# Dynamic phenomena arising from an extended Core Group model

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## Abstract

In order to obtain a reasonably accurate model for the spread of a particular infectious disease through a population, it may be necessary for this model to possess some degree of structural complexity. Many such models have, in recent years, been found to exhibit a phenomenon known as backward bifurcation, which generally implies the existence of two subcritical endemic equilibria. It is often possible to refine these models yet further, and we investigate here the influence such a refinement may have on the dynamic behaviour of a system in the region of the parameter space near  $R_0 = 1$ .

We consider a natural extension to a so-called core group model for the spread of a sexually transmitted disease, arguing that this may in fact give rise to a more realistic model. From the deterministic viewpoint we study the possible shapes of the resulting bifurcation diagrams and the associated stability patterns. Stochastic versions of both the original and the extended models are also developed so that the probability of extinction and time to extinction may be examined, allowing us to gain further insights into the complex system dynamics near  $R_0 = 1$ . A number of interesting phenomena are observed, for which heuristic explanations are provided.

## **1** Introduction

When using a system of deterministic differential equations to model the spread of an infectious disease within a population it is possible, by keeping all but one of the system parameter values fixed, to obtain a diagram showing how the endemic and the disease-free equilibrium solutions change as the remaining parameter,  $\alpha$  say, is varied. For some particular value of  $\alpha$  this diagram indicates a change in the qualitative behaviour of the system, at which point the disease-free equilibrium (DFE) bifurcates into a branch representing an endemic equilibrium and a further branch of the DFE. The parameter  $\alpha$  and its associated graph are, as a consequence, known as a *bifurcation parameter* and *bifurcation diagram* respectively. The point at which this change in behaviour occurs is termed the *bifurcation point*, and the curve emanating from it the *bifurcation curve*.

By considering the *basic reproduction ratio*  $R_0$  for the system it is straightforward to obtain the location of the bifurcation point. The basic reproduction ratio is defined to be the expected number of secondary cases produced in a population at the DFE by a typical infective individual during his or her entire infectious period. In general  $R_0$  will be a function of the system parameters (see Diekmann *et al.* (1990) for a precise mathematical formulation), and the value of  $\alpha$  at the bifurcation point corresponds to  $R_0 = 1$ .

The bifurcation diagrams of simple epidemic models always display *forward bifurcation*. In this case the bifurcation curve is such that as one moves along it from the bifurcation point, the level of infection increases as  $R_0$  increases, and the disease is able to persist in the population when  $R_0 > 1$  but dies out otherwise. In recent years, however, a phenomenon known as *backward bifurcation* has emerged whereby the disease can, for certain parameter values, persist even when  $R_0 \le 1$ . In this case the initial direction of the bifurcation curve is such that as one moves along it from the bifurcation point,  $R_0$  decreases as the level of infection increases. It seems that the potential for the existence of backward bifurcation in an epidemic model was first noted in similar papers by Castillo-Chavez *et al.* (1989*a*) and (1989*b*), and Huang *et al.* (1992). Some of the more recent papers in this area include those by Castillo-Chavez and Song (2004), Brauer (2004), Feng *et al.* (2000) and Song *et al.* (2006). An extensive literature survey of epidemic models exhibiting this phenomenon was carried out by Griffiths (2007).

The presence of backward bifurcation indicates the existence of two or more endemic equilibria for  $R_0 < 1$ , known as *subcritical endemic equilibria*. It has been demonstrated in some models exhibiting backward bifurcation (see Greenhalgh *et al.* (2000), for example) that there is the possibility for subcritical endemic equilibria to be locally asymptotically stable (LAS). This certainly has implications for disease control since the classical requirement for the eradication of the disease is no longer satisfied in such cases. It is now possible for the proportion of infected individuals in the population to remain at a steady level or even invade when  $R_0 \leq 1$ .

The phenomenon of backward bifurcation tends to arise in relatively complex deterministic epidemic models, often ones possessing some sort of group structure. Such models may, for example, incorporate disease-driven changes in behaviour or take account of the fact that the disease has several stages. Despite their complexity, many of these models may still be regarded as over-simplified representations of the progress of the disease through a population. In such cases it might be possible to extend the model in a natural way, thereby giving a better approximation to the true disease structure. Indeed, Greenhalgh and Griffiths (2009) argued that a three-stage model for the spread of bovine respiratory syncytial virus (BRSV) in cattle may be more realistic than the two-stage model studied by Greenhalgh *et al.* 

(2000). It might then be asked whether this increase in the complexity of the model provides scope for yet more complicated bifurcation diagrams and hence more complicated system dynamics in the region of the parameter space near  $R_0 = 1$ . Subsequent investigation of the three-stage BRSV model revealed that this was in fact the case.

In this paper we study, in connection with the points made above, the *Core Group* model (CG model) for the spread of a sexually transmitted disease as described by Hadeler and Castillo-Chavez (1995). The CG model is able to exhibit backward bifurcation, although we note here that there is a distinct structural difference between the CG and BRSV models. While the latter models a disease that passes through several stages, the CG model incorporates disease-driven changes in behaviour. After outlining the main features of this model, we carry out an analytic study of the local asymptotic stability of the endemic equilibria with the purpose of seeing whether a general result emerges relating the stability of endemic equilibria to their positions on the bifurcation curve. The CG model is then extended in a natural way in order to explore the possibility that more complicated bifurcation diagrams and stability patterns might appear, as was found when the BRSV model was extended from two to three stages.

Stochastic aspects of the CG models are also studied here. In particular, we explore the interaction between the deterministic phenomenon of backward bifurcation and the probability of extinction for the corresponding stochastic version of each model. Our main purpose here is to compare the theoretical probabilities of extinction for stochastic formulations of the model with the corresponding probabilities obtained via an extensive series of stochastic simulations. We would hope to be able to offer explanations for any observed discrepancies. Our investigations were carried out using analytical and numerical methods, and also by way of computer simulations. We have indeed found interesting links between the presence of backward bifurcation in the deterministic models and the probability of extinction in the stochastic versions. Furthermore, the expected time to extinction for the CG model is considered in order to see whether, in certain circumstances, it is possible to observe significant discrepancies between theoretical and simulated values, in contrast to the rather inconclusive results of Griffiths (2007) for the two-stage BRSV model. Some more unusual bifurcation diagrams are then obtained by using the full epidemic model (i.e. the model for the population as a whole rather than just that for the isolated sexually active core group).

When investigating backward bifurcation and associated phenomena in the two and three-stage BRSV models, Greenhalgh and Griffiths (2009) obtained, both from the deterministic and the stochastic point of view, a number of potentially interesting results. It cannot be assumed however, that such results, obtained by studying one model in isolation, will automatically carry over to other epidemic models. Although detailed analysis revealed a number of phenomena that were noteworthy in their own right, it may be that these were in fact

particular to the model that was being considered. Thus, in the light of the findings for the BRSV models, we would like to see which of these phenomena do actually transfer to other epidemic models.

# 2 The basic Core Group model

and

Hadeler and Castillo-Chavez (1995) consider the spread of a sexually transmitted disease. The population *P* is split into two classes; a sexually active and relatively small core group *C* and a weakly connected and sexually inactive remainder non-core group *A*. The core group is further subdivided into susceptible *S*, educated (or vaccinated) *V* and infected *I* individuals with C = S + V + I and P = A + C. Members of the core group are recruited from the non-core group.

This scenario is modelled by way of a general set of differential equations. In order to be able to draw some conclusions about the behaviour of this rather complex model, the following system of differential equations, modelling an isolated core population of constant size C, is studied in detail:

$$\frac{dS}{dt} = \mu C - \beta \frac{SI}{C} - \psi S + \alpha (1 - \gamma)I - \mu S , \qquad (2.1)$$

$$\frac{dV}{dt} = \psi S - \tilde{\beta} \frac{VI}{C} + \alpha \gamma I - \mu V$$
(2.2)

$$\frac{dI}{dt} = \frac{\beta SI + \tilde{\beta} VI}{C} - \alpha I - \mu I , \qquad (2.3)$$

where  $\mu$  is the (per capita) common birth and death rate,  $\beta$  is the transmission rate from infected to susceptible individuals,  $\tilde{\beta}$  is the transmission rate from infected to educated (vaccinated) individuals (with  $0 \le \tilde{\beta} \le \beta$ ),  $\alpha$  is the recovery rate,  $\gamma$  is the proportion of recovered individuals passing into the educated class and  $\psi$  is the rate of direct transition from the susceptible class to the educated class. Since the above system is homogeneous it can be normalised by setting C = 1, meaning that *S*, *V* and *I* then represent population proportions rather than numbers of individuals.

The reproduction ratios for initial populations consisting entirely of susceptible and educated individuals respectively are

$$R = \frac{\beta}{\alpha + \mu}$$
 and  $\tilde{R} = \frac{\tilde{\beta}}{\alpha + \mu}$ 

and the basic reproduction ratio is given by

$$R_{0}(\psi) = \frac{\mu}{\mu + \psi} R + \frac{\psi}{\mu + \psi} \widetilde{R} = \frac{\mu\beta + \psi\beta}{(\alpha + \mu)(\mu + \psi)},$$

where  $R_0$  is written as a function of  $\psi$  in order to indicate that  $\psi$  is to be utilised as the bifurcation parameter. We note here that R and  $\tilde{R}$  given above are denoted  $R_0$  and  $\tilde{R}_0$  respectively in the paper. We make this change in order to avoid confusion over the commonly accepted notation for the basic reproduction ratio that we have adopted here. The authors make the point that education is not necessary when R < 1, while if  $\tilde{R} > 1$  then education is not effective, so the interesting situation is  $\tilde{R} < 1 < R$ , and we shall assume that this is the case. When  $\tilde{R} < 1 < R$  the unique education rate unique education rate  $\psi$  for which  $R_0(\psi) = 1$  is given by

$$\psi^* = \frac{R-1}{1-\widetilde{R}}\mu > 0 \,.$$

#### **3** Locally asymptotically stable endemic equilibria and the bifurcation curve

In their detailed analysis of the endemic equilibria of the two-stage BRSV model, Greenhalgh *et al.* (2000) found that when backward bifurcation was present the upper subcritical endemic equilibrium was always LAS while the lower one was always unstable. This could be interpreted as saying that a particular endemic equilibrium was LAS if, and only if, the gradient of the bifurcation at that point was positive. Interestingly enough, Greenhalgh and Griffiths (2009) found that these straightforward stability patterns did not always occur when the BRSV model was extended to three stages. The vertical turning points of the more complicated bifurcation curves were not necessarily the points at which a change in the stability of the endemic equilibria occurred.

There would appear to be two possibilities here. Either the stability patterns observed in the two-stage BRSV model may also be found in other models exhibiting backward bifurcation but no more than two endemic equilibria or the stability patterns found in the twostage BRSV model are particular to this model. This is something that we now investigate. The CG model is ideal for this purpose since it is structurally different to the two-stage BRSV model yet provides us with another example of a system for which backward bifurcation can occur, but in which no more than two endemic equilibria are possible.

