST-E sub-province and the North Atlantic Drift Province (NADR) (See Longhurst 2007). 323

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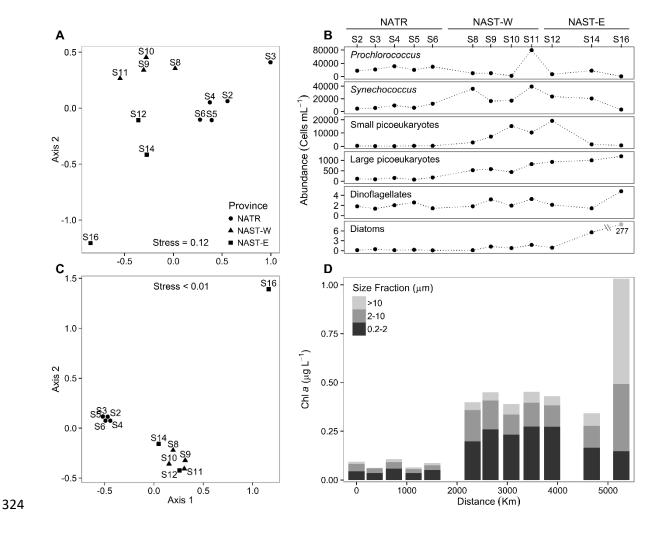


Fig. 2 Taxonomic composition and size structure of the phytoplankton community. (A) Two-dimensional configuration of stations obtained from multidimensional scaling (NMDS) for phytoplankton community taxonomic structure. NMDS stress, a measure of the goodness of fit, is indicated. (B) Abundances of Synechococcus, small picoeukaryotes, large picoeukaryotes, Prochlorococcus, dinoflagellates and diatoms in the stations where dilution experiments were performed. Note the different scales of the abundances. Diatom abundance at S16 is out of the scale represented; its value is showed below the dot. (C) Two-dimensional configuration of stations obtained from the NMDS for phytoplankton size structure. (D) Size fractionated Chl a concentrations in the stations where dilution experiments were conducted. Unsurprisingly, the phytoplankton community's size structure along the transect closely resembled the taxonomic structure of the community (Fig. 2C), with a correlation of r=0.79 between dissimilarity matrices. Once again, province was a determining factor in explaining the variance (R<sup>2</sup>= 0.53, PERMANOVA). NATR stations were clustered

together (Fig. 2C) mainly due to their low Chl a concentrations in all three size fractions

(Fig. 2D). NAST stations were grouped close together, with the exception of S16. They shared high Chl a concentrations caused by the aforementioned increases in phytoplankton abundance. S16 appeared as an outlier in the NMDS plot due to high concentrations of Chl a in the medium and large phytoplankton size fractions.

# 3.3. Phytoplankton growth and microzooplankton grazing rates

# 3.3.1. <u>Chl *a* analysis</u>

Net growth rates derived from Chl a measurements were analyzed using different models to estimate phytoplankton growth and microzooplankton grazing rates. Phytoplankton growth rates ranged between  $0.00 \pm 0.39 \, d^{-1}$  and  $1.19 \pm 0.18 \, d^{-1}$  for the large phytoplankton size fraction in S16 and the medium size fraction in S6, respectively (Fig. 3; Supplementary material Fig. 4). The range of grazing rates was narrower, between  $0.32 \pm 0.25 \, d^{-1}$  at S16 and  $0.74 \pm 0.26 \, d^{-1}$  at S4. In fact, the variation of phytoplankton growth rate was higher than the variation of microzooplankton grazing rate among provinces (Fig. 4), size fractions within each province (Fig. 4), stations and among size fractions within each station (Fig. 3; Supplementary material Fig. 4; see below).

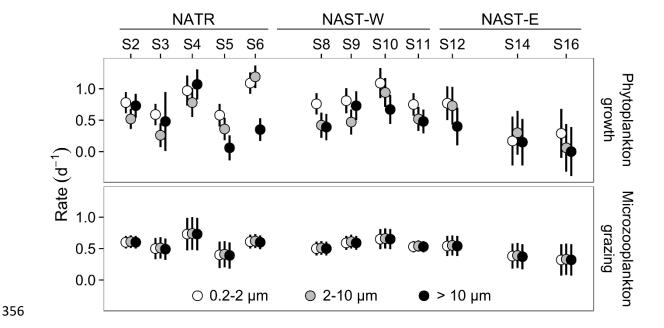


Fig. 3 Phytoplankton growth and microzooplankton grazing rates for each station and size fraction. Error bars represent 95% confidence intervals. Color indicates the phytoplankton size fraction. Geographical distance between stations has been kept.

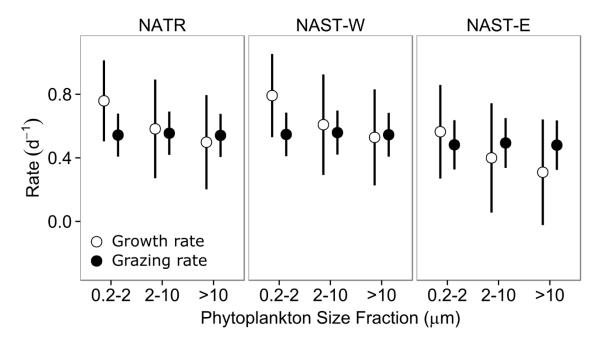


Fig. 4 Mean phytoplankton growth and microzooplankton grazing rates for each phytoplankton size fraction and province estimated from model averaging with models included in the 95% confidence set of models. Bars represent standard deviation.

Mean phytoplankton growth rates were similar in NATR and NAST-W and lower in NAST-E (Fig. 4). Mean grazing rates also decreased in NAST-E, but in a less

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pronounced manner than mean growth rates (Fig. 4). Mean phytoplankton growth rates diminished with the phytoplankton size class (Fig. 4). Nevertheless, mean grazing rates were almost the same for all size fractions (Fig. 4). In summary, province affected both phytoplankton growth and microzooplankton grazing, although this effect is less pronounced in grazing rates. Conversely, size fraction only affects phytoplankton growth rate. These effects were confirmed by measurements of relative variable importance by using scaled AICc weights: the sum of scaled AICc weights of models that included province and the interaction between dilution factor and province (dilution x province) in the fixed structure was 0.57 and 0.33, respectively. Nevertheless, in the case of size fraction and the interaction between dilution factor and size fraction that sum was 0.99 and 0.13, respectively. Thus, the differences in mean phytoplankton net growth rates among provinces and especially among size fractions within each province were mainly determined by the differences in growth rates rather than by differences in microzooplankton grazing rates. The mentioned effect of province on the size fractionated phytoplankton growth rate was also revealed by the PERMANOVA analysis ( $R^2 = 0.28$ . See also Fig. 5).

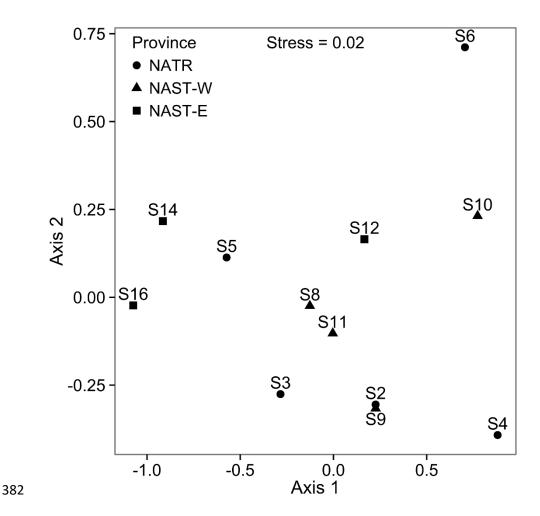


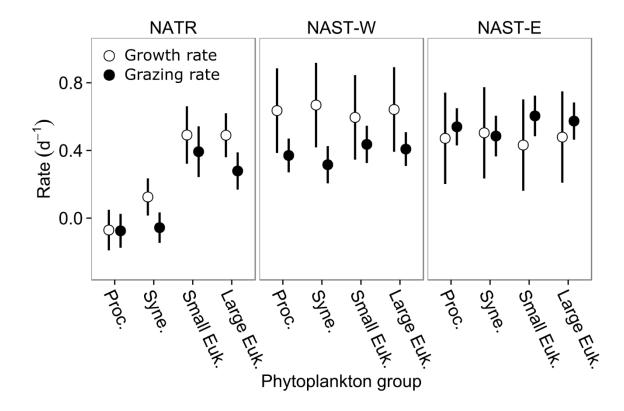
Fig. 5 Two-dimensional configuration of stations obtained from the non-metric multidimensional scaling (NMDS) analysis conducted with size fractionated phytoplankton growth rates. NMDS stress is also indicated.

The higher variability observed for phytoplankton growth rate than for microzooplankton grazing rate among stations (Fig. 3, see standard deviations in Fig. 4) and among size fractions within each station (Fig. 3) was also revealed by the sum of scaled AICc weights. For models including a varying coefficient for the intercept (growth) and the slope (grazing) the sum of scaled AICc weights were 1.00 and 0.65, respectively. In these models, size fraction was included in the coefficient for the intercept but not for the slope (Supplementary material Table 3). Again, differences in phytoplankton net growth rates, both among stations and size fractions within each station, would be mainly caused by differences in growth rates rather than by differences in microzooplankton grazing rates. Province does not greatly affect the

variability (standard deviation) among stations of both rates ( $\mu$  and m) (Fig. 4), in fact it was not included in the random structure of any of the models within the 95 % confidence set (Supplementary material Table 3).

# 3.3.2 Flow cytometry analysis

We estimated growth and microzooplankton grazing rates for picophytoplankton groups in the dilution experiments from FCM counts. As expected, the observed intercepts (phytoplankton growth rates) and slopes (microzooplankton grazing rates) were positive and negative, respectively, except in the case of cyanobacteria in NATR, where the contrary occurred (Supplementary material Fig.5). This effect on cyanobacteria has been previously reported in other dilution experiments, where it has been mainly attributed to the effect of trophic cascades (see Calbet and Saiz 2013 and references therein). The highest picophytoplankton growth and microzooplankton grazing rates were found in NAST-W and NAST-E sub-provinces, respectively (Fig. 6). Within NATR, growth and grazing rates were higher for picoeukaryotes than for cyanobacteria, whereas within NAST they were similar for the four picophytoplankton groups analyzed (Fig. 6). Additionally, in NAST-W picophytoplankton growth rate was higher than microzooplankton grazing rate; this difference was lower in the other provinces. Finally, we once again observed higher variations among stations in growth rates than in microzooplankton grazing rates (Fig. 6).



picophytoplankton group and province estimated from model averaging with models included in the 95% confidence set of models. Bars represent standard deviation.

We analyzed changes in FL3 and SSC signals between  $t_0$  and  $t_f$  to detect potential artifacts caused by dilution, diel growth cycles (some experiments lasted less than 24 h) or photoacclimation processes that might affect Chl a and FCM estimates of growth and grazing rates. We found no evidence of an effect of dilution treatment on relative FL3. Nevertheless, we observed positively correlated increases in mean FL3 and SSC signals of *Synechococcus* (estimated for each station) within the NATR province (r = 0.78, n = 5). Experiments in NATR lasted 21h and started when cells have just finished division (Table 1); therefore FL3 and SSC signals showed values near the lowest trough of *Synechococcus* light-dark growth cycle (Sweeney and Borgese 1989; Olson et al. 1990; Jacquet et al. 1998). However, experiments ended when cells were still dividing and the values of those signals were closer to the light-dark cycle peak. While we can not

Fig. 6 Mean phytoplankton growth and microzooplankton grazing rates for each

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discard the occurrence of photoacclimation processes, the estimates of *Synechococcus* growth rates from FCM counts could be underestimated.

# 3.3.3 Suitability of the method

In the case of the Chl a analysis, we compared the rates obtained by using mixed models and model averaging with those obtained by fitting separate linear regressions to each station and size fraction, the method traditionally employed (Supplementary material Fig. 4 and Table 7). Both approaches exhibited similar rates with only few exceptions. These exceptions occurred in experiments that showed a pattern far from norm, i.e. far from the rest of experiments, such as the 0.2-2 µm size fraction at S5, S6 and S8 or > 10 µm size fraction at S4 and S5 (Supplementary material Fig. 4 and Table 7). In those experiments, mixed models, by considering the entire data set and not only the data of the specific experiment, offered a more robust approach and a more stabilized inference, which was less influenced by extreme results or by potential errors occurred at specific experiments. Additionally, mixed models enabled the estimation of the rates for some factor levels without data (> 10 µm at S3) and improved the precision of the estimates in experiments with fewer observations (e.g. 2-10 µm at S3). In this way, the confidence intervals of the rates obtained by our approach were in general narrower than the ones obtained by fitting linear regressions (Supplementary material Table 7). The mean rates for each province and size fraction estimated from our approach and from averaging the rates obtained by fitting linear regressions to each experiment were in general similar too, although some differences were observed for both phytoplankton growth and microozooplankton grazing rate (Supplementary material Table 8), mainly in NAST-E.

It is worth emphasizing that the higher variability and differences observed for phytoplankton growth rate than for microzooplankton grazing rate among provinces, stations and size fractions were also observed when those rates were estimated by fitting linear regressions for every experiment (Supplementary material Table 8). Nevertheless, those variabilities were in general lower when they were estimated by following our approach, especially in the case of the microzooplankton grazing rate among size fractions within each station (Supplementary material Table 7).

# 3.4. Phytoplankton community properties and nutrient availability

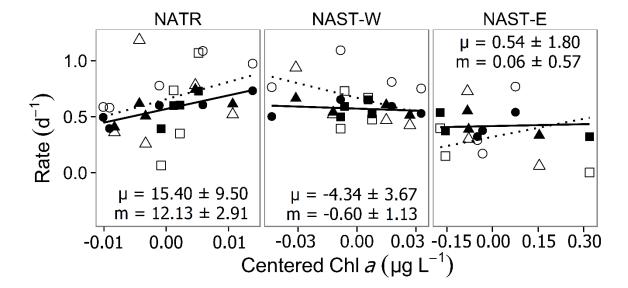
The PERMANOVA analysis revealed an effect of DIN and cumulative  $N^2$  on differences in taxonomic and size structure of phytoplankton community at the subtropical gyre spatial scale ( $R^2 \ge 0.16$ , Table 2). Explained variances were lower for silicate concentrations ( $R^2 \le 0.11$ , Table 2). All those relationships were lower at the within-province spatial scale (after removing province effects) ( $R^2 \le 0.08$ , Table 2). This means that differences in phytoplankton community structure are mainly driven by differences in nutrient concentrations and cumulative  $N^2$  among provinces rather than within province. Nevertheless, the high variance explained by the interaction between province and silicate concentrations, together with the high abundance of diatoms and the low silicate concentrations observed in S16, suggested that silicate concentrations were strongly related with community structure in NAST-E. We repeated the analysis using relative standardized abundances of phytoplankton (standardized abundances divided by the sum of all the standardized abundances of each station), obtaining very similar results (data not shown).

 Table 2. Variances explained (R<sup>2</sup>) for the relationships between phytoplankton community properties and the different covariates obtained by the PERMANOVA analysis. Rows show the covariates for which the relationships were estimated. Columns show the different community properties analyzed. Sub-columns "Subtropical" and "Within-prov." pointed out the spatial scale at which relationships were estimated. Subtropical: the relationships were obtained considering the effects of the covariates at a subtropical gyre spatial scale. Within-prov: the relationships were estimated after removing the effects of province. Sub-column "Interaction" indicates the variance explained by the interaction between the covariates and province (it was not estimated for models including the three covariates because the number of parameters was too high).

Covariate	Phytoplankton community taxonomic structure			Phytoplankton community size structure			Phytoplankton community size fractionated growth		
	Subtropical	Within-prov	Interaction	Subtropical	Within-prov	Interaction	Subtropical	Within-prov	Interaction
DIN	0.19	0.04	0.09	0.16	0.01	0.07	0.02	0.07	0.19
Silicates	0.09	0.08	0.17	0.11	0.07	0.37	0.13	0.08	0.13
Cum. N2	0.21	0.05	0.10	0.24	0.00	0.02	0.05	0.10	0.11
DIN+Silicates+Cum.N2	0.47	0.24		0.52	0.20		0.32	0.34	

Contrary to phytoplankton community structure measurements, phytoplankton growth rates were not influenced by either DIN or cumulative  $N^2$  at the subtropical gyre spatial scale ( $R^2 \le 0.05$ , Table 2). Thus, differences in DIN and cumulative  $N^2$  among provinces did not drive the differences in size fractionated phytoplankton growth rates. In fact, stations from NATR showed size fractionated phytoplankton growth rates similar to those observed at stations from NAST despite the general differences in DIN and cumulative  $N^2$  between the two provinces (Table 1, Fig. 4, Supplementary material Fig.1). The relationship between the differences in phytoplankton growth rates and silicate concentration was stronger, although it was highly influenced by S16; the exclusion of S16 from the analysis reduced the explained variance from 0.13 to 0.05. Conversely, the relationship between differences in phytoplankton growth and both DIN and cumulative  $N^2$  increased after removing the effects of the differences among provinces, indicating an effect of those covariates on phytoplankton dynamics at the within-province spatial scale, albeit a weak one ( $R^2 \ge 0.07$ , Table 2). Sure enough,

according to the explained variances for the interaction term, the relationship between 502 the differences in phytoplankton growth and nutrient availability differed between 503 504 provinces (Table 2). We obtained similar results when we repeated the analysis using phytoplankton growth rates estimated by fitting separate linear regressions for each 505 station and size fraction (data not shown). 506 507 Differences in size fractionated phytoplankton growth rates were uncoupled from differences in phytoplankton community structure at the subtropical gyre spatial scale. 508 509 We observed low correlations between the dissimilarity matrix of size fractionated phytoplankton growth rates and the dissimilarity matrices of both community 510 taxonomic structure and size structure (r = 0.13 and r = 0.24, respectively). However, 511 512 Chl a concentrations in all size fractions were positively correlated with the size fractionated growth and grazing rates within NATR (Fig. 7; Supplementary material 513 514 Table 10). In contrast, the relationships were weaker, and in some cases negative, in 515 both NAST sub-provinces (Fig. 7, Supplementary material Table 10).



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Fig. 7 Relationships between centered Chl a and both size fractionated phytoplankton growth ( $\mu$ ) and microzooplankton grazing rates (m) in the different provinces. Note the different scales of the x axes. White symbols indicate the phytoplankton growth rate and black symbols the microzooplankton grazing rate. Shapes signify the phytoplankton size fractions: 0.2-2  $\mu$ m size fraction (circles), 2- 10  $\mu$ m size fraction (triangles) and > 10  $\mu$ m size fraction (squares). Lines indicate the linear fit for the relationships between  $\mu$  and centered Chl  $\mu$  (dotted) and  $\mu$  and centered Chl  $\mu$  (solid).  $\mu$  is the slope (mean  $\mu$  standard error) of the relationship between phytoplankton growth rate and centered Chl  $\mu$  is the slope (mean  $\mu$  standard error) of the relationship between microzooplankton grazing rate and centered Chl  $\mu$ .

# 4. Discussion

We estimated size fractionated phytoplankton growth and microzooplankton grazing rates along a transect that covered a variety of conditions, which mirrored the geographical partition of the North Atlantic proposed by Longhurst (2007). Our results revealed that phytoplankton growth rate showed higher variability microzooplankton grazing rate among stations, provinces and size fractions. Phytoplankton community structure differed across provinces and was associated with nutrient availability at the subtropical gyre spatial scale. However, differences in phytoplankton growth rate showed a weak relationship with nutrient availability at that subtropical gyre spatial scale, being stronger at the within-province spatial scale.

Differences in phytoplankton growth rate and differences in community structure were only weakly correlated, although we observed a positive relationship between size-fractionated growth rate and size-fractionated Chl a within one of the provinces (NATR). Below, we discuss potential mechanisms for the observed variations in phytoplankton growth and microzooplankton grazing rates. Then, we discuss the relationship between nutrient availability, phytoplankton structure and phytoplankton dynamics at the two spatial scales considered.