Setting the time derivatives of the system of differential equations (2.1) to (2.3) to zero and solving for I gives the following equation in  $I_e$ , the proportion of infected individuals at endemic equilibrium:

$$\beta \tilde{\beta} I_e^2 + \left( \tilde{\beta} (\psi + \mu + \alpha (1 - \gamma) - \beta) + \beta (\mu + \alpha \gamma) \right) I_e + \psi (\mu + \alpha - \tilde{\beta}) + \mu (\mu + \alpha - \beta) = 0.$$
(3.1)

From this the equation of the bifurcation curve is given by

$$\psi(I_e) = -\frac{\beta \tilde{\beta} I_e^2 + \left(\tilde{\beta}(\mu + \alpha(1 - \gamma) - \beta) + \beta(\mu + \alpha\gamma)\right)I_e + \mu(\mu + \alpha - \beta)}{\tilde{\beta} I_e + (\mu + \alpha - \tilde{\beta})}.$$
(3.2)

As we are assuming that  $\tilde{R} < 1$  (i.e. that  $\tilde{\beta} < \mu + \alpha$ ) it is clear that the denominator is non-zero for all non-negative values of  $I_e$ , so that the bifurcation curve is continuous for such values of  $I_e$ . Since, as is easily checked,  $\psi(\mu + \alpha - \tilde{\beta}) + \mu(\mu + \alpha - \beta) = 0$  when  $R_0 = 1$ ,  $I_e = 0$  is a solution to equation (3.1) when  $R_0 = 1$ . The bifurcation curve therefore passes through the coordinate  $(\psi^*, 0)$ . Then, as there can be no more than two distinct endemic equilibria for a particular value of  $\psi$ , the bifurcation diagram will exhibit backward bifurcation if, and only if, it exhibits subcritical endemic equilibria. From (3.2) there might initially appear to be the possibility for the bifurcation parameter to be negative for all  $0 \le I_e \le 1$ . However, the assumption that R > 1 (i.e. that  $\beta > \mu + \alpha$ ) implies, by continuity, that there do exist positive values of  $I_e$  for which there are positive values of  $\psi(I_e)$ .

Hadeler and Castillo-Chavez (1995) showed, by analysing the gradient of the bifurcation curve at the bifurcation point, that backward bifurcation is possible in certain circumstances. They studied

$$\frac{\psi'(0)}{\alpha+\mu} = \frac{R-\tilde{R}}{(1-\tilde{R})^2} \left\{ \frac{\alpha(1-\gamma)}{\alpha+\mu} (1-\tilde{R}) - \frac{\mu}{\alpha+\mu} \tilde{R} \right\} - R$$

by keeping the parameters  $\alpha$ ,  $\gamma$  and  $\mu$  fixed while varying  $\beta$  and  $\tilde{\beta}$  subject to the constraint  $\tilde{R} < 1 < R$ , and concluded the following:

There is a backward bifurcation at  $\psi^*$  if the following conditions are satisfied:  $\alpha$  is large in comparison with  $\mu$ ,  $\gamma$  is small,  $\tilde{R}$  is far from both 0 and 1, and R is large.

Before moving on to the stability analysis of the model, we briefly indicate how the above results could have been arrived at via a slightly simpler method. The work in this paragraph does not purport to be an in-depth analysis of the situation; it merely illustrates an alternative method for analysing the direction of the bifurcation at the bifurcation point. As  $\psi(\mu + \alpha - \tilde{\beta}) + \mu(\mu + \alpha - \beta) = 0$  when  $R_0 = 1$ , the condition for backward bifurcation to occur is that that the coefficient of  $I_e$  in (3.1) is negative when  $R_0 = 1$ . Using the fact that

$$\psi^* = \frac{\mu(\beta - \mu - \alpha)}{\mu + \alpha - \tilde{\beta}}$$

then backward bifurcation occurs if

$$\beta > \frac{\mu(\beta - \mu - \alpha)}{\mu + \alpha - \widetilde{\beta}} + \mu + \alpha(1 - \gamma) + \frac{\beta}{\widetilde{\beta}}(\mu + \alpha\gamma).$$

Following Hadeler and Castillo-Chavez, we keep  $\alpha$ ,  $\gamma$  and  $\mu$  fixed while varying  $\beta$  and  $\tilde{\beta}$  subject to the constraint  $R > 1 > \tilde{R}$ . If  $\tilde{\beta} < \mu + \alpha \gamma$  then  $\beta(\mu + \alpha \gamma)/\tilde{\beta} > \beta$ , and the above inequality can never be satisfied. This shows that  $\tilde{R}$  must not be too close to 0 if backward bifurcation is to be possible. On the other hand, if

$$\mu + \alpha - \frac{\mu(\beta - \mu - \alpha)}{\beta} < \widetilde{\beta} < \mu + \alpha \quad \text{then} \quad \frac{\mu(\beta - \mu - \alpha)}{\mu + \alpha - \widetilde{\beta}} > \beta$$

and, once more, our inequality cannot be satisfied. From this it follows that  $\tilde{R}$  must not be too close to 1 for backward bifurcation to occur.

#### 3.1 Analytic stability analysis

Hadeler and Castillo-Chavez state that a stable endemic equilibrium always exists when  $\psi$  is decreased below  $\psi^*$ . This is certainly what might be expected, noting that decreasing  $\psi$  below  $\psi^*$  is equivalent to increasing  $R_0$  above 1. However, when there is backward bifurcation the authors claim that if  $\psi$  is decreased below  $\psi^*$  then, in the presence of even a very small proportion of infected individuals, the system jumps to the upper branch, and that if  $\psi$  is increased again the system stays on the upper branch all the way to the turning point. This initial behaviour is once more as we might expect, but it is the latter phenomenon that is interesting since it would imply that the endemic equilibrium at each point on the portion of the bifurcation curve with negative gradient is LAS. However, as no accompanying analysis was provided to support this claim, it is likely that this result was obtained numerically. We first provide an analytic proof that this is in fact the case.

**Theorem 3.1.1** A particular endemic equilibrium for the CG model is LAS if, and only if, the gradient of the bifurcation curve at the point corresponding to that endemic equilibrium is negative.

Proof Equation (3.2) may be rearranged to give

$$\psi(I_e) = (\beta - \tilde{\beta}) \frac{\alpha(1 - \gamma)I_e + \mu}{\tilde{\beta}I_e + (\alpha + \mu - \tilde{\beta})} - (\beta I_e + \mu), \qquad (3.1.1)$$

so that

$$\psi'(I_e) = (\beta - \tilde{\beta}) \frac{\alpha(1 - \gamma)(\alpha + \mu - \tilde{\beta}) - \mu \tilde{\beta}}{\left\{ \tilde{\beta}I_e + (\alpha + \mu - \tilde{\beta}) \right\}^2} - \beta.$$
(3.1.2)

From this it follows that the gradient of the bifurcation curve at  $(\psi(I_e), I_e)$  is negative if, and only if,

$$\beta > (\beta - \tilde{\beta}) \frac{\alpha(1 - \gamma)(\alpha + \mu - \tilde{\beta}) - \mu \tilde{\beta}}{\left\{ \tilde{\beta} I_e + (\alpha + \mu - \tilde{\beta}) \right\}^2}.$$

Next, on linearising the system of equations (2.1) to (2.3) about the endemic equilibrium solution  $(S_e, V_e, I_e)$  we obtain the Jacobian

$$\begin{pmatrix} -\beta I_{e} - \psi - \mu & 0 & -\beta S_{e} + \alpha(1-\gamma) \\ \psi & -\tilde{\beta} I_{e} - \mu & -\tilde{\beta} V_{e} + \alpha \gamma \\ \beta I_{e} & \tilde{\beta} I_{e} & \beta S_{e} + \tilde{\beta} V_{e} - \alpha - \mu \end{pmatrix}.$$

This can be reduced this to a  $2 \times 2$  matrix by putting  $V_e = 1 - S_e - I_e$  in order to eliminate  $V_e$  from the differential equations, thereby eliminating the corresponding row and column in the Jacobian. We obtain

$$\mathbf{J} = \begin{pmatrix} -\beta I_e - \psi - \mu & -\beta S_e + \alpha (1 - \gamma) \\ \beta I_e - \tilde{\beta} I_e & \beta S_e + \tilde{\beta} (1 - S_e - 2I_e) - \alpha - \mu \end{pmatrix},$$

leading to the quadratic following characteristic equation det $[\mathbf{J} - \lambda \mathbf{I}_2] = 0$  (where  $\mathbf{I}_2$  represents the 2×2 identity matrix):

$$\left(\left\{-\beta I_{e}-\psi-\mu\right\}-\lambda\right)\left(\left\{(\beta-\tilde{\beta})S_{e}+\tilde{\beta}(1-2I_{e})-\alpha-\mu\right\}-\lambda\right)-I_{e}(\beta-\tilde{\beta})\left(-\beta S_{e}+\alpha(1-\gamma)\right)=0.$$
 (3.1.3)

From (2.3) we obtain, with C = 1 and using  $V_e = 1 - S_e - I_e$ , that

$$S_{e} = \frac{\widetilde{\beta}I_{e} + (\alpha + \mu - \widetilde{\beta})}{\beta - \widetilde{\beta}}.$$

On substituting this into equation (3.1.3) we have, after some simplification, that

$$(-\beta I_{e} - \psi - \mu - \lambda)(-\widetilde{\beta} I_{e} - \lambda) + \beta \left\{ \widetilde{\beta} I_{e} + (\alpha + \mu - \widetilde{\beta}) \right\} I_{e} - \alpha (1 - \gamma)(\beta - \widetilde{\beta}) I_{e} = 0,$$

from which it follows that

$$\lambda^{2} + \left\{ (\beta + \tilde{\beta})I_{e} + \psi + \mu \right\} \lambda + \tilde{\beta} (\beta I_{e} + \psi + \mu)I_{e} + \beta \left\{ \tilde{\beta}I_{e} + (\alpha + \mu - \tilde{\beta}) \right\} I_{e} - \alpha (1 - \gamma)(\beta - \tilde{\beta})I_{e} = 0. \quad (3.1.4)$$

The endemic equilibrium is LAS if both roots of the characteristic equation (3.1.4) have strictly negative real parts, and is unstable if at least one root has a strictly positive real part. Now both roots of  $\lambda^2 + b\lambda + c = 0$ , where b and c are real, have strictly negative real parts if, and only if, b and c are both strictly positive. Thus the endemic equilibrium is LAS if, and only if,

$$\widetilde{\beta}(\beta I_{e} + \psi + \mu) + \beta \left\{ \widetilde{\beta} I_{e} + (\alpha + \mu - \widetilde{\beta}) \right\} - \alpha (1 - \gamma)(\beta - \widetilde{\beta}) > 0$$

Substituting the expression (3.1.1) for  $\psi(I_e)$  into the above inequality gives

$$\widetilde{\beta}(\beta-\widetilde{\beta})\frac{\alpha(1-\gamma)I_{e}+\mu}{\widetilde{\beta}I_{e}+(\alpha+\mu-\widetilde{\beta})}+\beta\left\{\widetilde{\beta}I_{e}+(\alpha+\mu-\widetilde{\beta})\right\}-\alpha(1-\gamma)(\beta-\widetilde{\beta})>0,$$

which can be rearranged as

$$\beta > (\beta - \tilde{\beta}) \frac{\alpha(1 - \gamma)(\alpha + \mu - \tilde{\beta}) - \mu \tilde{\beta}}{\left\{ \tilde{\beta} I_e + (\alpha + \mu - \tilde{\beta}) \right\}^2},$$

as required.