# 4.1. Suitability of the statistical method

By fitting mixed models and conducting model averaging we took into account the hierarchical organization of the data and achieved a robust inference, estimating both specific rates for each station and size fraction and average rates for each province and size fraction. In general, the rates estimated by our approach were close to the ones obtained by fitting linear regressions for each experiment. The observed differences between both methodologies were mainly caused by the model selection based on AICc, which prevents overfitting by dealing with the trade-off between the goodness of fit and the complexity (number of parameters) of the model (Burnham and Anderson 2002), and the subsequent model averaging. Also, those differences arose due to the use of mixed models: when estimating the rates for a particular station and size fraction, mixed models take advantage of the information contained in other stations and size fraction, mixed models assign a different weight to each experiment (depending on the information it contains). This does not occur when rates are estimated from the fitting of separate linear regressions for each experiment.

Our approach, both through using mixed models and model averaging, captured and unmasked the main patterns within the data without lead to overfitting. It enabled the detection of one of our major results, the higher variability in phytoplankton growth rate among provinces, stations and size fractions than in microzooplankton grazing rate, which could have been overlooked using traditional methods.

Based on our experience and the extensive literature on the use of mixed models (e.g. Gelman and Hill 2007), we encourage their application in future studies that aim to estimate mean rates in similar locations, depths or times, or studies focused on the variability of rates. Also, by conducting model selection and multimodel inference a more stable inference, i.e. more robust estimates of the rates, can be obtained. Furthermore, this procedure provides measurements on the importance of different predictors in explaining both the variability in phytoplankton growth and microzooplankton grazing rates (Burnham and Anderson 2002; Johnson and Omland 2004).

#### 4.2. Phytoplankton growth and microzooplankton grazing rates

The variability in phytoplankton growth rate among provinces, stations and size fractions was higher than the variability in microzooplankton grazing rate. Greater differences among habitats for phytoplankton growth rate than for microzooplankton grazing rate were previously reported by Calbet and Landry (2004). Thus, differences in phytoplankton net growth rate among provinces, stations and size fractions were mainly determined by differences in the phytoplankton growth rate rather than by differences in the microzooplankton grazing rate. Microzooplankton grazing is considered one of the main drivers of phytoplankton mortality in subtropical oceans (Calbet and Landry 2004), this could entail that phytoplankton growth rate rather than mortality rate is

driving the differences in phytoplankton net growth rate among subtropical areas or groups. Moreover, our present results on the high growth rates of the smallest size fraction, coupled with information on the low sedimentation and mortality rate due to mesozooplankton grazing found in the literature (Kiørboe 1993), would imply that the relative contribution of the small size fraction to the total phytoplankton biomass was increasing in most stations. Determining if in fact growth rate has a greater contribution to the variability of the phytoplankton net growth rate than mortality rate will be a crucial step in understanding phytoplankton dynamics, including phytoplankton blooms. Future studies analyzing the variability of the growth and all the mortality sources of phytoplankton (including viral lysis, mesozooplankton grazing and sedimentation in addition to microzooplankton grazing) are required to confirm this hypothesis and extrapolate it to other seasons or areas.

Phytoplankton growth rate tended to decrease as phytoplankton size increases in the three provinces, in agreement with the studies that analyzed the relationship between growth and size (Banse 1976; Tang 1995). The observed pattern could be due to a decrease in the maximum phytoplankton growth rates as phytoplankton size increases (Chisholm 1992; Edwards et al. 2012), although recent studies suggest that the highest growth rates can be found in species of intermediate size (c. 100 μm³, 5.76 μm spherical diameter) (Marañón et al. 2013). Our results contrast with research carried out in NAST-E in autumn or in other areas using the dilution technique, where large phytoplankton grew as fast or faster than small phytoplankton (Olson and Strom 2002; Calbet et al. 2008; Cáceres et al. 2013). In those cases, functional traits commonly more developed in large phytoplankton and advantageous when nutrients are supplied in an intermittent way, such as the maximum rate of nutrient uptake, the capacity to store

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nutrients or the ability to perform vertical migration, would influence the growth of phytoplankton populations (Reynolds 2006; Litchman et al. 2007).

According to our results, the microzooplankton grazing rate showed little differences among size fractions. This result contrasts with previous research, which stated large sizes provide phytoplankton protection against the predation by microzooplankton, thus microzooplankton grazing rates are expected to be lower for the large phytoplankton size fraction (Kiørboe 1993). Nevertheless, high grazing rates for the large phytoplankton size fraction have been previously observed in the subtropical Northeast Atlantic (Cáceres et al. 2013). The microzooplankton grazing rate depends on the ratio between phytoplankton biomass grazed and phytoplankton biomass. Therefore, if this ratio is constant across size fraction similar grazing rates are expected. In this way, the functional and numerical responses of predators to the abundance of preys would promote the association between phytoplankton biomass and phytoplankton biomass grazed. The fact that zooplankton might prey on different size fractions of phytoplankton, although with different efficiency (Hansen et al. 1994), could also contribute to equalizing grazing rates among size fractions. On the contrary, the specialization of grazers and the differences in their biology can lead to different grazing rates on each phytoplankton size fraction, as it has been reported for other seasons or places (Olson and Strom 2002; Calbet et al. 2008; Cáceres et al. 2013).

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#### 4.3. Nutrients and phytoplankton community structure and dynamics

The match between phytoplankton community structure, DIN and cumulative  $N^2$  at the 629 630 subtropical gyre scale could be caused by the selection of taxa with functional traits best suited to exploit the low nutrient concentrations in NATR (Litchman et al. 2007: Moore 631 et al. 2008; Edwards et al. 2013). In fact, the abundance of *Prochlorococcus*, probably 632 633 the most nutrient stress tolerant phytoplankton species (Reynolds 2006; Brun et al. 2015), was particularly high in NATR. That match is favored by the strong constraint 634 that nutrient availability imposes on phytoplankton in subtropical areas (Reynolds 2001). Differences in taxonomic composition and functional traits of phytoplankton 636 communities between biogeochemical provinces would lead to differences in growth-637 638 nutrient responses, promoting the weak relationship observed between phytoplankton growth and nutrients at a subtropical gyre scale. This situation was widely reported in 639 studies focused on phytoplankton at a species level instead of community (e.g. Grover 640 641 1997); species with different functional traits may have similar growth rates under different nutrient concentrations and vice versa. Even populations of the same species 642 may mitigate the effects of low nutrient concentrations due to phenotypic plasticity or 643 genotype diversity and selection in traits affecting nutrient acquisition (Martiny et al. 644 2006; Van Mooy et al. 2009; Bonachela et al. 2011; Lomas et al. 2014; Biller et al. 645 2015). This would highlight the importance of functional diversity in maintaining and stabilizing phytoplankton growth at the subtropical gyre spatial scale, as it was 647 previously determined by Díaz and Cabido (2001) for natural communities and 648 ecosystem functioning. Thus, the growth rate of phytoplankton communities in 649 oligotrophic subtropical gyres could be higher than the expected from the low nutrient 650 concentrations (Cullen et al. 1992). 651

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Other factors may contribute to the weak relationship between nutrient availability and phytoplankton growth at the subtropical gyre spatial scale, compensating for the low nutrient availability in NATR. Temperature stimulates chemical processes, metabolic reactions and phytoplankton growth (Eppley 1972; Raven and Geider 1988; Moore et al. 1995) and, as in other studies (Kamykowski and Zentara 1986), was negatively correlated with nutrients (Supplementary material Fig. 1). In addition, the large area encompassed by oligotrophic open ocean ecosystems like the NATR, together with the previous existence of stratified oceans (Falkowski and Oliver 2007), would favor the selection of species and ecotypes adapted to low nutrient concentrations. Furthermore, the stability of these areas could promote the match as well as the acclimation of phytoplankton communities to low nutrient concentration (see Venrick 1990). This match would be lower in areas with stronger seasonal cycles like NAST-E (see Longhurst 2007). Also, quick nutrient regeneration carried out by grazers and patches of high nutrient concentrations in these areas could increase nutrient availability for phytoplankton (Goldman 1984). Finally, differences in light conditions might also affect growth rate patterns and consequently their relationship with nutrients, although the careful selection of sampling depths would reduce that possibility. Silicates displayed a stronger relationship with differences in phytoplankton growth rate at a subtropical gyre spatial scale than DIN and cumulative  $N^2$ . This relationship was mainly influenced by the diatom bloom in S16, which prompted the depletion of silicates. In fact, considering the low silicate concentrations and the notably higher than 1 N:Si ratio, a common N:Si ratio for diatoms (Brzezinski 1985), diatoms growth could be limited by Si in S16, as it was reported at higher latitudes (Turner et al. 1998; Longhurst 2007). This explains why phytoplankton growth rates of the medium and large size fraction in S16 were lower than in contiguous stations and those reported in

other studies (Calbet and Landry 2004; Marañón 2005). These particularities in the biochemical properties of S16 could indicate that it was located in the frontier between NAST-E and the North Atlantic Drift Province (NADR), where spring phytoplankton blooms are more marked (Longhurst 2007). The diatom bloom could be responsible for the lower grazing rates observed in S16; the increase in phytoplankton biomass would have not been counterbalanced yet due to the lag in the zooplankton response. Similarly, lower grazing rates associated to high phytoplankton biomasses have been previously reported in other areas (Olson and Strom 2002).