This result parallels that found by Greenhalgh et al. (2000) for the two-stage BRSV model.

## 4 Extending the Core Group Model

We now extend, in a natural way, the CG model (that shall henceforth be termed the *basic CG model*) to obtain an *extended CG model*. This will be effected by incorporating two educated classes,  $E_1$  and  $E_2$ , into the model. It shall be assumed that individuals can be 'educated' with regard to the disease in one of two different ways. Of course, the type of education on offer would depend on the disease. It could take the form of vaccines, creams, drugs, condoms or behaviour modification through education programmes, amongst others. For a sexually transmitted disease,  $E_1$  might correspond to those individuals using a particular type of condom while  $E_2$  represents those using a type that is less effective with regard to protection from the sexually transmitted disease. Alternatively, these could be classes of individuals with reduced levels of promiscuity due to the availability of two different sex-education programmes.

Another possible interpretation is that the  $E_1$  and  $E_2$  classes respond to the education in different ways.

This does, in fact, have a highly topical flavour since the *Bill Gates Foundation* is currently offering, for projects that it deems worthy, significant funding for developing new ways of reducing the impact of major diseases (see Sam Lister (2005) and Bill & Melinda Gates Foundation (2006)). This might conceivably give rise to 'competing' forms of education for a particular infectious disease.

We assume here that membership of these educated classes is mutually exclusive in that an individual cannot simultaneously belong to both. It might of course be argued that this will not necessarily be the case and that in certain circumstances some individuals may be regarded as being in both classes. This more complicated scenario could be modelled by incorporating three educated classes into the model, the extra class containing those individuals that belong to both  $E_1$  and  $E_2$ . However, our model would be a special case of this situation and is certainly more feasible to study initially. The susceptible individuals may be regarded as those that are completely uneducated with respect to the disease. This leads to the following system of differential equations:

$$\begin{split} \frac{dS}{dt} &= \mu C - \beta_3 \frac{SI}{C} - \psi_1 S - \psi_2 S + \alpha (1 - \gamma_1 - \gamma_2) I - \mu S \\ \frac{dE_1}{dt} &= \psi_1 S - \beta_1 \frac{E_1 I}{C} + \alpha \gamma_1 I - \mu E_1, \\ \frac{dE_2}{dt} &= \psi_2 S - \beta_2 \frac{E_2 I}{C} + \alpha \gamma_2 I - \mu E_2 \\ \frac{dI}{dt} &= \frac{\beta_3 SI + \beta_1 E_1 I + \beta_2 E_2 I}{C} - \alpha I - \mu I. \end{split}$$

and

There is no need explicitly to define all of the parameters here since their definitions follow on in an obvious way from those in the basic CG model. Note that this model does not allow for direct migration between the two types of educated class, but this extra feature could be incorporated at a later stage. Our simplifying assumptions are not overly restrictive in the sense that any interesting behaviour we do observe for the above model will be a subset of all possible behaviours observable in the more general models.

As before, since the above system is homogeneous it can be normalised by setting C = 1, meaning that S,  $E_1$ ,  $E_2$  and I now represent population proportions rather than numbers of individuals, so we have  $S + E_1 + E_2 + I = 1$ . The equilibrium solutions are obtained by setting the time derivatives of the above system of equations to zero. With  $I_e$  denoting an endemic equilibrium solution, it is found that  $WI_e^3 + XI_e^2 + YI_e + Z = 0$ , where

$$\begin{split} W &= \beta_1 \beta_2 \beta_3, \\ X &= \beta_1 \beta_2 \{ \alpha (1 - \gamma_1 - \gamma_2) + \psi_1 + \psi_2 + \mu \} + \beta_1 \beta_3 (\alpha \gamma_2 + \mu) + \beta_2 \beta_3 (\alpha \gamma_1 + \mu) - \beta_1 \beta_2 \beta_3, \\ Y &= \alpha \beta_1 \{ (1 - \gamma_1) (\psi_2 + \mu) + \gamma_2 \psi_1 \} + \alpha \beta_2 \{ (1 - \gamma_2) (\psi_1 + \mu) + \gamma_1 \psi_2 \} \\ &+ \mu (\beta_1 + \beta_2) (\psi_1 + \psi_2 + \mu) - \beta_1 \beta_2 (\psi_1 + \psi_2) + \mu \beta_3 (\alpha \gamma_1 + \alpha \gamma_2 + \mu - \beta_1 - \beta_2), \\ Z &= \mu \{ (\alpha + \mu) (\psi_1 + \psi_2 + \mu) - \psi_1 \beta_1 - \psi_2 \beta_2 - \mu \beta_3 \} \\ &= \mu (\alpha + \mu) (\psi_1 + \psi_2 + \mu) \{ 1 - R_0 \}, \end{split}$$

on noting that the basic reproduction ratio is given by

$$R_0 = \frac{\psi_1 \beta_1 + \psi_2 \beta_2 + \mu \beta_3}{(\alpha + \mu)(\psi_1 + \psi_2 + \mu)}.$$

If  $\psi_1$  is used as the bifurcation parameter then the bifurcation curve is given by

$$\psi_{1}(I_{e}) = -\frac{WI_{e}^{3} + \hat{X}I_{e}^{2} + \hat{Y}I_{e} + \hat{Z}}{\beta_{1}\beta_{2}I_{e}^{2} + \{\alpha\beta_{1}\gamma_{2} + \alpha\beta_{2}(1-\gamma_{2}) + \mu(\beta_{1}+\beta_{2}) - \beta_{1}\beta_{2}\}I_{e} + \mu(\alpha + \mu - \beta_{1})}$$

where

$$\begin{split} \hat{X} &= \beta_1 \beta_2 \{ \alpha (1 - \gamma_1 - \gamma_2) + \psi_2 + \mu \} + \beta_1 \beta_3 (\alpha \gamma_2 + \mu) + \beta_2 \beta_3 (\alpha \gamma_1 + \mu) - \beta_1 \beta_2 \beta_3 ,\\ \hat{Y} &= \alpha \beta_1 (1 - \gamma_1) (\psi_2 + \mu) + \alpha \beta_2 \{ \mu (1 - \gamma_2) + \gamma_1 \psi_2 \} \\ &+ \mu (\beta_1 + \beta_2) (\psi_2 + \mu) - \beta_1 \beta_2 \psi_2 + \mu \beta_3 (\alpha \gamma_1 + \alpha \gamma_2 + \mu - \beta_1 - \beta_2) ,\\ \hat{Z} &= \mu \{ (\alpha + \mu) (\psi_2 + \mu) - \psi_2 \beta_2 - \mu \beta_3 \}. \end{split}$$

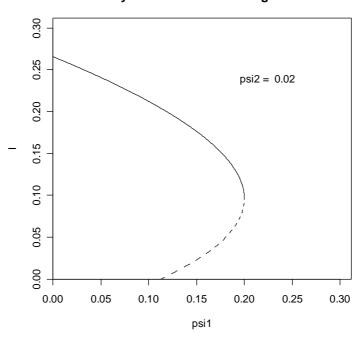
and

# 5 Bifurcation diagrams for the extended CG model

We now investigate the possibility that more complicated bifurcation diagrams, of the type observed for the three-stage BRSV model by Greenhalgh and Griffiths (2009), arise from the extended CG model. Since this paper already provides, via a detailed equilibrium analysis of the three-stage BRSV model, proofs that such bifurcation diagrams are indeed possible for epidemic models, there is no need here for further lengthy analytic existence proofs. Instead we carry out a numerical investigation in which our choice of parameter values is, to certain extent, guided by those used by Hadeler and Castillo-Chavez for the basic CG model. These were as follows:

$$\alpha = 4.0$$
,  $\mu = 0.2$ ,  $\beta = 6.0$ ,  $0 \le \tilde{\beta} \le 3.3$  and  $\gamma = 0.025$ ,

remembering that  $\psi$  is the bifurcation parameter. It is clear that  $\gamma$  is dimensionless while the remaining parameters all have the dimension time<sup>-1</sup>, although no specific units were given. In fact, the specific units used for the parameters need not concern us here as we are only really interested in their values relative to one another. In order for there to be any potential for realism in the parameter values we would certainly expect, for the extended CG model, that the transmission rates from infected individuals to the two types of educated individuals are both less than the transmission rate from infected to susceptible individuals. In other words, we shall require that  $\beta_3 > \max{\{\beta_1, \beta_2\}}$ .



**Dynamic Bifurcation Diagram** 

FIGURE 5.1: A bifurcation diagram for the extended CG model exhibiting backward bifurcation.

Programmes were written in R (2004) to plot dynamic bifurcation diagrams for which  $\psi_1$  was the bifurcation parameter and  $\psi_2$  the dynamic parameter. A point on a section of solid curve corresponds to a LAS endemic equilibrium while a point in a dashed section represents an unstable endemic equilibrium. Sets of parameter values were found that produced bifurcation diagrams exhibiting either backward bifurcation or forward bifurcation along with multiple subcritical and supercritical endemic equilibria. Figures 5.1, 5.2 and 5.3 give a sequence of snapshots taken from a dynamic bifurcation diagram. The set of parameter values given by

$$\alpha = 4.0$$
,  $\mu = 0.08$ ,  $\beta_1 = 2.5$ ,  $\beta_2 = 5.3$ ,  $\beta_3 = 6.0$ ,  $\gamma_1 = 0.06$  and  $\gamma_2 = 0.04$ 

were kept fixed throughout the sequence. The values of  $\psi_2$  used for each diagram are displayed on the plots.

**Dynamic Bifurcation Diagram** 

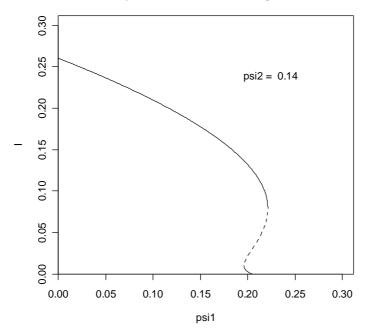


FIGURE 5.2: A bifurcation diagram for the extended CG model exhibiting forward bifurcation accompanied by two subcritical and three supercritical endemic equilibria.