The drivers for community structure differed among scales. Contrary to what was absented at the subtractical graze and a grazely provided at the subtractical grazely provided grazely pr

observed at the subtropical gyre scale, DIN and cumulative  $N^2$  had little influence on the community structure at the within-province spatial scale, possibly caused by the fickle nature of nutrient differences at this scale (Johnson et al. 2010). This would hinder the match of the phytoplankton community structure to nutrient availability, or restrict that match to very short time periods, making it difficult to detect. In fact, the high concentration of DIN and silicates in S10, associated with the presence of a negative sea level anomaly which entailed the ascent of enriched subsurface waters, did not cause any marked increase in the abundance of any phytoplankton group. Nevertheless, differences in phytoplankton community structure associated to fleeting nutrient inputs have been reported for subtropical areas (McAndrew et al. 2007; McGillicuddy et al. 2007; Brown et al. 2008). That weak relationship between differences in community structure and both DIN and cumulative  $N^2$  at the within-province spatial scale would imply that phytoplankton communities within each province would exhibit similar functional traits associated with nutrient acquisition and growth. Thus, we would expect a similar response to nutrients in these communities. This promoted the emergence of the relationship observed between both DIN and cumulative  $N^2$  and differences in size 702 fractionated phytoplankton growth at a within-province spatial scale, which does not occur at the larger subtropical gyre scale. In this way, phytoplankton growth rates 703 704 estimated from both Chl a concentrations and FCM counts were high at S10, coinciding 705 with the mentioned enhanced concentration of DIN and silicates. Studies in the subtropical North Atlantic relating phytoplankton growth and nutrients at a within-706 707 province scale are scarce, although increases in phytoplankton growth linked to nutrient 708 inputs associated to mesoscale features has been suggested (McGillicuddy et al. 1998). The uncoupling between phytoplankton community structure and growth at a 709 710 subtropical gyre spatial scale, possibly favored by the response of those properties to 711 nutrient availability, was reverted within the NATR province. The positive relationships 712 observed between size fractionated  $\mu$  and centered Chl a in NATR could be promoted 713 by the also positive relationship found between size fractionated m and centered Chl a. 714 Higher grazing rates when phytoplankton biomasses are higher entail higher nutrient 715 regenerations (Bergquist and Carpenter 1986; Sterner 1986) and avoid increases in 716 phytoplankton biomass that would lead to nutrient scarcity. The similar relationships with centered Chl a of both phytoplankton growth and microzooplankton grazing rates 717 imply a coupling between growth and grazing, which has been previously reported in 718 719 oligotrophic subtropical gyres (e.g. Quevedo and Anadón 2001) and argued to explain 720 the high phytoplankton growth rates measured in those areas (Goldman 1984). 721 In conclusion, the relationships between nutrient availability and both the differences in 722 phytoplankton community structure and growth were subject to change according to the scale at which they were analysed. Therefore, it is crucial to consider the spatial scale in 723 724 the study of phytoplankton ecology (Levin 1992). Furthermore, the relationship between nutrient availability and phytoplankton growth rate is particularly complex. Here, we 725 have observed the impact of scale and phytoplankton community structure on this 726

- 727 relationship. At the subtropical gyre spatial scale, we observed a weak relationship
- between the differences in phytoplankton growth and nutrient availability, which was
- 729 promoted by the match between phytoplankton community structure and nutrient
- 730 availability. This highlights the importance of taking into account the structure of
- biological communities when analysing their functioning and response to changes.

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# **Supplementary material**

Material and methods

#### Phytoplankton growth and microzooplankton grazing rates.

We fitted mixed models to estimate phytoplankton growth ( $\mu$ ) and microzooplankton grazing rates (m). These were based on the linear regression model proposed by Landry and Hasset (1982), which estimates  $\mu$  and m from phytoplankton apparent growth rate (r) and the dilution factor (f):

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$$r = \mu + mf$$

This model would allow us to estimate  $\mu$  and m for each phytoplankton group (phytoplankton size fraction or flow cytometry group) in each station by running it separately. However, we were also interested in estimating size fractionated  $\mu$  and m for each province. Therefore, we included effects of province, phytoplankton group and station (as random factor) in the previous model, obtaining the following two global mixed models (note the different random structures):

$$r_{ijkl} = \mu_0 + \mu_{Prov} + \mu_{group} + \mu_{Prov.\ group} + \alpha_{station(Prov.)} + (m_0 + m_{Prov.} + m_{group} + m_{Prov.\ group} + \beta_{station(Prov.)})f + e_{ijkl}$$

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$$r_{ijkl} = \mu_0 + \mu_{Prov.} + \mu_{group} + \mu_{Prov. group} + \alpha_{station, group} + (m_0 + m_{Prov.} + m_{group} + m_{Prov. group} + \beta_{station, group}) f + e_{ijkl}$$

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$$(\alpha_{station(Prov.)}, \beta_{station(Prov.)}) \sim N(0, \sum_{station(Prov.)})$$

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$$(\alpha_{station, group}, \beta_{station, group}) \sim N(0, \sum_{station group})$$

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$$e_{ijkl} \sim N(0, \sigma^2).$$

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Where  $r_{ijkl}$  is the net growth rate when province = province<sub>i</sub>, group = group<sub>i</sub>, station= station<sub>k</sub> and dilution factor  $(f) = f_l$ .  $\mu_0$  is the intercept of the reference level.  $m_0$  is the slope of the reference level. Province and group are fixed effects on both intercept  $(\mu_{Prov.}, \mu_{group})$  and slope  $(m_{Prov.}, \mu_{group})$  $m_{group}$ ), whose interaction is also considered ( $\mu_{Prov.\ group}$ ,  $m_{Prov.\ group}$ ). Station is a random effect also acting on both intercept ( $\alpha$ ) and slope ( $\beta$ ), being nested in province ( $\alpha_{station(Prov.)}$ ,  $\beta_{station(Prov.)}$ ) or interacting with phytoplankton group ( $\alpha_{station, group}$ ,  $\beta_{station, group}$ ) depending on the global model considered. This allowed the intercepts and slopes to vary between stations, estimating at the same time different variances for intercepts and slopes depending on the province or the phytoplankton group. Because of the relative low number of observations, we cannot include in the same model random structures considering province and group. Random coefficients follow a normal distribution with mean equal 0 and a variance which is estimated by model fitting. In the case of size fractionated Chl a data,  $\sum_{station\ Prov.}$  and  $\sum_{station\ group}$  are 6 x 6 symmetric covariance matrices containing each one 21 parameters: three intercept variances (one for each province or phytoplankton size fraction), three slope variances (one for each province or phytoplankton size fraction) and 15 covariances. The error term is represented by  $e_{ijkl}$ . Mixed models nested in the two previous global models were fitted using the *lmer* function from the R package *lme4* (Bates et al. 2013). We fitted models containing the interaction between the covariate (dilution factor) and the two fixed factors considered (province or phytoplankton group) even when the main effects were not included in the model. Those models are equivalent to the hypothesis that grazing rate was affected by the analyzed factors whereas phytoplankton growth rate remained unaffected. We employed the second order Akaike information criterion (AICc) to perform model selection (see below), instead of AIC, because of the low ratio between sample size (n) and the number of estimated parameters (K) (Burnham and Anderson 2002).

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From AICc we computed AICc weight (AICc w) for every model, a measurement of the strength of evidence of each model. In doing that, we used the R package AICcmodavg (Mazerolle 2013). Model selection procedure was based on Zuur et al. (2009), but we performed model averaging to estimate  $\mu$  and m from a 95 % confidence set of models, which may include several fixed and random structures, if the AICc w of the best model was < 0.9 (Burnham and Andersson 2002). We firstly determined the best random structures of the q random structures considered (Supplementary material Table 1) using the most complex fixed structure (see Zuur et al. 2009). Restricted maximum likelihood (REML) was used to fit the models because we compared random structures. We interpreted the AICc weights (AICc w random str al complex fixed str) as the probability of each random structure q being the best among the whole set of random structures considered. Instead of only selecting the best random structure, we obtained the 95 % confidence set of models by adding AICc weights from the highest to the lowest until the sum ( $\sum$  AICc w) was  $\geq 0.95$  (Burnham and Anderson 2002). Then, we scaled the AICc weights of those models including the best random structures q' to sum one (scaled AICc w random str  $q \mid complex$  fixed str). Subsequently, we took each random structure q' and combined it with the different fixed structures p (Supplementary material Table 2). Because we were comparing models with different fixed structures but the same random structure, models were fitted using maximum likelihood (ML). We obtained the weight of selecting a model with fixed structure p given the random structure  $q'(AICc \ w \ fixed \ str_p | random \ str_q')$ . This can be combined with the above estimate to yield the weight of the model associated to fixed structure p accounting for the uncertainty in the selection of the random structure  $q'(AICc w_{pq'})$ .

 $AICc \ w_{pq'} = (AICc \ w_{fixed \ str \ p \ | \ random \ str \ q'}) \ (scaled \ AICc \ w \ random \ str \ q' \ | \ complex \ fixed \ str)$ 

Again, we obtained the 0.95 confidence set of models by summing AICc weights of models from the highest to the lowest until the sum was ≥ 0.95. Then, we scaled AICc weights to sum one.

Model averaging to estimate coefficients  $(\widetilde{\beta}_j)$ , i.e. the rates, was performed using the zero method proposed in Burnham and Anderson (2002):

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$$\widetilde{\widetilde{\beta}}_{j} = \sum_{i=1}^{R} model \ AICc \ w_{i} \ \widehat{\beta}_{j,i}$$

Where  $\widehat{\beta_{j,i}}$  is the estimate of  $\beta_j$  for model i. If the predictor j was not included in the model  $\widehat{\beta_{j,i}}$  was set to zero. This method entails the use of all R models included in the final set of models. The unconditional variances  $(\widehat{Var})$ , which include both within and between model variation, were estimated using the equation 6.12 proposed by Burnham and Anderson (2002):

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$$\widehat{Var}\left(\widetilde{\overline{\beta}}_{j}\right) = \sum_{i=1}^{R} model \ AICc \ w_{i} \ \left[\widehat{Var}\left(\widehat{\beta}_{j,i} \mid g_{i}\right) + \left(\widehat{\beta}_{j,i} - \widetilde{\overline{\beta}}_{j}\right)^{2}\right]$$

We calculated unconditional standard error (se) as the square root of the unconditional variance estimator (Burnham and Anderson 2002). Unconditional 95 % CI was estimated multiplying unconditional standard error by two (Burnham and Anderson 2002).

In the case of flow cytometry data, we did not analyze all the experiments together because of the positive slopes commonly detected for *Prochlorococcus* and *Synechococcus* in NATR. If all the data were analyzed together, those unrealistic microzooplankton grazing rates would affect rates of the other groups, or the rates of *Prochlorococcus* and *Synechococcus* in the other two provinces, due to the analytical procedure of mixed models. Thus, we performed three separate analyses disaggregating the data in the following form: *Prochlorococcus* and *Synechococcus* in

NATR, large eukaryotes and small eukaryotes in NATR, and the four FCM groups together in NAST-W and NAST-E. We did not estimate relative importance of variables.

#### Flow cytometry analysis

Initial cell counts of some groups were very low in some experiments. When initial cell counts < 330 in the undiluted treatment, we estimated initial cell abundances in diluted containers multiplying cell concentrations in undiluted containers by the corresponding nominal dilution.