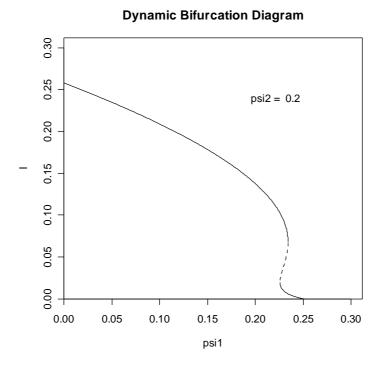


FIGURE 5.3: A bifurcation diagram for the extended CG model exhibiting forward bifurcation accompanied by three supercritical endemic equilibria, but for which no subcritical endemic equilibria are present.

The type of bifurcation curve as given in Figure 5.1 can also be observed for the basic CG model. On the other hand, curves with more complicated shapes, as shown in Figures 5.2 and 5.3, cannot be observed for the simpler version of the model. Of particular interest to us is the fact that these correspond to some of the bifurcation diagrams found for the three-stage BRSV model.

One of the necessary conditions Hadeler and Castillo-Chavez gave for backward bifurcation to occur for the basic CG model was that  $\tilde{R}$  should be far from 0, remembering that it is also assumed that  $\tilde{R} < 1$ . This corresponds to a poor education programme, and equates to  $\tilde{\beta}$  not being too small compared to  $\beta$ . On considering the sets of parameter values used for the bifurcation diagrams in Figures 5.1, 5.2 and 5.3 it would appear that we have a similar, if slightly more complicated, situation for the extended CG model. In order for multiple equilibria (rather than simply backward bifurcation) to be present it seems as though neither of  $\beta_1$  or  $\beta_2$  can be too small compared to  $\beta_3$ . In other words, it might be conjectured that for multiple equilibria to be at all possible in the extended CG model, whether via forward or backward bifurcation, neither of the education programmes or methods must be too effective. It is also worth noting that the death rate must not be too large for such curves to be possible here.

# 5.1 Stability patterns

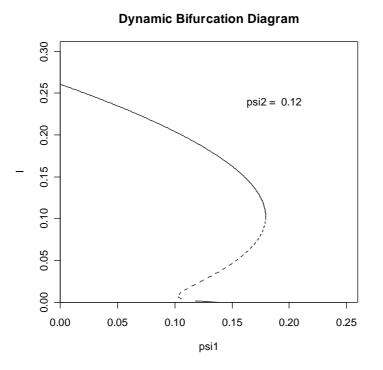


FIGURE 5.4: A bifurcation diagram demonstrating that it is possible for a particular endemic equilibrium to be unstable while the gradient of the bifurcation curve at the point corresponding to this equilibrium is negative.

Now it has been established that complex bifurcation diagrams of the type observed for the three-stage BRSV model are also possible for the extended CG model, we turn our attention to the issue of stability. On the limited evidence of the diagrams given in Figures 5.1, 5.2 and 5.3 it might be conjectured that a particular endemic equilibrium is LAS if, and only if, the gradient at the point on the bifurcation curve corresponding to this equilibrium is negative. However, not only would such a conjecture for the extended CG model be false, but we also observed patterns that we had not seen in any of the literature on backward bifurcation. Before considering this latter phenomenon we give an example to show that stability is not related to the gradient of the bifurcation curve in the simple way that it was for the basic CG model. The diagram in Figure 5.4 was obtained using the following parameter values:

$$\alpha = 4.0$$
,  $\mu = 0.016$ ,  $\beta_1 = 2.5$ ,  $\beta_2 = 5.5$ ,  $\beta_3 = 6.0$ ,  $\gamma_1 = 0.07$  and  $\gamma_2 = 0.04$ .

We now demonstrate an interesting stability pattern arising from the extended CG model. The two-stage BRSV and basic CG models have the property that for any particular set of parameter values corresponding to  $R_0 > 1$  there exists a unique LAS endemic equilibrium. By extending the two-stage BRSV model to three stages we found that it was possible for multiple supercritical endemic equilibria to exist. Furthermore, when multiple supercritical endemic equilibria are present it is quite possible that one of them is unstable. However, for all of the examples that were considered for the three-stage BRSV model it was found that at least one of the resulting supercritical equilibria is LAS. In particular, when there is a unique supercritical endemic equilibrium, it is LAS. On extending the CG model we find that this is not necessarily the case. The bifurcation diagram given in Figure 5.5 was obtained using the following set of parameter values:

$$\alpha = 4.5, \ \mu = 0.01, \ \beta_1 = 2.5, \ \beta_2 = 5.5, \ \beta_3 = 6.0, \ \gamma_1 = 0.11 \text{ and } \ \gamma_2 = 0.04.$$

It can be seen that there are no subcritical endemic equilibria but there are multiple supercritical endemic equilibria for a narrow range of values of the bifurcation parameter  $\psi_1$ . Bifurcation curves with such shapes were also observed for the three-stage BRSV model. What is new here, however, is that there exists a *unique but unstable supercritical endemic equilibrium* for a range of values of the bifurcation parameter. In other words, we have found regions of the parameter space for which  $R_0 > 1$  but for which no LAS endemic equilibrium exists. This does raise some interesting questions. Do stable limit cycles exist for these parameter values, or does some sort of chaotic behaviour result?

**Dynamic Bifurcation Diagram** 

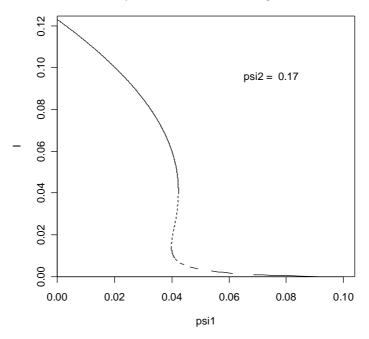


FIGURE 5.5: This bifurcation diagram demonstrates the existence of a unique and unstable supercritical endemic equilibrium for regions in the parameter space.

We emphasise here that the phenomenon observed in Figure 5.5 is not as a result of a lack of sensitivity in the plotting routine. In order to demonstrate this fact we now perform a precise numerical calculation with  $\psi_1 = 0.05$  in order to check numerically the stability result implied by the bifurcation curve in Figure 5.5. On linearising the system of equations about the endemic equilibrium solution  $(S_e, E_{e1}, E_{e2}, I_e)$  we obtain the Jacobian

$$\begin{pmatrix} -\beta_{3}I_{e} - \psi_{1} - \psi_{2} - \mu & 0 & 0 & -\beta_{3}S_{e} + \alpha(1 - \gamma_{1} - \gamma_{2}) \\ \psi_{1} & -\beta_{1}I_{e} - \mu & 0 & -\beta_{1}E_{e1} + \alpha\gamma_{1} \\ \psi_{2} & 0 & -\beta_{2}I_{e} - \mu & -\beta_{2}E_{e2} + \alpha\gamma_{2} \\ \beta_{3}I_{e} & \beta_{1}I_{e} & \beta_{2}I_{e} & \beta_{3}S_{e} + \beta_{1}E_{e1} + \beta_{2}E_{e2} - \alpha - \mu \end{pmatrix}$$

This can be reduced to a  $3 \times 3$  matrix by using  $E_{e^2} = 1 - S_e - E_{e^1} - I_e$  in order to eliminate  $E_{e^2}$  from the differential equations, thereby eliminating the corresponding row and column in the Jacobian:

$$\begin{pmatrix} -\beta_{3}I_{e} -\psi_{1} -\psi_{2} -\mu & 0 & -\beta_{3}S_{e} +\alpha(1-\gamma_{1} -\gamma_{2}) \\ \psi_{1} & -\beta_{1}I_{e} -\mu & -\beta_{1}E_{e1} +\alpha\gamma_{1} \\ (\beta_{3} -\beta_{2})I_{e} & (\beta_{1} -\beta_{2})I_{e} & \beta_{3}S_{e} +\beta_{1}E_{e1} +\beta_{2}(1-S_{e} -E_{e1} -2I_{e}) -\alpha -\mu \end{pmatrix}.$$

From the fourth equilibrium equation we have  $\beta_3 S_e + \beta_1 E_{e1} + \beta_2 E_{e2} - \alpha - \mu = 0$ , so that entry (3,3) of this matrix can be simplified considerably:

$$\beta_{3}S_{e} + \beta_{1}E_{e1} + \beta_{2}(1 - S_{e} - E_{e1} - 2I_{e}) - \alpha - \mu = \beta_{3}S_{e} + \beta_{1}E_{e1} + \beta_{2}E_{e2} - \beta_{2}I_{e} - \alpha - \mu$$
$$= -\beta_{2}I_{e}.$$

The characteristic equation is thus given by

$$\det \begin{pmatrix} -\beta_3 I_e - \psi_1 - \psi_2 - \mu - \lambda & 0 & -\beta_3 S_e + \alpha (1 - \gamma_1 - \gamma_2) \\ \psi_1 & -\beta_1 I_e - \mu - \lambda & -\beta_1 E_{e1} + \alpha \gamma_1 \\ (\beta_3 - \beta_2) I_e & (\beta_1 - \beta_2) I_e & -\beta_2 I_e - \lambda \end{pmatrix} = 0.$$

At this point we note that a study of the local asymptotic stability of the endemic equilibria could, in theory, be performed by analysing the coefficients of the resultant cubic equation in  $\lambda$  via the Routh-Hurwitz criteria. However, we already know that no straightforward stability result will be forthcoming, and so this is not pursued here.

The bifurcation parameter  $\psi_1$  was set at 0.05, giving  $R_0 = 1.079726...$  With the given parameter values the unique endemic equilibrium is calculated to be

$$(S_{_e}, E_{_{e1}}, E_{_{e2}}, I_{_e}) = (0.091709..., 0.339074..., 0.565830..., 0.003387...) \,.$$

These equilibrium values have been written as 0.091709..., and so on, to emphasise the fact that all calculations were performed to the highest possible degree of accuracy in order to avoid coming to false conclusions through the accumulation of rounding errors. When substituting these into the characteristic equation we find that two of the eigenvalues have a real part approximately equal to 0.00348, showing that this endemic equilibrium is indeed unstable. This instability does occur at very low levels of infection, and over a relatively narrow range of values of the bifurcation parameter (and, as a consequence, over a particularly narrow range of values of  $R_0$ ). This phenomenon is investigated further in the following section.