This was the case of large eukaryotes in all the stations, small eukaryotes in NATR stations, S14 and S16, and *Prochlorococcus* in S16. Departures from the nominal dilution caused by inexact bottle fillings would be unaccounted for with this approach and could be a source of error in the estimated rates (Worden and Binder 2003). Nevertheless, we discarded this potential mistake by graphically checking that observed initial abundances of the more abundant groups (*Prochlorococcus* and *Synechococcus*) in diluted samples were similar to the abundances obtained multiplying observed abundances at the undiluted samples by the nominal dilutions (data not shown).

#### Relation between size fractionated Chl a and growth

We fitted the following mixed models to estimate the relationship between phytoplankton community size structure and size fractionated phytoplankton growth and grazing rates in each province and subprovince:

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$$\mu_{ijk} \vee m_{ijk} = a_0 + a_{Prov.} + \alpha_{size} + (b_0 + b_{Prov.} + \beta_{size}) Chl \ a^* + e_{ijk}$$
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$$(\alpha_{size}, \beta_{size}) \sim N \ (0, \sum_{size})$$
1033 
$$e_{ijk} \sim N \ (0, \sigma^2)$$

Where  $\mu_{ijk}$  and  $m_{ijk}$  are phytoplankton growth and microzooplankton grazing rates, respectively, when province = province<sub>i</sub>, size = size fraction<sub>j</sub>, station = station<sub>k</sub>, and *Chl*  $a^*$  = (*Chl*  $a_{ijk}$  -  $\overline{Chl} a_{ij}$ ). In this way, we estimated a general relationship for the three size fractions without considering differences in Chl a concentrations between size fractions, *i.e.* to isolate within group effects (e.g. van de Pol and Wright 2006).  $a_0$  and  $b_0$  are the intercept and the slope, respectively, for the reference level. Province is a fixed effect acting on both intercept ( $a_{Prov.}$ ) and slope ( $b_{Prov.}$ ).  $\alpha_{size}$  and  $\beta_{size}$  are random effects of size fraction on intercept and slope, respectively.  $\sum_{size}$  is a 2 x 2 symmetric covariance matrix containing 3 parameters: a variance for the intercept, a variance for the slope and a covariance between them.  $e_{ijk}$  is the error term.

#### Results

### Sea water properties and classification of the stations

Potential temperature at 10 m depth, the depth of the chlorophyll maximum and the variability of the N:P ratio decreased toward Iberian Peninsula. In contrast, fluorescence at 10 m depth, DIN and N:Si ratio increased toward Iberian Peninsula (Supplementary material Fig. 1). The geographic and depth patterns of DIN mimic the ones of NO<sub>3</sub><sup>-</sup>, which was much more variable than NO<sub>2</sub><sup>-</sup> and NH<sub>4</sub><sup>+</sup> (data not shown). Singularities were observed along the transect. This is the case of potential temperature in S4; salinity, potential temperature and nutrients in S10; or potential temperature, fluorescence and nutrients in S16 (Supplementary material Fig. 1). Singularities in S4 and S10 could be promoted by the presence of sea level anomalies (Supplementary material Fig. 2A).

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The NMDS ordination of the sea-water properties helped classify stations in the corresponding provinces and sub-provinces. The low NMDS stress, a measure of the goodness of fit, supported the obtained configuration (Supplementary material Fig. 3). S2 to S6 have similar values on the axis 1, mainly defined by the depth of the chlorophyll maximum and fluorescence and temperature at 10 m depth. We classified them as stations from NATR (Supplementary material Fig. 3). The S7 showed marked differences from the contiguous stations, because of its location at the boundary between NATR and NAST (Supplementary material Figs. 1 and 3). According to Longhurst (2007), the front between both provinces is defined by the position of the Subtropical convergence (STC), which in winter (near to our sampling time) matches the surface end of the 20°C isotherm. This is in agreement with the grouping of S7 with NAST-W stations (S7 surface  $T = 19.8^{\circ}$  C). The division of the group of the S7 to S16 stations, corresponding to the separation of NAST province into NAST-W and NAST-E, was supported by the observed geostrophic velocities (Supplementary material Fig. 2B). The boundary between both subprovinces was located between S11 and S12, coinciding with the topography of the Mid Atlantic Ridge (see Fig 1), which limits the entrance of water from the western Atlantic (Longhurst 2007).

## Figures and tables

Table 1. Different random structures considered in models fitted to parameterize phytoplankton growth  $(\mu)$  and microzooplankton grazing rates (m). An I letter means that intercept, i.e. phytoplankton growth, can change between stations. Consequently a standard deviation (sd) for  $\mu$  is estimated. An S letter means that slope, i.e. grazing, may change between stations and standard deviation is estimated for m. If I or S appears in columns  $Station \times Prov$ . or  $Station \times Group$ , a standard deviations for  $\mu$  or m, respectively, is estimated for each level of the fixed factor.

		Random eff	ects
Structure	Station	Station x Prov.	Station x Group
1			
2	1		
3		1	
4			1
5	S		
6		S	
7			S
8	1 & S		
9		1 & S	
10			1 & S
11	S	1	
12	1	S	
13	S		1
14	1		S

Table 2. Different fixed structures included in models fitted to parameterize phytoplankton growth ( $\mu$ ) and microzooplankton grazing rates (m). A cross (x) in the *Dilution* column means that dilution factor, i.e. grazing, was included in the model. An I letter means that a different intercept, i.e. an effect on phytoplankton growth rate, was estimated for every level of the factor. An S letter means that a different slope, i.e. an effect on grazing rate, was estimated for each level of the factor.

		Fixe	ed effec	ets
Structure	Dilution			Group x Prov.
1				
2	X			
3	X		1	
4	X		S	
5	X		1 & S	
6	X	I		
7	x	S		
8	x	1 & S		
9	x	I	1	
10	x	I	1 & S	
11	x	1 & S	- 1	
12	X	I	S	
13	X	S	- 1	
14	X	S	S	
15	X	I	1 & S	
16	X	1 & S	1	
17	X	1 & S	1 & S	
18	X	I	1	1
19	X	I	1 & S	1
20	X	1 & S	1	1
21	X	1 & S	1 & S	1
22	X	S	S	S
23	X	S	1 & S	S
24	X	1 & S	S	S
25	X	1 & S	1 & S	S
26	х	1 & S	1 & S	1 & S
27			1	
28		I		
29		1	1	
30		I	1	1

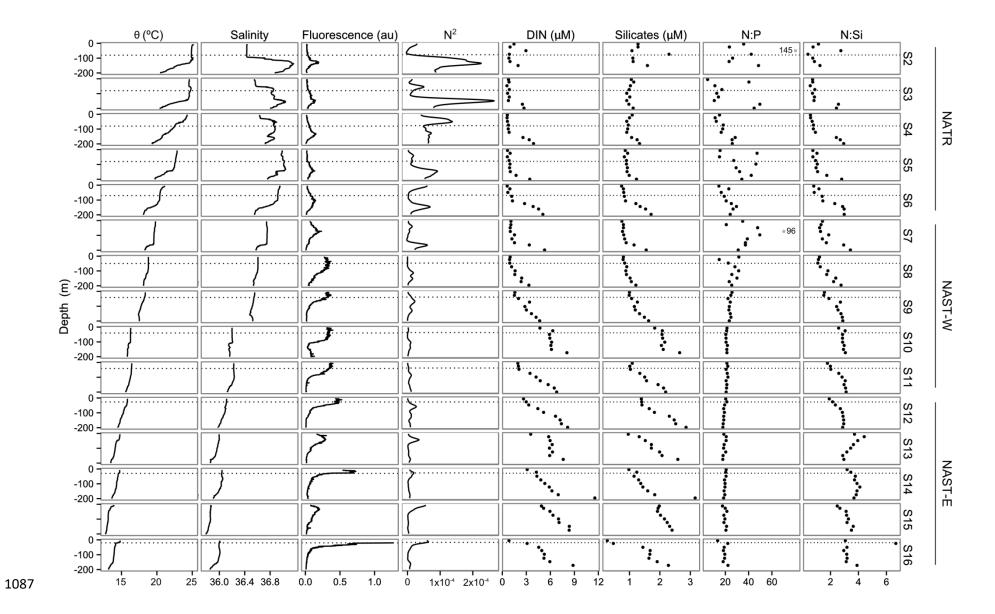


Fig. 1 Vertical profiles of potential temperature (θ), salinity, fluorescence, square Brunt-Väisälä frequency ( $N^2$ ), dissolved inorganic nitrogen (DIN), silicates, N:P ratio (N:P) and N:Si ratio (N:Si). Horizontal dotted lines indicate sampling depths of dilution experiments.

Fluorescence values of S16 up to 20m depth were excluded in order to increase the resolution of the panels at lower fluorescence values.  $N^2$  profiles were smoothed. Grey points in N:P ratio profiles at S2 and S7 show values out of the scale, their values are indicated close to the points

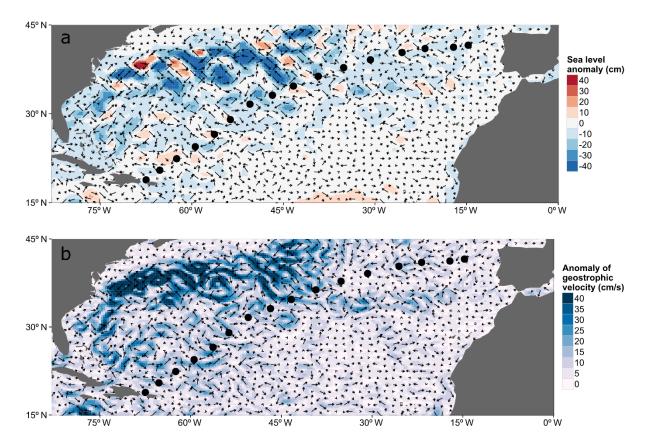


Fig 2. Satellite images showing average sea level anomalies (a) and geostrophic velocities (b) during the cruise. Black dots show the location of the stations. Scale colors indicate the magnitude of the sea level anomaly or geostrophic velocity. Arrows indicate directions of the flow.

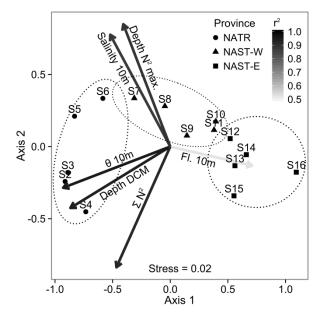


Fig. 3 Biplot showing the ordination of all stations retrieved from the NMDS, the directions of maximum correlation between the covariates used in the NMDS and the axes, and the classification of the stations. Stations are showed as points and variables included in NMDS are displayed as arrows. The symbols indicate the province or sub-province. Arrows point out the direction of maximum correlation between variables and the axes. Arrow heads indicate normalized linear regression coefficients between each variable and the axes (see methods). Arrow color shows values of R<sup>2</sup>. Stress, a measure of goodness of fit, is indicated.

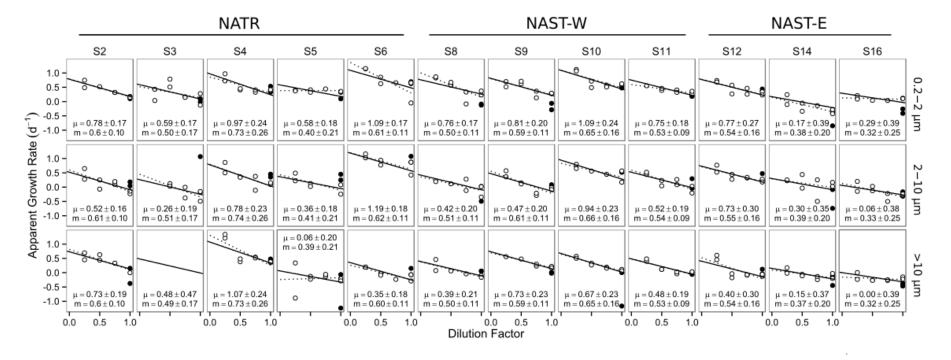


Fig. 4 Plots of dilution experiments from Chl a data for the different phytoplankton size fractions analyzed. White dots point out phytoplankton apparent growth rate (r) in treatments without nutrient addition. Black dots indicate apparent growth rate in treatments with added nutrients. Black solid lines show the fitting obtained from mixed models and model averaging. Black dotted lines show the fitting obtained from simple linear regression models for every station and size fraction.  $\mu$ : phytoplankton growth rate  $\pm$  95 % confidence interval (CI) obtained from mixed models and model averaging. m: microzooplankton grazing rate  $\pm$  95 % CI obtained from mixed models and model averaging. Because of the low Chl a concentration there are not data for the size fraction > 10  $\mu$ m in S3, although the use of mixed models allowed us estimating the parameters.

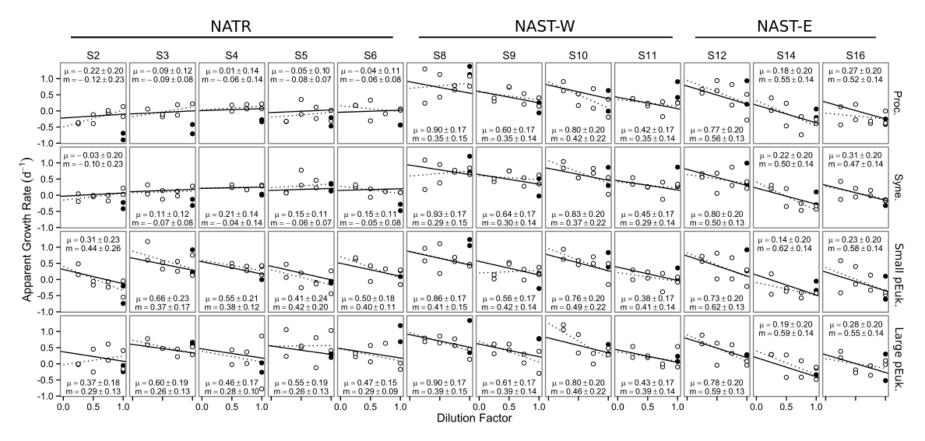


Fig. 5 Plots of dilution experiments from flow cytometry data for the different picophytoplankton groups analyzed. White dots point out phytoplankton apparent growth rate (r) in treatments without nutrient addition. Black dots indicate apparent growth rate in treatments with added nutrients. Black solid lines show the fitting obtained from mixed models and model averaging. Black dotted lines show the fitting obtained from simple linear regression models for every station and size fraction.  $\mu$ : phytoplankton growth rate  $\pm$  95 % confidence interval

- 1120 (CI) obtained from mixed models and model averaging. m: microzooplankton grazing rate  $\pm$  95 % CI obtained from mixed models and
- 1121 model averaging.

Table 3. 95 % Confidence set of models fitted with data of Chl a from dilution experiments. Models are ranked by AICc w. A cross (x) in *Dilution* column means that dilution factor was included in the model. Fixed and random effect columns show the different fixed and random factors included in models. The letter I means that an intercept (phytoplankton growth rate) was estimated for every level of the factor. The letter S means that a slope (microzooplankton grazing rate) was estimated in each level of the factor. K: number of parameters. AICc w scaled random str: scaled AICc w to obtain  $\sum$  AICc w = 1 considering models with different random structures and the most complex fixed structure included in the 95 % confidence set of models. AICc w Fixed str: AICc w of models with different fixed structures conditioned on some of the better random structures. AICc w Model: AICc w obtained multiplying scaled AICc w of random structures by AICc w of fixed structures. AICc w scaled Model: scaled Model AICc w to obtain  $\sum$  AICc w = 1.  $\sum$  AICc w Model: Cumulative Model AICc w.  $\sum$  AICc w scaled. Model: cumulative model AICc w using scaled model AICc w.

Rank		Fixed	effects		Rand	dom effects	V		AICc w				∑ AICc w
Kalik	Dilution	Size	Prov.	Size:Prov.	Station	Station x Size	K	scaled random str.	Fixed str.	Model	scaled Model	Model	scaled Model
1	х	1			S	1	15	0.6885	0.3933	0.2708	0.2838	0.2708	0.2838
2	x	1	1		S	1	17	0.6885	0.1913	0.1317	0.1381	0.4025	0.4218
3	x	1	1 & S			1	15	0.3115	0.3801	0.1184	0.1241	0.5209	0.5459
4	x	1	1 & S		S	1	19	0.6885	0.1219	0.0839	0.0879	0.6048	0.6339
5	x	1	1			1	13	0.3115	0.2459	0.0766	0.0803	0.6814	0.7141
6	x	1 & S			S	1	17	0.6885	0.0708	0.0488	0.0511	0.7301	0.7652
7	x	1	S		S	1	17	0.6885	0.0543	0.0374	0.0392	0.7675	0.8044
8	x	1	1	1	S	1	21	0.6885	0.0463	0.0319	0.0334	0.7994	0.8379
9	x	1				1	11	0.3115	0.0821	0.0256	0.0268	0.8250	0.8647
10	x	1 & S	1		S	1	19	0.6885	0.0328	0.0226	0.0237	0.8476	0.8883
11	x	1 & S	1 & S			1	17	0.3115	0.0655	0.0204	0.0214	0.8680	0.9097

12	Х	1	1 & S	1		1	19	0.3115	0.0497	0.0155	0.0162	0.8835	0.9259
13	Х	1 & S	1 & S		S	1	21	0.6885	0.0201	0.0139	0.0145	0.8973	0.9405
14	Х	1 & S	I			1	15	0.3115	0.0444	0.0138	0.0145	0.9112	0.9550
15	Х	1	I	1		1	17	0.3115	0.0338	0.0105	0.0110	0.9217	0.9660
16	Х	1	1 & S	1	S	1	23	0.6885	0.0131	0.0090	0.0095	0.9307	0.9755
17	Х	1	S			1	13	0.3115	0.0273	0.0085	0.0089	0.9392	0.9844
18	Х	S			S	1	15	0.6885	0.0123	0.0085	0.0089	0.9477	0.9932
19	Х	1	S		S	1	19	0.6885	0.0094	0.0065	0.0068	0.9541	1.0000

Table 4. 95 % confidence set of models fitted with flow cytometry data from dilution experiments carried out in NAST. Models are ranked by AICc w. A cross (x) in *Dilution* column means that dilution factor was included in the model. Fixed and random effect columns show the different fixed and random factors included in models. The letter *I* means that an intercept (phytoplankton growth rate) was estimated for every level of the factor. The letter *S* means that the interaction with dilution factor (microzooplankton grazing rate) was estimated in each level of the factor. K: number of parameters estimated by the model. AICc w scaled random str: scaled AICc w to obtain  $\sum$  AICc w = 1 considering models with different random structures and the most complex fixed structure included in the 95 % confidence set of models. AICc w Fixed str: AICc w of models fitted with different fixed structures and some of the better random structures. AICc w Model: AICc w obtain  $\sum$  AICc w fixed AICc w of fixed structures. AICc w scaled Model: scaled Model AICc w to obtain  $\sum$  AICc w Model: Cumulative Model AICc w.  $\sum$  AICc w scaled Model: cumulative model AICc w using scaled model AICc w.