Before moving on to consider the situation from a stochastic point of view, we note that the work considered in the current section poses some interesting questions with regard to dynamic phenomena. Indeed, these might be worth pursuing in a subsequent paper. The *Principle of Exchange of Stability* (see Boldin (2006), for example) is certainly relevant here. This is valid when one stationary state loses its stability to another stationary state in a smooth way. At a saddle-point bifurcation (also known as a tangential bifurcation or turning point bifurcation) one of the eigenvalues is zero, and in this case the stability of the branches representing the various endemic equilibria will depend on the value of the other eigenvalues at the turning point (see Kuznetsov (2004) for details). It might be possible to look at the nongeneric case of a vertical bifurcation and unfold it with two parameters (see van den Bosch *et al.* (1988) and Kuznetsov (2004) once more). The dimension of the state space also needs to be taken into account since Hopf bifurcation may occur when this dimension is at least two. Indeed, it may be the case that Hopf bifurcation is present in Figure 5.5.

#### 6 Probability of extinction for the basic CG model

We now calculate, for the stochastic version of the basic CG model, the theoretical probability of extinction, given that one infected individual is introduced into the population at the DFE. Before proceeding with the calculation, it is worth pointing out that in general the probability of extinction does depend on the initial conditions. If, for example, the initial number of infected individuals was greater than one then we would expect the corresponding probability of extinction to be reduced. Should  $R_0$  be large then the introduction of even one infected individual into the population could lead to a huge outbreak of the disease in the population. If the spread of the disease was particularly rapid then this might give rise to a bottle neck in the supply of susceptible individuals, which may in turn cause the pathogen to go extinct. Since we are only considering here values of  $R_0$  close to one, the possibility of 'extinction after the first outbreak' is not particularly relevant here, although an account of this phenomenon can be found in Taxidis (2008).

At the DFE the state of the system is given by

$$(S_{DFE}, V_{DFE}, I_{DFE}) = \left(\frac{\mu}{\mu + \psi}, \frac{\psi}{\mu + \psi}, 0\right).$$

Let us consider the discrete branching process  $(\hat{X}_k)$  consisting of the *k*th generation size of the continuous branching process (I(t)), given that  $\hat{X}_0 = I(0) = 1$ . The size of the first generation,  $\hat{X}_1$ , is the total number of individuals that have been directly infected by the initially infected individual, then  $\hat{X}_2$  gives the total number of individuals that have been directly infected by the  $\hat{X}_1$  individuals in the first generation, and so on. We note that I(t) = 0 for some time *t* if, and only if,  $\hat{X}_k = 0$  for some generation *k*, and so the processes (I(t)) and  $(\hat{X}_k)$  share the same extinction probabilities (see Grimmett and Stirzaker (2001), p. 177).

We next formulate the probability generating function for  $(\hat{X}_{k})$ . Let

$$\Delta = \beta S_{DFE} + \tilde{\beta} V_{DFE} + \alpha + \mu = \frac{\beta \mu}{\mu + \psi} + \frac{\tilde{\beta} \psi}{\mu + \psi} + \alpha + \mu .$$

It shall be assumed that the length of time that an infected individual remains infectious follows a negative exponential distribution with mean  $1/(\alpha + \mu)$ . Let N be the random variable representing the number of individuals directly infected by one infected individual through its period of infectiousness. It is clear that  $P(N = 0) = (\alpha + \mu)/\Delta$ . We also have

$$P(N=1) = \frac{\Delta - (\alpha + \mu)}{\Delta} \times \frac{\alpha + \mu}{\Delta}$$

Continuing in this way

$$P(N=n) = \frac{\alpha + \mu}{\Delta} \left( \frac{\Delta - (\alpha + \mu)}{\Delta} \right)^n,$$

from which it follows that the probability generating function of the number of individuals infected by one infected individual:

$$G(x) = E(x^{N}) = \frac{\alpha + \mu}{\Delta} \sum_{n=0}^{\infty} \left\{ \left( \frac{\Delta - (\alpha + \mu)}{\Delta} \right) x \right\}^{n} = \frac{\alpha + \mu}{\Delta} \left\{ 1 - \left( \frac{\Delta - (\alpha + \mu)}{\Delta} \right) x \right\}^{-1}$$

The probability of extinction,  $P(\infty)$ , is given by the smallest positive solution to x = G(x). Solving the equation

$$x = \frac{\alpha + \mu}{\Delta} \left\{ 1 - \left(\frac{\Delta - (\alpha + \mu)}{\Delta}\right) x \right\}^{-1}$$

gives

$$x = 1$$
 or  $x = \frac{\alpha + \mu}{\Delta - (\alpha + \mu)} = \frac{(\alpha + \mu)(\psi + \mu)}{\beta \mu + \tilde{\beta} \psi} = \frac{1}{R_0}$ 

Thus extinction is certain unless  $R_0 > 1$ , in which case  $P(\infty) = 1/R_0$ .

#### 6.1 Stochastic simulations

Stochastic simulations were carried out in order to see whether the presence of backward bifurcation might lead to results different from the analytic probabilities of extinction given above. Programmes in C++ were developed in order to perform the simulations, and we began by testing the programmes for situations in which forward bifurcation is present. The important points made by Griffiths (2007) concerning stochastic simulations apply here also. These include details relating to the population size, run times, random number generation and the definition of *quasi-stationary distributions conditional on non extinction* (see also Renshaw (1991), Clancy *et al.* (2001), Dickman and Vidigal (2002), Bean *et al.* (1997) and Pollett

(2004)). Using similar parameter values to those used in the paper by Hadeler and Castillo-Chavez, it is found that forward bifurcation occurs when:

$$\alpha = 3.0, \ \mu = 0.2, \ \beta = 8.0, \ \beta = 0.4 \text{ and } \gamma = 0.025.$$

The corresponding bifurcation diagram is given in Figure 6.1.1, noting that the unique supercritical endemic equilibrium is always stable, as given by Theorem 3.1.1 for the basic CG model. We have that  $\psi^* = \psi(R_0 = 1) = 0.3428...$  here, as can be seen in Figure 6.1.1.

Table 6.1.1 compares, for a range of values of the bifurcation parameter  $\psi$ , the theoretical values for  $P(\infty)$  with those obtained via the simulations. The population size and the number of simulations used to estimate each of these extinction probabilities were 5000 and 4000 respectively. By using a normal approximation to the binomial distribution we may construct approximate confidence intervals from each of the simulated probabilities of extinction in Table 6.1.1 to see if they contain the corresponding theoretical values. For example, an approximate 90% confidence interval for the probability of extinction when  $\psi = 0.2$  is given by [0.7456,0.7680], which does contains the theoretical value. In fact, as is easily checked, all such confidence intervals contain the corresponding theoretical value for  $P(\infty)$ . We see therefore that there is excellent agreement between theoretical and simulated outcomes in this case. It turns out that the simulated values for the probability of extinction do generally match the theoretical values very closely for sets of parameter values leading to forward bifurcation.



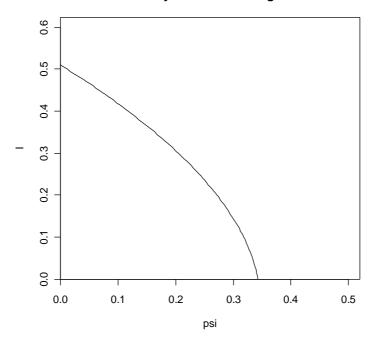


FIGURE 6.1.1: A bifurcation diagram for the basic CG model, exhibiting forward bifurcation.

Ψ	$R_{_0}$	Theoretical $P(\infty)$	Simulated $P(\infty)$
0.5	0.8036	1.0000	1.0000
0.4	0.9167	1.0000	1.0000
0.3	1.0750	0.9302	0.9310
0.2	1.3125	0.7619	0.7568
0.1	1.7083	0.5853	0.5805
0.05	2.0250	0.4938	0.4818
0.0	2.5000	0.4000	0.4113

TABLE 6.1.1: A comparison of theoretical and simulated probabilities of extinction for the basic CG model using a set of parameter values leading to forward bifurcation.

It should be noted that increasing the population size from 5000 does not, for the values of  $\psi$  used in Table 6.1.1, significantly affect the simulated extinction probabilities. In fact, the population size needs to be reduced considerably for these probabilities to be altered in any way. This is because, as can be seen from Figure 6.1.1, the levels of disease prevalence within the population are relatively high for this particular set of parameter values, even when  $R_0$  is reasonably close to 1. Therefore even fairly small populations will have a relatively large number of infected individuals at quasi-equilibrium, thereby reducing the possibility that a particular random fluctuation will result in extinction.

### **Stability Bifurcation Diagram**

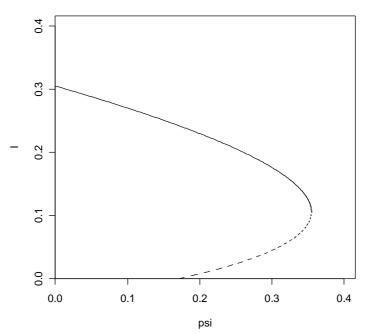


FIGURE 6.1.2: A bifurcation diagram for the basic CG model, exhibiting backward bifurcation.

We now turn our attention to the situation in which backward bifurcation is present. The following set of parameter values lead to the bifurcation diagram given in Figure 6.1.2, noting that the stability pattern is as predicted by Theorem 3.1.1:

$$\alpha = 4.0$$
,  $\mu = 0.1$ ,  $\beta = 6.0$ ,  $\beta = 3.0$  and  $\gamma = 0.02$ .

A series of simulations was carried out with  $\psi = 0.18$ . This gives  $R_0 = 0.9930$ , so the theoretical probability of extinction for this set of parameter values is 1. Table 6.1.2 gives the simulated probability of extinction as the population size N is varied. Each probability of extinction was estimated using 10000 simulations. As well as corroborate the results obtained for the probabilities of extinction for the two-stage BRSV model by Griffiths (2007), our findings offer a fascinating glimpse of the subtle link between deterministic and stochastic phenomena associated with our model.

It is possible to provide heuristic explanations for the interesting relationship observed in Table 6.1.2 between the population size and the probability of extinction estimated via the simulations. As can be seen from the bifurcation curve in Figure 6.1.2, there exists, for  $\psi = 0.18$ , an upper LAS subcritical endemic equilibrium and an unstable lower one. The unstable endemic equilibrium lies on the separatrix for the domains of attraction for the LAS DFE and the LAS endemic equilibrium. It would seem plausible that these deterministic domains of attraction correspond in some way to probabilistic domains of attraction (for the absorbing state I(t) = 0 and the quasi-equilibrium respectively) in the stochastic version of the model. The probability of extinction may be related to both the likelihood with which these domains of attraction can be reached and, in cases for which a quasi-equilibrium has been attained, the ability to avoid an extinction due to random fluctuations.

For particularly small population sizes, the low numbers of individuals at quasiequilibrium make it relatively likely that, over ecologically relevant periods of time (i.e. over several generations), a random fluctuation will result in extinction (since in this case any quasiequilibrium will be close to the absorbing state). Thus, as when considering the possibility of a supercritical endemic invasion when forward bifurcation is present, we would expect the simulated probability of extinction to tend to 1 as the population size is decreased.