Dank		Fixed	d effec	ts	Ran	dom effects	V		AICc w			2	∑ AICc w
Rank	Dilution	Group	Prov.	Group x Prov.	Station	Station x Prov.	K	scaled random str.	Fixed str.	Model	scaled Model	Model	scaled Model
1	Х	S	S		1		8	0.6778	0.1703	0.1154	0.1214	0.1154	0.1214
2	X	1	S		1		8	0.6778	0.1516	0.1028	0.1080	0.2182	0.2294
3	X	S	1 & S		1		9	0.6778	0.1313	0.0890	0.0936	0.3072	0.3230
4	X	I	1 & S		1		9	0.6778	0.1169	0.0792	0.0833	0.3864	0.4063
5	X	S	I		1		8	0.6778	0.0958	0.0649	0.0683	0.4513	0.4746
6	X	I	I		1		8	0.6778	0.0854	0.0579	0.0609	0.5093	0.5355
7	X	S	S		1 & S		10	0.2457	0.1494	0.0367	0.0386	0.5460	0.5740
8	X	S	1		1 & S		10	0.2457	0.1380	0.0339	0.0356	0.5798	0.6097
9	X	1	S		1 & S		10	0.2457	0.1326	0.0326	0.0342	0.6124	0.6439

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10	Х	I	1		1 & S		10	0.2457	0.1225	0.0301	0.0316	0.6425	0.6756
11	Х	S			I		7	0.6778	0.0414	0.0281	0.0295	0.6706	0.7051
12	X	1			1		7	0.6778	0.0369	0.0250	0.0263	0.6956	0.7314
13	X	S	1 & S		1 & S		11	0.2457	0.0936	0.0230	0.0242	0.7186	0.7556
14	х	1	1 & S		1 & S		11	0.2457	0.0831	0.0204	0.0215	0.7390	0.7770
15	Х	1 & S	S		1		11	0.6778	0.0269	0.0182	0.0192	0.7573	0.7962
16	х	S			1 & S		9	0.2457	0.0635	0.0156	0.0164	0.7729	0.8126
17	х	1			1 & S		9	0.2457	0.0564	0.0139	0.0146	0.7867	0.8272
18	х	1 & S	1 & S		1		12	0.6778	0.0201	0.0137	0.0144	0.8004	0.8416
19	х	S	S			1	10	0.0765	0.1747	0.0134	0.0141	0.8137	0.8556
20	х		S		I		5	0.6778	0.0197	0.0133	0.0140	0.8271	0.8696
21	х	1	S			I	10	0.0765	0.1556	0.0119	0.0125	0.8390	0.8822
22	х	S	S S		I		11	0.6778	0.0172	0.0117	0.0123	0.8507	0.8944
23	х		1 & S		I		6	0.6778	0.0155	0.0105	0.0111	0.8612	0.9055
24	х	1 & S	1		I		11	0.6778	0.0149	0.0101	0.0106	0.8713	0.9161
25	х	S	1 & S			I	11	0.0765	0.1207	0.0092	0.0097	0.8805	0.9258
26	х	S	1 & S	S	I		12	0.6778	0.0129	0.0087	0.0092	0.8893	0.9350
27	х	1	1 & S			I	11	0.0765	0.1075	0.0082	0.0086	0.8975	0.9437
28	х		1		I		5	0.6778	0.0118	0.0080	0.0084	0.9054	0.9521
29	х	S	1			I	10	0.0765	0.0899	0.0069	0.0072	0.9123	0.9593
30	х	1	1			I	10	0.0765	0.0802	0.0061	0.0065	0.9185	0.9657
31	х	1 & S	S		1 & S		13	0.2457	0.0231	0.0057	0.0060	0.9241	0.9717
32	х	1 & S	1		1 & S		13	0.2457	0.0213	0.0052	0.0055	0.9294	0.9772
33	х	1	1 & S	1	I		12	0.6778	0.0076	0.0051	0.0054	0.9345	0.9826
34	х	1 & S			I		10	0.6778	0.0066	0.0045	0.0047	0.9390	0.9873
35	x	S				I	9	0.0765	0.0562	0.0043	0.0045	0.9433	0.9918
36	x		S		1 & S		7	0.2457	0.0159	0.0039	0.0041	0.9472	0.9960
37	x	1	1	1	I		11	0.6778	0.0057	0.0038	0.0040	0.9511	1.0000

Table 5. 95 % confidence set of models fitted with flow cytometry data of cyanobacteria from dilution experiments carried out in NATR. Models are ranked by AICc w. A cross (x) in *Dilution* column means that dilution factor was included in the model. Fixed and random effect columns indicate the different fixed and random factors included in models. The letter *I* means that an intercept (phytoplankton growth rate) was estimated for every level of the factor. The letter *S* means that the interaction with dilution factor (microzooplankton grazing rate) was estimated in each level of the factor. K: number of parameters estimated by the model. AICc w scaled random str: scaled AICc w to obtain  $\sum$  AICc w = 1 using models with different random structures and the most complex fixed structure included in the 95% confidence set of models. AICc w Fixed str: AICc w of models with different fixed structures and some of the better random structures. AICc w Model: AICc w obtained multiplying scaled AICc w of random structures by AICc w of fixed structures. AICc w scaled Model: scaled Model AICc w to obtain  $\sum$  AICc w = 1.  $\sum$  AICc w Model: Cumulative Model AICc w.  $\sum$  AICc w scaled Model: cumulative model AICc w using scaled model AICc w.

Dank	Fixed e	ffects	Rar	ndom effects	V		AICc w				∑ AICc w
Rank	Dilution	Group	Station	Station x Group	K	scaled random str.	Fixed str.	Model	scaled Model	Model	scaled Model
1		I	1		4	0.6159	0.4046	0.2491	0.2590	0.2491	0.2590
2	X	I	I		5	0.6159	0.3887	0.2394	0.2488	0.4885	0.5077
3		I	1 & S		6	0.1976	0.5764	0.1139	0.1184	0.6024	0.6261
4	X	1 & S	I		6	0.6159	0.1792	0.1104	0.1147	0.7128	0.7409
5	X	I	1 & S		7	0.1976	0.2823	0.0558	0.0580	0.7686	0.7989
6		I		1	6	0.1093	0.3586	0.0392	0.0407	0.8078	0.8396
7	X	I		1	7	0.1093	0.3311	0.0362	0.0376	0.8440	0.8772
8		I	S		4	0.0772	0.4025	0.0311	0.0323	0.8751	0.9095
9	X	I	S		5	0.0772	0.3892	0.0300	0.0312	0.9051	0.9407

10	Х	1 & S	1 & S		8	0.1976	0.1239	0.0245	0.0255	0.9296	0.9662
11	Х	S	1		5	0.6159	0.0268	0.0165	0.0171	0.9461	0.9833
12	х	1 & S		I	8	0.1093	0.1466	0.0160	0.0167	0.9621	1.0000

Table 6. 95 % confidence set of models fitted with Flow cytometry data of eukaryotes from dilution experiments carried out in NATR. Models are ranked by AICc w. A cross (x) in *Dilution* column means that dilution factor was included in the model. Fixed and random effect columns show the different fixed and random factors included in models. The letter *I* means that an intercept (phytoplankton growth rate) was estimated for every level of the factor. The letter *S* means that the interaction with dilution factor (microzooplankton grazing rate) was estimated in each level of the factor. K: number of parameters estimated by the model. AICc w scaled random str: scaled AICc w to obtain  $\sum$  AICc w = 1 using models with different random structures and the most complex fixed structure included in the 95 % confidence set of models. AICc w Fixed str: AICc w of models fitted with different fixed structures and some of the better random structures. AICc w Model: AICc w obtained multiplying scaled AICc w of random structures by AICc w of fixed structures. AICc w scaled Model: scaled Model AICc w to obtain  $\sum$  AICc w = 1.  $\sum$  AICc w Model: Cumulative Model AICc w.  $\sum$  AICc w scaled Model: cumulative model AICc w using scaled model AICc w.

Dank	Fixed e	ffects	Rar	ndom effects	V		AICc w			Σ	AICc w
Rank	Dilution	Group	Station	Station x Group	K	scaled random str.	Fixed str.	Model	scaled Model	Model	scaled Model
1	Х			1	6	0.4384	0.3627	0.1590	0.1660	0.1590	0.1660
2	X	S		1	7	0.4384	0.3214	0.1409	0.1471	0.2999	0.3131
3	X	S	1		5	0.3573	0.3505	0.1252	0.1307	0.4252	0.4438
4	x		1		4	0.3573	0.2572	0.0919	0.0959	0.5170	0.5397
5	X	1	1		5	0.3573	0.1939	0.0693	0.0723	0.5863	0.6120
6	x	1		1	7	0.4384	0.1368	0.0600	0.0626	0.6463	0.6746
7	X	1 & S		1	8	0.4384	0.1288	0.0565	0.0590	0.7027	0.7335
8	x	1 & S	1		6	0.3573	0.1300	0.0464	0.0485	0.7492	0.7820
9	X			S	6	0.0850	0.4088	0.0347	0.0363	0.7839	0.8183

10	x	S	S		5	0.0851	0.3139	0.0267	0.0279	0.8107	0.8461
11	x		S		4	0.0851	0.2592	0.0221	0.0230	0.8327	0.8692
12	x	S		S	7	0.0850	0.1908	0.0162	0.0169	0.8489	0.8861
13				1	5	0.4384	0.0362	0.0159	0.0166	0.8648	0.9027
14	x	I	S		5	0.0851	0.1793	0.0153	0.0159	0.8801	0.9186
15			I		3	0.3573	0.0401	0.0143	0.0150	0.8944	0.9335
16				S	5	0.0850	0.1535	0.0130	0.0136	0.9074	0.9472
17	x	S	1 & S		7	0.0342	0.3292	0.0113	0.0118	0.9187	0.9589
18	x	I		S	7	0.0850	0.1240	0.0105	0.0110	0.9292	0.9699
19		I	1		4	0.3573	0.0283	0.0101	0.0106	0.9393	0.9805
20	x	1 & S	S		6	0.0851	0.1169	0.0100	0.0104	0.9493	0.9909
21	х		1 & S		6	0.0342	0.2561	0.0088	0.0091	0.9581	1.0000

μm

lm

0.72 ± 0.60

1169

1170

1171

1172

Table 7. Phytoplankton growth and microzooplankton grazing rates for every station and size fraction estimated by fitting mixed models and conducting model averaging (lmm + ma) or by fitting separate linear regression models for each experiment (lm). The 95% confidence intervals ( $\pm$ ) are also indicated.