However, things are very different to the previous example when the population is increased rather than decreased. When considering the possibility of a subcritical invasion of the disease in the presence of backward bifurcation, it needs to be borne in mind that the probabilistic domain of attraction for the quasi-equilibrium tends to get further from the absorbing state as the population size is increased. Thus, on the introduction of one infected individual into the population at the DFE, the larger the population size the less opportunity there is for initial random fluctuations to result in a state lying in this domain of attraction. For very large population sizes this domain of attraction becomes virtually impossible to reach, and extinction is very likely to occur.

From the above discussion it would seem that when backward bifurcation is present, subcritical probabilities of extinction will always tend to 1 as the population becomes either extremely large or extremely small. It is the intermediate population sizes that are of more interest. In these cases the population is still sufficiently small for there to be a realistic prospect of a quasi-equilibrium being attained, while not so small that this quasi-equilibrium is not sustainable over ecologically relevant periods of time.

Population size N	Simulated $P(\infty)$	
50	1.0000	
100	0.9963	
150	0.9846	
200	0.9840	
250	0.9766	
300	0.9718	
400	0.9702	
500	0.9739	
800	0.9797	
1000	0.9800	
2000	0.9834	
5000	0.9870	
10000	0.9919	

TABLE 6.1.2: An illustration of how, for a set of parameter values allowing subcritical endemic equilibria to exist, the estimated probability of extinction varies with the population size.

Population size N	90% confidence interval	99% confidence interval
300	[0.9691, 0.9745]	[0.9675, 0.9761]
400	[0.9674, 0.9730]	[0.9658, 0.9746]
500	[0.9713, 0.9765]	[0.9698, 0.9780]
800	[0.9774, 0.9820]	[0.9761, 0.9833]
1000	[0.9777, 0.9823]	[0.9764, 0.9836]

TABLE 6.1.3: A series of confidence intervals for the extinction probabilities

We may calculate confidence intervals for the simulated values of  $P(\infty)$  shown in Table 6.1.2 in order to add further weight to our argument. Table 6.1.3 provides, for some of the intermediate population sizes, both 90% and 99% confidence intervals for these probabilities. The results in Table 6.1.2 seem to indicate the existence of an 'optimal' population size for which the probability of extinction is least. This will be some function of the parameter values and consequently of the level of the disease in the population at quasi-equilibrium. For our particular example the prevalence of the disease can be sustained in a relatively small population. For a situation in which the proportion of infected individuals is at a lower level at endemic equilibrium it is likely that this optimal population size will be greater.

## 7 Probability of extinction for the extended CG model

Further programmes were developed to allow stochastic simulations of the extended CG model. Of particular interest here is the probability of extinction for the case in which there is a unique but unstable supercritical endemic equilibrium in the deterministic version of the model. For  $R_0 > 1$  the theoretical probability of extinction for the stochastic version of the extended CG model is given by

$$P(\infty) = \frac{(\alpha + \mu)(\psi_1 + \psi_2 + \mu)}{\psi_1 \beta_1 + \psi_2 \beta_2 + \mu \beta_3}.$$

The calculation is similar to that for the probability of extinction for the basic CG model, noting that now  $\Delta = \beta_1 E_{DFE1} + \beta_2 E_{DFE2} + \beta_3 S_{DFE} + \alpha + \mu$ . By considering Figure 5.5 it can be seen that the parameter values

$$\psi_1 = 0.05, \ \psi_2 = 0.17, \ \alpha = 4.5, \ \mu = 0.01, \ \beta_1 = 2.5, \ \beta_2 = 5.5 \ \beta_3 = 6.0, \ \gamma_1 = 0.11 \text{ and}$$
  
 $\gamma_2 = 0.04$ 

do in fact lead to a unique but unstable supercritical endemic equilibrium. For this set of parameter values  $P(\infty) = 0.9262$ .

The stochastic simulations were initialised by introducing one infected individual into the population at the DFE, given by

$$(S_{DFE}, E_{DFE1}, E_{DFE2}, I_{DFE}) = \left(\frac{\mu}{\psi_1 + \psi_2 + \mu}, \frac{\psi_1}{\psi_1 + \psi_2 + \mu}, \frac{\psi_2}{\psi_1 + \psi_2 + \mu}, 0\right).$$

Population sizes ranging from one hundred to one hundred thousand were used. The presence of the absorbing state I = 0 in conjunction finite and approximately constant population sizes means that the stochastic model cannot possess a true endemic stationary distribution. However, when the parameter values are such that a deterministic LAS endemic equilibrium exists, the stochastic versions of both the basic and extended CG models do frequently appear to settle down into some sort of steady-state behaviour, resulting in what we have previously termed quasi-equilibria. In these cases it was found that the disease is able to persist within the population for extremely long time periods. Furthermore, these quasi-equilibria tend to match the LAS endemic equilibrium very well indeed. This relationship between the deterministic and stochastic versions of the CG models was also observed when studying the corresponding versions of the BRSV models.

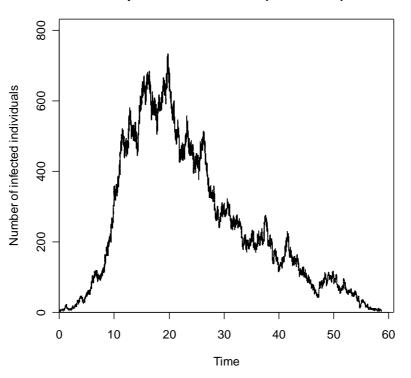
We may clarify this quasi-equilibrium phenomenon by way of an example. A series of simulations of the stochastic version of the basic CG model was performed using the following parameter values (those leading to the bifurcation diagram shown in Figure 6.1.1):

$$\alpha = 3.0, \ \mu = 0.2, \ \beta = 8.0, \ \beta = 0.4 \text{ and } \gamma = 0.025,$$

with the bifurcation parameter  $\psi$  set at 0.3. As can be seen in Figure 6.1.1, there is a unique deterministic supercritical endemic equilibrium for this set of parameter values, and the proportion of infected individuals at this equilibrium is approximately 10%. The simulations were initialised by introducing one infected individual into a population of 1000 individuals at the DFE. On simulations for which extinction did not occur we saw evidence for the existence of a quasi-equilibrium for which the average number of infected individuals over time is approximately 100, or 10% of the population, demonstrating a quantitative matching with the deterministic endemic equilibrium. Despite the relatively small number of infected individuals, it was found that once the system had reached this quasi-equilibrium then random fluctuations tended not to lead to extinction over the considerable run times of our simulations, thereby implying the presence of a quasi-stationary distribution. There were occasions when a random fluctuation resulted in extinction from an apparent quasi-equilibrium, but these proved to be extremely rare occurrences, and such events became even scarcer as the population size was increased. It would appear that an infected state is acting as a weak attractor in a probabilistic sense.

By way of contrast, we now consider the outcomes of simulations of the stochastic version of the extended CG model for the situation in which there exists a unique but unstable supercritical endemic equilibrium in the deterministic version (with parameter values given at the beginning of this section). Each of the simulations was initialised by introducing one infected individual into a population of 10000 individuals at the DFE. A larger population size was used here than since the proportion of infected individuals at the deterministic endemic equilibrium is relatively small, and we do require the absolute number of infected individuals in the population corresponding to this endemic equilibrium to be large enough to determine whether or not there is any evidence of quasi-equilibrium behaviour.

The outcomes of our simulations are potentially very interesting. With a population size of 10000, there are approximately 34 infected individuals at the unstable deterministic endemic equilibrium. However, in none of our simulations did we observe any quasi-equilibrium behaviour for which the average number of infected individuals was approximately this number. On the other hand, potential quasi-equilibrium behaviour was observed for which the average number of infected individuals was considerably greater than 34, typically of the order of 500 individuals. We found either that extinction occurred almost immediately or the number of infected individuals built up over time to several hundred. We may, on the basis of these observations, ponder over the mechanism driving this apparent quasi-equilibrium and then consider whether or not there is a quasi-stationary distribution present. In what follows we may merely speculate on the answers to these questions, and provide some heuristic explanations. Indeed, a more detailed analysis could prove to be extremely difficult.



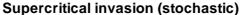


FIGURE 7.1: A realisation of a stochastic supercritical invasion of the disease for the extended CG model where the parameters are such that there exists a unique unstable deterministic supercritical endemic equilibrium.

We considered those realisations for which extinction did not occur in the very early stages of the run. Similar initial behaviour was observed for each of these realisations, with invasions resulting in up to 700 infected individuals being present in the population. In these situations it was almost always the case that, once the number of infected individuals had peaked, a decline occurred, eventually resulting in extinction. An example of such a realisation is shown in Figure 7.1. There were also isolated examples of the disease persisting at these higher levels for slightly longer periods of time. However, even in these instances, it would seem that the system does not remain at the quasi-equilibrium for long enough to indicate the presence of a quasi-stationary distribution.

Having established, empirically at least, that a quasi-stationary distribution is unlikely in this case, we may turn our attention to the business of speculating as to which aspect of the model is driving this behaviour. Let us consider, once more, the bifurcation diagram given in Figure 5.5. The geometry of the bifurcation curve means that  $\psi_1$  does not have to be reduced much below 0.05 before there exists a deterministic endemic equilibrium that is both LAS and at a considerably higher level than the unstable endemic equilibrium present when  $\psi_1 = 0.05$ . In fact, if  $\psi_1 = 0.04$  there is a deterministic LAS endemic equilibrium for which approximately 6% of the population consists of infected individuals, corresponding to 600 infected individuals in a population of 10000 individuals. This could, of course, be a numerical coincidence, but we cannot rule out the possibility that the absence of a deterministic LAS endemic equilibrium when  $\psi_1 = 0.05$  leads to stochastic behaviour in which the system is weakly attracted to a quasi-equilibrium corresponding to a nearby deterministic LAS endemic equilibrium. If this were the case then the lack of a quasi-stationary distribution may be a manifestation of the fact that the parameter values do not quite correspond to this deterministic LAS endemic equilibrium. Although, initially, the changes in population structure caused by the invasion of the disease favour further invasion, the parameter values are such that this cannot be maintained.

We now consider the problem of comparing the simulated and theoretical extinction probabilities for the stochastic version of the extended CG model with parameter values such that a unique unstable deterministic supercritical endemic equilibrium is present. The probabilities of extinction are calculated analytically under the assumption that the population size is effectively infinite in the sense that our supply of susceptible and educated individuals is not only inexhaustible but also always in the same ratio, that at the DFE. However, in the stochastic simulations the population sizes are finite and the number of individuals in each class is allowed to change, as would be the case in a real-life epidemic. Therefore the proportions of susceptible and educated individuals will not generally remain the same as that at the DFE during the course of the simulation. To a large extent, however, the extinction probabilities are determined by the behaviour of the epidemic in the initial stages of an invasion and it is generally possible to obtain a very close match between the simulated and analytic extinction probabilities, even by using relatively short run times in the simulations. In fact, so long as the run times were not too short, we found that our simulations gave estimates that were quite close to the theoretical value 0.9262 obtained from the parameter values given at the beginning of this section. This was also true when obtaining estimates for the probability of extinction for the basic CG model with parameter values as given earlier. In this case estimates close to the analytic extinction probability of 0.9302 resulted from even quite short run times.