					NATR				NAS	T-W			NAST-E	
			S2	<b>S3</b>	S4	S5	S6	S8	S9	S10	S11	S12	S14	S16
-1)	0.2-2	lmm + ma	0.78 ± 0.17	0.59 ± 0.17	0.97 ± 0.24	0.58 ± 0.18	1.09 ± 0.17	0.76 ± 0.17	0.81 ± 0.20	1.09 ± 0.24	0.75 ± 0.18	0.77 ± 0.27	0.17 ± 0.39	0.29 ± 0.39
P	μm	lm	0.79 ± 0.25	0.52 ± 0.61	0.86 ± 0.33	0.38 ± 0.12	1.37 ± 0.55	1.01 ± 0.45	0.80 ± 0.32	1.09 ± 0.38	0.59 ± 0.14	0.78 ± 0.28	0.15 ± 0.37	0.13± 0.17
rate	2-10	lmm + ma	0.52 ± 0.16	0.26 ± 0.19	0.78 ± 0.23	0.36 ± 0.18	1.19 ± 0.18	0.42 ± 0.20	0.47 ± 0.20	0.94 ± 0.23	0.52 ± 0.19	0.73 ± 0.30	0.30 ± 0.35	0.06 ± 0.38
	μm	lm	0.60 ± 0.38	0.45 ± 0.80	0.80 ± 0.42	$0.43 \pm 0.41$	1.20 ± 0.44	0.36 ± 0.65	0.56 ± 0.40	0.83 ± 0.38	0.60 ± 0.29	0.71 ± 0.24	0.29 ± 0.62	0.13 ± 0.36
Growth	> 10	lmm + ma	0.73 ± 0.19	0.48 ± 0.47	1.07 ± 0.24	0.06 ± 0.20	0.35 ± 0.18	0.39 ± 0.21	0.73 ± 0.23	0.67 ± 0.23	0.48 ± 0.19	0.40 ± 0.30	0.15 ± 0.37	0.00 ± 0.39
9	μm	lm	0.81 ± 0.36	-	1.31 ± 0.54	-0.24 ± 0.77	0.26 ± 0.25	0.38 ± 0.36	0.69 ± 0.16	0.64 ± 0.12	0.49 ± 0.22	0.52 ± 0.45	0.09 ± 0.23	-0.15 ± 0.21
-1)	0.2-2	lmm + ma	0.60 ± 0.10	0.50 ± 0.17	0.73 ± 0.26	0.40 ± 0.21	0.61 ± 0.11	0.50 ± 0.11	0.59 ± 0.11	0.65 ± 0.16	0.53 ± 0.09	0.54 ± 0.16	0.38 ± 0.20	0.32 ± 0.25
P)	μm	lm	0.63 ± 0.39	$0.38 \pm 0.89$	0.55 ± 0.48	0.00 ± 0.20	1.07 ± 0.82	0.85 ± 0.66	0.56 ± 0.46	0.61 ± 0.56	$0.30 \pm 0.19$	0.53 ± 0.41	0.50 ± 0.55	0.05 ± 0.24
rate	2-10	lmm + ma	0.61 ± 0.10	0.51 ± 0.17	0.74 ± 0.26	0.41 ± 0.21	0.62 ± 0.11	0.51 ± 0.11	0.61 ± 0.11	0.66 ± 0.16	0.54 ± 0.09	0.55 ± 0.16	0.39 ± 0.20	0.33 ± 0.25
ngı	μm	lm	0.76 ± 0.56	0.80 ± 1.03	0.76 ± 0.62	0.54 ± 0.62	0.57 ± 0.66	0.47 ± 0.86	0.78 ± 0.58	0.47 ± 0.55	0.67 ± 0.43	0.51 ± 0.35	0.30 ± 0.90	0.47 ± 0.52
razing	> 10	lmm + ma	0.60 ± 0.10	0.49 ± 0.17	0.73 ± 0.26	0.39 ± 0.21	0.60 ± 0.11	0.50 ± 0.11	0.59 ± 0.11	0.65 ± 0.16	0.53 ± 0.09	0.54 ± 0.16	0.37 ± 0.20	0.32 ± 0.25

 $1.05 \pm 0.79$   $-0.07 \pm 1.13$   $0.50 \pm 0.37$ 

0.52 ± 0.52

 $0.52 \pm 0.21$   $0.59 \pm 0.17$ 

 $0.56 \pm 0.33$ 

 $0.74 \pm 0.66$   $0.27 \pm 0.33$   $0.08 \pm 0.30$ 

Table 8. Mean phytoplankton growth and microzooplankton grazing rates for every province and size fraction estimated by fitting mixed models and conducting model averaging (lmm + ma) or by averaging the rates obtained by fitting separate linear regression models for each experiment (lm). In the latter case, we assigned a value equal to 0 to the negative rates estimated at some stations. Standard deviations ( $\pm$ ) are also indicated.

			NATR	NAST-W	NAST-E
1)	0.2.2	lmm + ma	0.76 ± 0.25	0.79 ± 0.26	0.56 ± 0.29
(d <sup>-1</sup> )	0.2-2 μm	lm	0.78 ± 0.38	0.87 ± 0.22	0.35 ± 0.37
ate	2.10	lmm + ma	0.58 ± 0.31	0.71 ± 0.32	0.40 ± 0.34
th I	2-10 μm	lm	0.70 ± 0.32	0.59 ± 0.19	0.38 ± 0.30
Growth rate	> 10 · · · · ·	lmm + ma	0.50 ± 0.30	0.53 ± 0.30	0.31 ± 0.33
g	> 10 µm	lm	0.60 ± 0.58	0.55 ± 0.14	0.20 ± 0.28
1)	0.2.2	lmm + ma	0.54 ± 0.14	0.55 ± 0.14	0.48 ± 0.16
(d <sup>-1</sup> )	0.2-2 μm	lm	0.53 ± 0.39	0.58 ± 0.23	0.36 ± 0.27
rate	2.10	lmm + ma	0.55 ± 0.14	0.56 ± 0.14	0.49 ± 0.16
ng i	2-10 μm	lm	0.69 ± 0.12	0.60 ± 0.15	0.43 ± 0.11
Grazing ı	> 10 um	lmm + ma	0.54 ± 0.14	0.55 ± 0.14	0.48 ± 0.16
9	> 10 µm	lm	0.57 ± 0.44	0.55 ± 0.03	0.36 ± 0.34

Table 9. Phytoplankton abundances (cells mL<sup>-1</sup>) observed at the different stations.

Station	S2	S3	S4	S5	S6	S8	S9	S10	S11	S12	S14	S16
Cyanobacteria												
Prochlorococcus spp.	17414	21229	31088	19953	29530	10091	10111	2293	79636	6865	17718	431
Rhizomonas setigera¹	-	-	-	-	0.04	-	-	-	-	-	-	0.04
Synechococcus spp.	4327	5458	9429	5748	12170	35833	16432	17028	39360	23520	20448	2888
Diatoms												
Chaetoceros atlanticus	-	-	-	-	-	-	-	-	-	-	-	0.92
Chaetoceros lorenz	-	-	-	-	-	-	-	-	-	-	-	0.32
Chaetoceros peruvianum	-	-	-	-	-	-	0.04	-	-	-	0.88	0.6
Corethron criophillum	-	-	-	-	-	-	-	0.04	-	0.36	0.04	2.2
Coscinodiscus spp.	-	0.04	-	-	-	-	-	-	-	-	-	-
Dactyliosolen fragilissimus	-	-	-	-	-	-	0.04	-	-	-	-	-
Guinardia striata	-	-	-	-	-	-	-	-	-	-	-	1.28
Hemiaulus spp.	-	0.04	-	0.12	-	-	-	-	-	-	-	-
Navicula spp.	0.12	-	-	-	-	-	-	0.08	0.56	0.04	0.6	3.28
Nitzschia spp.	-	0.32	0.12	0.12	0.04	0.08	1.12	-	0.96	0.48	3.92	-
Nitzschia delicatissima	-	-	-	-	-	-	-	0.6	-	-	-	268
Nitzschia longissima	-	-	-	-	-	-	-	-	0.08	-	-	-
Pleurosigma spp.	-	-	-	-	-	-	-	-	-	-	0.12	0.28
Proboscia alata	-	-	-	-	-	-	-	-	0.04	-	-	-
Rhizosolenia hebetata	-	0.04	-	-	0.04	0.04	0.04	-	0.08	-	-	-
Rhizosolenia imbricata	-	-	-	-	-	-	-	-	-	-	-	0.12
Thalassionema nitzschioides	-	-	-	-	-	-	-	-	-	-	-	0.6
Dinoflagellates												
Amphidinium spp.	0.16	-	-	-	-	0.04	-	-	-	-	-	0.04
Amphidoma caudata	-	-	-	-	-	-	-	-	-	-	-	0.04
Ceratium spp.	-	-	0.04	-	-	-	-	-	-	-	-	-
Dinophysis schuettii	-	0.04	-	-	-	-	-	-	-	-	-	-
Gymnodinium spp.	1.2	1.08	1.52	2.24	0.88	1.28	2.88	1.68	2.68	1.68	1.12	2.28
Gyrodinium spp.	-	-	0.16	0.04	0.24	0.32	0.2	0.16	0.32	0.16	0.16	1.88
Gyrodinium spirale	0.12	0.04	-	-	-	-	-	-	-	0.04	0.04	-
Heterocapsa niei	-	-	-	-	-	-	-	-	-	-	-	0.08
Katodinium glaucum	0.04	0.04	0.08	0.16	0.16	-	-	0.04	0.2	0.08	0.08	0.04
Oxytoxum scolopax	-	-	-	-	-	0.04	0.04	0.04	0.04	-	-	-
Podolampas palmipes	-	0.04	-	-	-	-	-	-	-	-	-	-
prorocentrum spp.	0.12	-	0.08	-	-	-	-	-	-	0.04	0.04	-
Prorocentrum compresum	-	-	-	-	-	-	-	-	-	-	-	0.44
Protopteridinium steinii	-	-	-	-	-	-	-	-	-	0.04	-	-
Scrippsiella trochoidea	0.16	0.12	0.12	0.08	0.12	0.12	-	0.04	0.04	-	-	-
Torodinium robustum	0.04	-	-	0.08	0.04	-	0.08	-	-	-	-	-
Torodinium spp.	-	-	0.04	-	-	0.04	-	-	-	0.08	-	-

Silicoflagellates												
Dictyocha fibula	-	0.04	-	-	-	-	0.04	0.04	-	-	0.04	1.68
Non taxonomic groups												
Large Eukaryotes	115	91	155	75	175	524	577	433	815	918	997	1188
Small Eukaryotes	320	152	128	312	325	2872	7257	15205	10424	19183	1445	621

<sup>&</sup>lt;sup>1</sup> Rhizomonas setigera abundances are expressed in colonies mL<sup>-1</sup>.

Table 10. Slopes for each province and size fraction for the relationships between size fractionated centered Chl *a* and both size fractionated phytoplankton growth rate and size fractionated microzooplankton grazing rate.

Province	Size fraction (µm)	Slope growth	Slope grazing
NATR	0.2-2	18.41	13.01
NATR	2-10	13.62	11.56
NATR	> 10	14.17	11.81
NAST-W	0.2-2	-1.32	0.28
NAST-W	2-10	-6.12	-1.17
NAST-W	> 10	-5.57	-0.92
NAST-E	0.2-2	3.55	0.94
NAST-E	2-10	-1.24	-0.51
NAST-E	> 10	-0.69	-0.26

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