On the other hand, from the definition of the probability of extinction, we might expect to obtain ever better agreement between the simulated and theoretical extinction probabilities by extending the run times of our simulations. However, we also know that for finite and approximately constant population sizes, as in our simulations, extinction is bound to occur eventually. It is at this point that we need to take account of whether or not an apparent quasiequilibrium is able to persist for ecologically relevant periods of time. When a quasi-stationary distribution is present, then, so long as the population size (and hence the number of infected individuals at quasi-equilibrium) is not too small, we may increase the run times considerably in order to obtain more accurate estimates for the theoretical probability of extinction. However, the presence of phenomena such as limit cycles in the deterministic model might mean that increasing the run time gives considerably more opportunities for random fluctuations to result in extinction, and in this case simulated probabilities of extinction are bound to show a corresponding increase. This latter scenario does appear to be the case for the particular set of parameter values we are currently considering for the extended CG model. Simulations do indeed bear this out, with the simulated probability of extinction tending to 1 as the run times increase. Thus the degree to which the simulated and theoretical extinction probabilities agree can, in such situations, be very sensitive with regard to the run times.

This would appear to be a difficult problem to resolve completely, and indeed it may be worth considering alternative approaches such as those adopted by Kuske et al. (2008) or Allen and van den Driessche (2006). For example, in the latter paper the authors utilise a stochastic differential equation model, noting that their deterministic and stochastic models agree very well when  $N \ge 1000$ . There is also the possibility that  $R_0$  may depend on N in some way. Indeed, Nasell (2002) notes that, with regard to the stochastic version of the model he is studying, there are three identifiable regions of the parameter space with qualitatively different behaviours, and the boundaries between these regions are dependent on N. He also acknowledges that this is a mathematically challenging problem.

# 8 Expected time to extinction for the basic CG model

For the two-stage BRSV model Griffiths (2007) found some interesting results in this regard, although, because of the analytic complications, it was not possible to state any definitive results. However, the current situation is more promising since we are able to obtain a precise analytic result for the expected time to extinction given that one infected individual enters the population at the DFE. Note that this is in contrast to the approach adopted by Nasell (2002), who calculates the time to extinction from the quasi-stationary distribution.

Let T be the random variable representing the time to extinction given that one infected individual is introduced into the population at the DFE. Then, using a result from Karlin and Taylor (1975, pp. 149-150), it follows that

$$E(T) = \frac{\mu + \psi}{\beta \mu + \widetilde{\beta} \psi} \ln \left( \frac{\alpha + \mu}{\alpha + \mu - \frac{\beta \mu + \widetilde{\beta} \psi}{\mu + \psi}} \right) = \frac{\mu + \psi}{\beta \mu + \widetilde{\beta} \psi} \ln \left( \frac{(\alpha + \mu)(\mu + \psi)}{(\alpha + \mu)(\mu + \psi) - \beta \mu + \widetilde{\beta} \psi} \right)$$

This may also be written as

$$\mathrm{E}(T) = \frac{1}{R_0(\alpha + \mu)} \ln\left(\frac{1}{1 - R_0}\right),$$

which is not defined for  $R_0 \ge 1$ , in which case we write  $E(T) = \infty$ . We are thus interested in only the subcritical case here. In particular we consider whether the presence of subcritical endemic equilibria exert an influence on E(T) to the extent that significant differences between the theoretical and simulated values are observed.

We first demonstrate that our programmes are working correctly in the sense that they lead to simulated estimates for E(T) that are in good agreement with the corresponding analytic values under what we may regard as 'ideal' conditions. We start therefore with the following set of parameter values leading to forward bifurcation:

$$\alpha = 3.0, \ \mu = 0.2, \ \beta = 8.0, \ \widetilde{\beta} = 0.4 \text{ and } \gamma = 0.025.$$

The corresponding bifurcation diagram is given in Figure 6.1.1, and  $\psi(R_0 = 1) = 0.3429$  in this case. Various values of  $\psi$  such that  $R_0 < 1$  were used in the simulations, and the results of four successive runs using each of these values are shown in Table 8.1. The population size used for each of these simulations was 3000, and each estimate for E(T) was obtained by taking the mean extinction time from 5000 runs. It can be seen that there is very good agreement between the simulated and theoretical values given in Table 8.1.

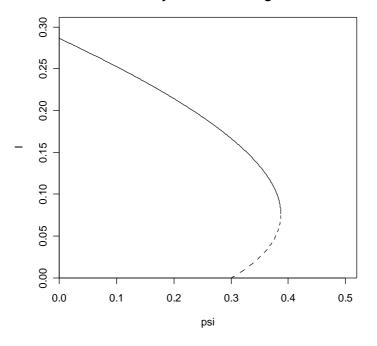
Ψ	$R_0$	Theoretical $E(T)$	Simulated $E(T)$
0.5	0.8036	0.6329	0.6284, 0.6483, 0.6294, 0.6301
0.75	0.6250	0.4904	0.4900, 0.4879, 0.4921, 0.4864
1.0	0.5208	0.4414	0.4361, 0.4542, 0.4431, 0.4544

TABLE 8.1: A comparison of theoretical and simulated expected times to extinction for the basic CG model using a set of parameter values leading to forward bifurcation.

We next run simulations for the system with parameter values given by:

$$\alpha = 4.0$$
,  $\mu = 0.2$ ,  $\beta = 6.0$ ,  $\tilde{\beta} = 3.0$  and  $\gamma = 0.025$ .

The corresponding bifurcation diagram, exhibiting backward bifurcation, is shown in Figure 8.1.



# **Stability Bifurcation Diagram**

FIGURE 8.1: A bifurcation diagram for the basic CG model exhibiting backward bifurcation.

It can be shown that the vertical turning point has coordinate  $(\psi(I_e^{\#}), I_e^{\#})$  where

$$I_{e}^{*} = \frac{1}{\widetilde{\beta}} \left( \sqrt{\frac{\beta - \widetilde{\beta}}{\beta}} \left\{ \alpha (1 - \gamma)(\alpha + \mu - \widetilde{\beta}) - \mu \widetilde{\beta} \right\} - (\alpha + \mu - \widetilde{\beta}) \right)$$

although the estimate of  $\psi(I_e^{\#}) \approx 0.4$  from the bifurcation diagram will certainly suffice for present purposes.

We compare theoretical and analytic values for E(T) using various values of  $\psi$  greater than  $\psi(I_e^{\#})$ . In particular, we are interested in determining whether or not there is any evidence that the deterministic LAS subcritical endemic equilibrium present when  $\psi$  is slightly smaller than  $\psi(I_e^{\#})$  is able to influence the stochastic behaviour of the system for slightly larger values of  $\psi$  (for which no deterministic endemic equilibria are present). Indeed, if this were the case, and an ecologically unstable quasi-equilibria occasionally resulted when  $\psi$  was just above  $\psi(I_e^{\#})$ , then we might expect the simulated values of E(T) to be significantly greater than the corresponding analytic values. The results of the simulations, using the parameter values leading to the bifurcation diagram in Figure 8.1, are given in Table 8.2. The conditions for each simulation were the same as those used to obtain the results in Table 8.1.

Ψ	$R_0$	Theoretical $E(T)$	Simulated $E(T)$
0.42	0.9447	0.7296	0.7664, 0.7680, 0.7515, 0.7704
0.44	0.9375	0.7041	0.7534, 0.7366, 0.7113, 0.7241
0.6	0.8929	0.5956	0.6168, 0.5862, 0.6082, 0.6140
1.0	0.8333	0.5119	0.5116, 0.5119, 0.5063, 0.5174

TABLE 8.2: A comparison of theoretical and simulated expected times to extinction for the basic CG model using a set of parameter values leading to backward bifurcation.

In Table 8.2, it may be seen that values of  $\psi$  just above  $\psi(I_e^*)$  tend to lead to the simulations giving significant overestimates for the theoretical values of E(T), but as we increase  $\psi$  there is ever better agreement between the simulated and theoretical values. We feel that these results, though not totally conclusive, do lend weight to our argument concerning the possibility that ecologically unstable quasi-equilibria are present for values of  $\psi$  just above  $\psi(I_e^*)$ , and that the presence of a nearby deterministic endemic equilibrium does manifest itself as a very weak attractor for the stochastic system. Indeed, we found on considering a number of

individual runs, that the overestimates mentioned above, when  $\psi$  was just above  $\psi(I_e^*)$ , tended to be caused by a very small proportion of relatively long-lived, but still ecologically unstable, quasi-equilibria.

# 9 More unusual bifurcation diagrams

We have thus far only analysed the system consisting of an isolated core population of constant size. It is now time to consider the population as a whole. We expand on some of the interesting ideas that were briefly considered by Hadeler and Castillo-Chavez towards the end of their paper. We make the point here that much of the content of this particular section is not new. However, due to the brevity of the discussion in the original paper we feel that these ideas do deserve to be explained in a little more detail, and ought also to be considered within the context of the extended CG model.

The basic CG model arose from the following more general model:

$$\frac{dA}{dt} = \mu P - Ar(I,C) - \mu A \tag{9.1}$$

$$\frac{dS}{dt} = Ar(I,C) - \beta \frac{SI}{C} - \psi S + \alpha (1-\gamma)I - \mu S, \qquad (9.2)$$

$$\frac{dV}{dt} = \psi S - \tilde{\beta} \frac{VI}{C} + \alpha \gamma I - \mu V$$
(9.3)

$$\frac{dI}{dt} = \frac{\beta SI + \tilde{\beta} VI}{C} - \alpha I - \mu I , \qquad (9.4)$$

where the total population size *P* is constant but the core group size *C* is not, in general, constant. We have that P = A + C = A + S + V + I where *A* is the sexually inactive non-core group. The function r(I,C) describes the rate of recruitment into the core group from the non-core group. We may, once more, look for endemic equilibria by setting the time derivatives to zero. If  $(A_e, S_e, V_e, I_e)$  is an endemic equilibrium of the above system then  $C_e = S_e + V_e + I_e = P - A_e$ . Equation (9.1) gives  $\mu(P - A_e) - A_e r(I_e, C_e) = 0$ , from which we see that  $A_e r(I_e, C_e) = \mu C_e$ . On substituting this into (9.2) we obtain the following set of equilibrium equations:

$$\mu C_e - \beta \frac{S_e I_e}{C_e} - \psi S_e + \alpha (1 - \gamma) I_e - \mu S_e = 0$$
  
$$\psi S_e - \tilde{\beta} \frac{V_e I_e}{C_e} + \alpha \gamma I_e - \mu V_e = 0$$

and

$$\frac{\beta S_e I_e + \beta V_e I_e}{C_e} - \alpha I_e - \mu I_e = 0$$

and

which is, except for the fact that C is now  $C_e$  instead of 1, identical to the equilibrium equations for the isolated core group.

Our previous equilibrium results for the basic CG model have been in terms of proportions of individuals in the core group. In order to use these results to obtain information about absolute numbers of individuals at equilibrium, we adopt notation allowing us to distinguish between proportions in the core and absolute numbers. In keeping with Hadeler and Castillo-Chavez we use an overbar to denote proportions in the core group, so that

$$\overline{S} = \frac{S}{C}, \ \overline{V} = \frac{V}{C} \text{ and } \overline{I} = \frac{I}{C}.$$

Thus if  $(\overline{S}_e, \overline{V}_e, \overline{I}_e)$  is an equilibrium solution for proportions in the core then an equilibrium solution for the core group of the above model will be  $(C_e \overline{S}_e, C_e \overline{V}_e, C_e \overline{I}_e)$ . From  $A_e r(I_e, C_e) = \mu C_e$  we obtain  $(P - C_e) r(I_e, C_e) = \mu C_e$ , giving us

$$C_e = \frac{r(I_e, C_e)P}{r(I_e, C_e) + \mu}.$$

Let us now assume that the function r(I,C) is in fact a function of the proportion of infected individuals in the core group,  $\overline{I} = I/C$ , so that we may write r(I,C) as  $r(\overline{I})$ . This might be the case if the figure reported, with regard to the prevalence of the disease, was the proportion of infected individuals in the core group rather than the absolute number of infected individuals. We then have that

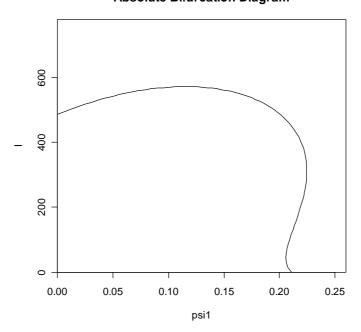
$$C_e = \frac{r(\bar{I}_e)P}{r(\bar{I}_e) + \mu} = P\left\{1 - \frac{\mu}{r(\bar{I}_e) + \mu}\right\},\,$$

which shows that  $C_e$  is an increasing function of  $r(\bar{I}_e)$ . It would seem sensible to expect that the rate of recruitment into the core decreases when the proportion of infected in the core group increases, so that  $r(\bar{I})$  is a decreasing function of  $\bar{I}$ . If this were the case then we see that  $C_e$  is a decreasing function of  $\bar{I}_e$ . The absolute number of infected individuals at equilibrium is given by

$$I_e = C_e \bar{I}_e = \frac{r(\bar{I}_e)P\bar{I}_e}{r(\bar{I}_e) + \mu}.$$

We thus have the total number of infected individuals at equilibrium,  $I_e$ , as a function of the proportion of infected individuals at equilibrium,  $\bar{I}_e$ . This really is quite interesting since we also have the bifurcation parameter  $\psi$  as a function of  $\bar{I}_e$  (see (3.1.1)). This means that we have a pair of parametric equations connecting  $I_e$  and  $\psi$ , and may thus plot bifurcation curves for the total number of infected individuals rather than just the proportion in the core. What is also interesting is that even though  $C_e$  is, under our current assumptions, a decreasing function of  $\bar{I}_e$ , it is quite possible that  $I_e$  may not be a decreasing of  $\bar{I}_e$ . It turns out, as a consequence of this, that these new bifurcation curves can have rather surprising shapes. Hadeler and Castillo-Chavez cite  $r(\bar{I}) = k_1(1-k_2\bar{I})$  as an example for the recruitment rate, where  $k_1$  and  $k_2$  are positive real numbers. In this case the parametric equations are given by

$$I_{e}(\bar{I}_{e}) = \frac{k_{1}P\bar{I}_{e}(1-k_{2}\bar{I}_{e})}{(k_{1}+\mu)-k_{1}k_{2}\bar{I}_{e}} \quad \text{and} \quad \psi(\bar{I}_{e}) = (\beta-\tilde{\beta})\frac{\alpha(1-\gamma)\bar{I}_{e}+\mu}{\tilde{\beta}\bar{I}_{e}+(\alpha+\mu-\tilde{\beta})} - (\beta\bar{I}_{e}+\mu).$$



**Absolute Bifurcation Diagram** 

FIGURE 9.1: A bifurcation diagram showing the endemic equilibria for the absolute number of infected individuals in the core for the extended CG model.

For the extended CG model we have, for absolute numbers of infected individuals, the bifurcation curve defined by the following parametric equations:

$$I_{e}(\bar{I}_{e}) = \frac{k_{1}P\bar{I}_{e}(1-k_{2}\bar{I}_{e})}{(k_{1}+\mu)-k_{1}k_{2}\bar{I}_{e}},$$

$$\psi_{1}(\bar{I}_{e}) = -\frac{W\bar{I}_{e}^{3} + \hat{X}\bar{I}_{e}^{2} + \hat{Y}\bar{I}_{e} + \hat{Z}}{\beta_{1}\beta_{2}\bar{I}_{e}^{2} + \{\alpha\beta_{1}\gamma_{2} + \alpha\beta_{2}(1-\gamma_{2}) + \mu(\beta_{1}+\beta_{2}) - \beta_{1}\beta_{2}\}\bar{I}_{e} + \mu(\alpha+\mu-\beta_{1})}$$

with  $W, \hat{X}, \hat{Y}$  and  $\hat{Z}$  as given previously. This leads to some particularly complex and interesting shapes of bifurcation curves. For example, the following set of parameter values (similar to the ones used by Hadeler and Castillo-Chavez):

$$\alpha = 4.0$$
,  $\mu = 0.1$ ,  $\beta_1 = 2.5$ ,  $\beta_2 = 5.5$ ,  $\beta_3 = 6.0$ ,  $\gamma_1 = 0.07$  and  $\gamma_2 = 0.04$ ,

gives rise to the bifurcation diagram in Figure 9.1 where we have used  $k_1 = 0.1$  and  $k_2 = 3.0$  (again, as suggested by Hadeler and Castillo-Chavez) and a total population size of 10000.

From Figure 9.1 it can be seen that there is not only the potential for multiple endemic equilibria to be present but that there is also the possibility that distinct values of the bifurcation parameter may give rise to the same absolute number of infected individuals at endemic equilibrium. It needs to be borne in mind, however, that in such cases the size of the core group will not generally be the same. This highlights a potential difficulty in this case with regard to estimating the size of the population at risk or the size of the core group from prevalence data.

#### **10 Conclusions**

In this paper we have investigated some of the potential behaviours, in the region of the parameter space near  $R_0 = 1$ , of both a core group model for the spread of a sexually transmitted disease and an extended version of this basic model incorporating two educated classes. While it might appear obvious that increased model complexity will give rise to more complicated dynamics, the actual nature of the resulting phenomena might not be so clear-cut. Our aim here has been to highlight a number of the interesting dynamic possibilities that do arise when considering models of increasing complexity.

We proved first that the local asymptotic stability of any endemic equilibrium for the basic CG model is related to the gradient of the bifurcation curve at the point representing this equilibrium. This implies in particular that in the presence of backward bifurcation the resultant upper and lower endemic equilibria are always LAS and unstable respectively. Although in the deterministic model this means that there is the potential for a subcritical endemic equilibrium to exist, it is not generally possible for the disease to invade at the DFE when  $R_0 < 1$  on the introduction of a small number of infected individuals.

For the extended model we found that there is the possibility for bifurcation diagrams to exhibit forward bifurcation while simultaneously possessing two subcritical endemic equilibria for some values of the bifurcation parameter. These diagrams possess similar shapes to the ones obtained for the three-stage BRSV model. The stability patterns also become less predictable, and there was certainly no simple result corresponding to the one for the basic CG model relating the local asymptotic stability of endemic equilibria to the gradient of the bifurcation curve. Indeed, we even encountered an example of a unique but unstable supercritical endemic equilibrium (noting here that this was never observed on any of the bifurcation diagrams obtained for the three-stage BRSV model).

Stochastic versions of the CG models were then developed. These were used in conjunction with extensive simulations in order to investigate the probability of extinction of the disease. In particular we observed subcritical probabilities of extinction that were less than one in circumstances for which backward bifurcation was present in the deterministic model (bearing in mind the caveats in this regard discussed by Griffiths (2007)). It was found that population size was a crucial factor in determining these probabilities. The disease seems most able to invade and then sustain a long-term quasi-equilibrium when  $R_0 < 1$  for intermediate population sizes. Heuristic explanations for this phenomenon were provided. In the simulations we also observed the invasion of a disease which was, because of the changing population structure, not able to sustain itself in the form of a long-lived quasi-equilibrium.

Whereas the behaviour with regard to the probability of extinction was comparable to that seen in the two and three-stage BRSV models, the results on the expected time to extinction did prove to be rather more conclusive than for the BRSV case. For the extended CG model we observed a situation in which the presence of backward bifurcation resulted in the simulated expected time to extinction being somewhat larger than the theoretical value. We argued that the presence of a nearby deterministic endemic equilibrium may indeed manifest itself as a very weak attractor for the stochastic system, resulting in a very small proportion of relatively long-lived, but still ecologically unstable, quasi-equilibria.

Bifurcation diagrams were also plotted for the full population models. This was achieved by obtaining a pair of parametric equations for each of these curves. The resultant bifurcation curves for the extended version of the full model proved to be particularly interesting in that we were able to observe not only the presence of multiple equilibria but also the existence of endemic equilibria with the same number of infected individuals for distinct values of the bifurcation parameter.

In the future some of the work carried out here may help explain certain subtle dynamic phenomena observed in real-life epidemics. Indeed, we might ask ourselves whether any of our findings do actually relate to previously unexpected behaviours observed during the spread of a disease through a population. Further realism (and hence complexity) of the CG model might be achieved by extending it to yet more educated classes or by building in some spatial element.

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More work is also needed to ascertain whether some of the interesting bifurcation diagrams observed here will ever occur for realistic parameter values. This of course raises the issue of parameter estimation.

We have started to explore the interaction between deterministic and stochastic phenomena arising when attempting to make an epidemic model more realistic and, as a consequence, more complex. There is certainly still a long way to go in this regard, although it would appear that the presence of backward bifurcation in the deterministic differential equation model has repercussions for the subcritical dynamic behaviour in the stochastic formulation. In investigating the CG model we have been able to corroborate many of our findings, with respect to the phenomena of backward bifurcation and multiple endemic equilibria, for the BRSV models. In addition, a number of new behaviours have been observed.

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